

Package ‘mev’

April 15, 2026

Type Package

Title Modelling of Extreme Values

Version 2.2

Description Various tools for the analysis of univariate, multivariate and functional extremes. Exact simulation from max-stable processes (Dombry, Engelke and Oestling, 2016, <[doi:10.1093/biomet/asw008](https://doi.org/10.1093/biomet/asw008)>, R-Pareto processes for various parametric models, including Brown-Resnick (Wadsworth and Tawn, 2014, <[doi:10.1093/biomet/ast042](https://doi.org/10.1093/biomet/ast042)>) and Extremal Student (Thibaud and Opitz, 2015, <[doi:10.1093/biomet/asv045](https://doi.org/10.1093/biomet/asv045)>). Threshold selection methods, including Wadsworth (2016) <[doi:10.1080/00401706.2014.998345](https://doi.org/10.1080/00401706.2014.998345)>, and Northrop and Coleman (2014) <[doi:10.1007/s10687-014-0183-z](https://doi.org/10.1007/s10687-014-0183-z)>. Multivariate extreme diagnostics. Estimation and likelihoods for univariate extremes, e.g., Coles (2001) <[doi:10.1007/978-1-4471-3675-0](https://doi.org/10.1007/978-1-4471-3675-0)>.

License GPL-3

URL <https://lbelzile.github.io/mev/>, <https://github.com/lbelzile/mev/>

BugReports <https://github.com/lbelzile/mev/issues/>

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Rsolnp, stats,

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| | |
|---------------|------------------------|
| <i>abisko</i> | <i>Abisko rainfall</i> |
|---------------|------------------------|

Description

Daily non-zero rainfall measurements in Abisko (Sweden) from January 1913 until December 2014.

Arguments

| | |
|--------|-------------------------|
| date | Date of the measurement |
| precip | rainfall amount (in mm) |

Format

a data frame with 15132 rows and two variables

Source

Abisko Scientific Research Station, Swedish Meteorological and Hydrological Institute, distributed under a Creative Commons Attribution 4.0 SE license

References

A. Kiriliouk, H. Rootzen, J. Segers and J.L. Wadsworth (2019), *Peaks over thresholds modeling with multivariate generalized Pareto distributions*, *Technometrics*, **61**(1), 123–135, <doi:10.1080/00401706.2018.1462738>

Description

Estimation of the bivariate angular dependence function

Usage

```
adf(
  xdat,
  qllev = 0.95,
  estimator = c("hill", "mle", "bayes"),
  level = 0.95,
  ties.method = "random",
  angles = seq(0, 1, by = 0.02),
  plot = TRUE
)
```

Arguments

| | |
|-------------|--|
| xdat | an n by 2 matrix of multivariate observations |
| qllev | quantile level on uniform scale at which to threshold data. Default to 0.95 |
| estimator | string indicating the estimation method |
| level | level for confidence intervals, default to 0.95 |
| ties.method | method for handling of ties in rank transformation |
| angles | vector of angles at which to evaluate the angular dependence function The confidence intervals are based on normal quantiles. The standard errors for the hill are based on the asymptotic covariance and that of the mle derived using the delta-method. Bayesian posterior predictive interval estimates are obtained using ratio-of-uniform sampling with flat priors: the shape parameters are constrained to lie within the triangle, as are frequentist point estimates which are adjusted post-inference. |
| plot | logical indicating whether to plot the function, defaults to TRUE |

Value

a plot of the angular dependence function if plot=TRUE, plus an invisible list with components

- angle the sequence of angles in (0,1) at which the lambda values are evaluated
- coef point estimates of the angular dependence function
- lower level% confidence interval for lambda (lower bound)
- upper level% confidence interval for lambda (upper bound)

References

J.L. Wadsworth and J.A. Tawn (2013). A new representation for multivariate tail probabilities, *Bernoulli*, 19(5B), 2689-2714.

Examples

```
set.seed(12)
dat <- mev::rmev(n = 1000, d = 2, model = "log", param = 0.1)
adf(xdat = dat, estimator = 'hill')
```

angmeas

Rank-based transformation to angular measure

Description

The method uses the pseudo-polar transformation for suitable norms, transforming the data to pseudo-observations, then marginally to unit Frechet or unit Pareto. Empirical or Euclidean weights are computed and returned alongside with the angular and radial sample for values above threshold(s) thresh, specified in terms of quantiles of the radial component R or marginal quantiles. Only complete tuples are kept.

Usage

```
angmeas(
  xdat,
  thresh,
  Rnorm = c("l1", "l2", "linf"),
  Anorm = c("l1", "l2", "linf", "arctan"),
  marg = c("frechet", "pareto"),
  wgt = c("empirical", "euclidean"),
  region = c("sum", "min", "max"),
  is.angle = FALSE,
  ...
)
```

Arguments

| | |
|--------|---|
| xdat | an n by d sample matrix |
| thresh | threshold of length 1 for 'sum', or d marginal thresholds otherwise. |
| Rnorm | character string indicating the norm for the radial component. |
| Anorm | character string indicating the norm for the angular component. arctan is only implemented for $d = 2$ |
| marg | character string indicating choice of marginal transformation, either to Frechet or Pareto scale |
| wgt | character string indicating weighting function for the equation. Can be based on Euclidean or empirical likelihood for the mean |

| | |
|----------|---|
| region | character string specifying which observations to consider (and weight). 'sum' corresponds to a radial threshold $\sum x_i > \text{thresh}$, 'min' to $\min x_i > \text{thresh}$ and 'max' to $\max x_i > \text{thresh}$. |
| is.angle | logical indicating whether observations are already angle with respect to region. Default to FALSE. |
| ... | additional arguments |

Details

The empirical likelihood weighted mean problem is implemented for all thresholds, while the Euclidean likelihood is only supported for diagonal thresholds specified via region=sum.

Value

a list with arguments ang for the $d - 1$ pseudo-angular sample, rad with the radial component and possibly wts if Rnorm='l1' and the empirical likelihood algorithm converged. The Euclidean algorithm always returns weights even if some of these are negative.

a list with components

- ang matrix of pseudo-angular observations
- rad vector of radial contributions
- wts empirical or Euclidean likelihood weights for angular observations

Author(s)

Leo Belzile

References

Einmahl, J.H.J. and J. Segers (2009). Maximum empirical likelihood estimation of the spectral measure of an extreme-value distribution, *Annals of Statistics*, **37**(5B), 2953–2989.

de Carvalho, M. and B. Oumow and J. Segers and M. Warchol (2013). A Euclidean likelihood estimator for bivariate tail dependence, *Comm. Statist. Theory Methods*, **42**(7), 1176–1192.

Owen, A.B. (2001). *Empirical Likelihood*, CRC Press, 304p.

Examples

```
x <- rmev(n = 25, d = 3, param = 0.5, model = 'log')
wts <- angmeas(xdat = x, Rnorm = 'l1', Anorm = 'l1', marg = 'frechet', wgt = 'empirical')
wts2 <- angmeas(xdat = x, Rnorm = 'l2', Anorm = 'l2', marg = 'pareto')
```

angmeasdir

Dirichlet mixture smoothing of the angular measure

Description

This function computes the empirical or Euclidean likelihood estimates of the spectral measure and uses the points returned from a call to `angmeas` to compute the Dirichlet mixture smoothing of de Carvalho, Warchol and Segers (2012), placing a Dirichlet kernel at each observation.

Usage

```
angmeasdir(
  xdat,
  thresh,
  Rnorm = c("l1", "l2", "linf"),
  Anorm = c("l1", "l2", "linf", "arctan"),
  marg = c("frechet", "pareto"),
  wgt = c("empirical", "euclidean"),
  region = c("sum", "min", "max"),
  is.angle = FALSE,
  ...
)
```

Arguments

| | |
|-----------------------|---|
| <code>xdat</code> | an n by d sample matrix |
| <code>thresh</code> | threshold of length 1 for 'sum', or d marginal thresholds otherwise. |
| <code>Rnorm</code> | character string indicating the norm for the radial component. |
| <code>Anorm</code> | character string indicating the norm for the angular component. <code>arctan</code> is only implemented for $d = 2$ |
| <code>marg</code> | character string indicating choice of marginal transformation, either to Frechet or Pareto scale |
| <code>wgt</code> | character string indicating weighting function for the equation. Can be based on Euclidean or empirical likelihood for the mean |
| <code>region</code> | character string specifying which observations to consider (and weight). 'sum' corresponds to a radial threshold $\sum x_i > \text{thresh}$, 'min' to $\min x_i > \text{thresh}$ and 'max' to $\max x_i > \text{thresh}$. |
| <code>is.angle</code> | logical indicating whether observations are already angle with respect to region. Default to FALSE. |
| <code>...</code> | additional arguments |

Details

The cross-validation bandwidth is the solution of

$$\max_{\nu} \sum_{i=1}^n \log \left\{ \sum_{k=1, k \neq i}^n p_{k,-i} f(\mathbf{w}_i; \nu \mathbf{w}_k) \right\},$$

where f is the density of the Dirichlet distribution, $p_{k,-i}$ is the Euclidean weight obtained from estimating the Euclidean likelihood problem without observation i .

Value

an invisible list with components

- nu bandwidth parameter obtained by cross-validation;
- dirparmat n by d matrix of Dirichlet parameters for the mixtures;
- wts mixture weights.

Examples

```
set.seed(123)
x <- rmev(n = 100, d = 2L, param = 0.5, model = 'log')
out <- angmeasdir(x)
```

build.blocks

Compute block maxima and order them by block

Description

Given a time series of observations in `xdat`, compute the maximum of blocks of size `block` (b), and then order them by further blocks of size `m`, increasing by row from left to right. If the length of `xdat` is not a multiple of `block`, the last observations are discarded without warning.

Usage

```
build.blocks(xdat, block = 1L, m = 2L)
```

Arguments

| | |
|--------------------|---|
| <code>xdat</code> | vector of length <code>n</code> |
| <code>block</code> | integer, size of block over which to compute maxima |
| <code>m</code> | number of columns for further sub-blocking |

Value

a matrix with $\lfloor n/b \rfloor$ observations, ordered by row, with `m` columns.

| | |
|----------------|-----------------------------------|
| cheeseborowind | <i>Cheeseboro wind speed data</i> |
|----------------|-----------------------------------|

Description

Daily measurements of wind speed during the month of January and including February 1st, from the Cheeseboro (California) weather station. between 1996 and 2026.

Usage

```
cheeseborowind
```

Format

A data frame with 992 rows and 3 variables:

date date of observation

direction angle (in degrees) of the wind

gust maximum daily wind speed (in meters per second)

Source

Raw US Climate Archive, <https://raws.dri.edu/cgi-bin/rawMAIN.pl?caCCHB>, maintained by the Western Regional Climate Center, Desert Research Institute based in Reno, Nevada

| | |
|---------------|--|
| confint.eprof | <i>Confidence intervals for profile likelihood objects</i> |
|---------------|--|

Description

Computes confidence intervals for the parameter psi for profile likelihood objects. This function uses spline interpolation to derive level confidence intervals

Usage

```
## S3 method for class 'eprof'  
confint(  
  object,  
  parm,  
  level = 0.95,  
  prob = c((1 - level)/2, 1 - (1 - level)/2),  
  print = FALSE,  
  method = c("cobs", "smooth.spline"),  
  boundary = FALSE,  
  ...  
)
```

Arguments

| | |
|----------|--|
| object | an object of class eprof, normally the output of <code>gpd.pll</code> or <code>gev.pll</code> . |
| parm | a specification of which parameters are to be given confidence intervals, either a vector of numbers or a vector of names. If missing, all parameters are considered. |
| level | confidence level, with default value of 0.95 |
| prob | percentiles, with default giving symmetric 95% confidence intervals |
| print | should a summary be printed. Default to FALSE. |
| method | string for the method, either <code>cobs</code> (constrained robust B-spline from eponym package) or <code>smooth.spline</code> |
| boundary | logical; if TRUE, the null distribution is assumed to be a mixture of a point mass and half a chi-square with one degree of freedom. |
| ... | additional arguments passed to functions. Providing a logical <code>warn=FALSE</code> turns off warning messages when the lower or upper confidence interval for ψ are extrapolated beyond the provided calculations. |

Value

returns a 2 by 3 matrix containing point estimates, lower and upper confidence intervals based on the likelihood root and modified version thereof

dgeoaniso

Distance matrix with geometric anisotropy

Description

The function computes the distance between locations, with geometric anisotropy. Consider real parameters θ_1 and θ_2 , and the transformation $\psi = \arctan(\theta_1/\theta_2)/2$ and $r = 1 + \theta_1^2 + \theta_2^2$. The dilation and rotation matrix is

$$\begin{pmatrix} \sqrt{r} \cos(\rho) & -\sqrt{r} \sin(\rho) \\ \sin(\rho)/\sqrt{r} & \cos(\rho)/\sqrt{r} \end{pmatrix}.$$

The parametrization is convenient for optimization purposes, as the parameter vector is unconstrained and the transformation has unit Jacobian.

Usage

```
dgeoaniso(loc, theta)
```

Arguments

| | |
|-------|--|
| loc | a d by 2 matrix of locations giving the coordinates of a site per row. |
| theta | numeric vector of length 2, real parameters |

Value

a d by d square matrix of pairwise distance

References

Rai, K. and Brown, P.E. (2025), A parameter transformation of the anisotropic Matérn covariance function. Canadian Journal of Statistics e11839. doi:10.1002/cjs.11839

 egp

Extended generalised Pareto families

Description

This function provides the log-likelihood and quantiles for the three different families presented in Papastathopoulos and Tawn (2013) and the two proposals of Gamet and Jalbert (2022), plus exponential tilting. All of the models contain an additional parameter, $\kappa \geq 0$. All families share the same tail index as the generalized Pareto distribution, while allowing for lower thresholds. For most models, the distribution reduce to the generalised Pareto when $\kappa = 1$ (for models `gj-tnorm` and `logist`, on the boundary of the parameter space when $\kappa \rightarrow 0$).

`egp.retlev` gives the return levels for the extended generalised Pareto distributions

Arguments

| | |
|---------------------|---|
| <code>xdat</code> | vector of observations, greater than the threshold |
| <code>thresh</code> | threshold value |
| <code>par</code> | parameter vector (κ, σ, ξ) . |
| <code>model</code> | a string indicating which extended family to fit |
| <code>show</code> | logical; if TRUE, print the results of the optimization |
| <code>p</code> | extreme event probability; <code>p</code> must be greater than the rate of exceedance for the calculation to make sense. See Details . |
| <code>plot</code> | logical; if TRUE, a plot of the return levels |

Details

For return levels, the `p` argument can be related to T year exceedances as follows: if there are n_y observations per year, than take `p` to equal $1/(Tn_y)$ to obtain the T -years return level.

Value

`egp.ll` returns the log-likelihood value, while `egp.retlev` returns a plot of the return levels if `plot=TRUE` and a list with tail probabilities `p`, return levels `retlev`, thresholds `thresh` and model name `model`.

Usage

```
egp.ll(xdat, thresh, model, par)
egp.retlev(xdat, thresh, par, model, p, plot=TRUE)
```

Author(s)

Leo Belzile

References

Papastathopoulos, I. and J. Tawn (2013). Extended generalised Pareto models for tail estimation, *Journal of Statistical Planning and Inference* **143**(3), 131–143, <doi:10.1016/j.jspi.2012.07.001>.

Gamet, P. and Jalbert, J. (2022). A flexible extended generalized Pareto distribution for tail estimation. *Environmetrics*, **33**(6), <doi:10.1002/env.2744>.

Examples

```
set.seed(123)
xdat <- rgp(1000, loc = 0, scale = 2, shape = 0.5)
par <- fit.egp(xdat, thresh = 0, model = 'gj-beta')$par
p <- c(1/1000, 1/1500, 1/2000)
# With multiple thresholds
th <- c(0, 0.1, 0.2, 1)
opt <- tstab.egp(xdat, thresh = th, model = 'gj-beta')
egp.retlev(xdat = xdat, thresh = th, model = 'gj-beta', p = p)
opt <- tstab.egp(xdat, th, model = 'pt-power', plots = NA)
egp.retlev(xdat = xdat, thresh = th, model = 'pt-power', p = p)
```

egp.pll

Profile log likelihood for extended generalized Pareto models

Description

Computes the profile log likelihood over a grid of values of ψ for various parameters, including return levels.

Usage

```
egp.pll(
  psi,
  model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
    "logist"),
  param = c("kappa", "scale", "shape", "retlev"),
  mle = NULL,
  xdat,
  thresh = NULL,
  plot = FALSE,
```

```

    method = c("Nelder", "nlminb", "BFGS"),
    p,
    ...
)

```

Arguments

| | |
|--------|---|
| psi | grid of values for the parameter to profile |
| model | string; choice of extended eneralized Pareto model. |
| param | string; parameter to profile |
| mle | a vector or matrix with maximum likelihood estimates of kappa, scale, shape. This can be a matrix if there are multiple threshold |
| xdat | vector of observations |
| thresh | vector of positive thresholds. If NULL, defaults to zero. |
| plot | logical; if TRUE, returns a plot of the profile log likelihood |
| method | string giving the optimization method for the outer optimization in the augmented Lagrangian routine; one of nlminb or BFGS |
| p | tail probability for return level if param="retlev". |
| ... | additional arguments, currently ignored |

Value

an object of class eprof

| | |
|---------|---|
| egpdist | <i>Extended generalized Pareto distribution</i> |
|---------|---|

Description

Density function, distribution function, quantile function and random number generation for various extended generalized Pareto distributions

Usage

```

pegp(
  q,
  scale,
  shape,
  kappa,
  model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
    "logist"),
  lower.tail = TRUE,
  log.p = FALSE
)

```

```

degp(
  x,
  scale,
  shape,
  kappa,
  model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
            "logist"),
  log = FALSE
)

qegp(
  p,
  scale,
  shape,
  kappa,
  model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
            "logist"),
  lower.tail = TRUE,
  log.p = FALSE
)

regp(
  n,
  scale,
  shape,
  kappa,
  model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
            "logist")
)

```

Arguments

| | |
|------------|---|
| scale | scale parameter, strictly positive. |
| shape | shape parameter. |
| kappa | shape parameter for the tilting distribution. |
| model | string giving the distribution of the model |
| lower.tail | logical; if TRUE (default), the lower tail probability $\Pr(X \leq x)$ is returned. |
| log.p, log | logical; if FALSE (default), values are returned on the probability scale. |
| x, q | vector of quantiles |
| p | vector of probabilities |
| n | scalar number of observations |

References

Papastathopoulos, I. and J. Tawn (2013). Extended generalised Pareto models for tail estimation, *Journal of Statistical Planning and Inference* **143**(3), 131–143, <doi:10.1016/j.jspi.2012.07.001>.

Gamet, P. and Jalbert, J. (2022). A flexible extended generalized Pareto distribution for tail estimation. *Environmetrics*, **33**(6), <doi:10.1002/env.2744>.

 emplik

Self-concordant empirical likelihood for a vector mean

Description

Self-concordant empirical likelihood for a vector mean

Usage

```
emplik(
  dat,
  mu = rep(0, ncol(dat)),
  lam = rep(0, ncol(dat)),
  eps = 1/nrow(dat),
  M = 1e+30,
  thresh = 1e-30,
  itermax = 100
)
```

Arguments

| | |
|---------|---|
| dat | n by d matrix of d-variate observations |
| mu | d vector of hypothesized mean of dat |
| lam | starting values for Lagrange multiplier vector, default to zero vector |
| eps | lower cutoff for $-\log$, with default $1/nrow(dat)$ |
| M | upper cutoff for $-\log$. |
| thresh | convergence threshold for log likelihood (default of $1e-30$ is aggressive) |
| itermax | upper bound on number of Newton steps. |

Value

a list with components

- loge1r log empirical likelihood ratio.
- lam Lagrange multiplier (vector of length d).
- wts n vector of observation weights (probabilities).
- conv boolean indicating convergence.
- niter number of iteration until convergence.
- ndec Newton decrement.
- gradnorm norm of gradient of log empirical likelihood.

Author(s)

Art Owen, C++ port by Leo Belzile

References

Owen, A.B. (2013). Self-concordance for empirical likelihood, *Canadian Journal of Statistics*, **41**(3), 387–397.

eskrain

Eskdalemuir Observatory Daily Rainfall

Description

This dataset contains exceedances of 30mm for daily cumulated rainfall observations over the period 1970-1986. These data were aggregated from hourly series.

Format

a vector with 93 daily cumulated rainfall measurements exceeding 30mm.

Details

The station is one of the rainiest of the whole UK, with an average 1554mm of cumulated rainfall per year. The data consisted of 6209 daily observations, of which 4409 were non-zero. Only the 93 largest observations are provided.

Source

Met Office.

expme

Exponent measure for multivariate generalized Pareto distributions

Description

Integrated intensity over the region defined by $[0, z]^c$ for logistic, Huesler-Reiss, Brown-Resnick and extremal Student processes.

Usage

```
expme(
  z,
  par,
  model = c("log", "neglog", "hr", "br", "xstud"),
  method = c("TruncatedNormal", "mvtnorm", "mvPot")
)
```

Arguments

| | |
|--------|---|
| z | vector at which to estimate exponent measure |
| par | list of parameters |
| model | string indicating the model family |
| method | string indicating the package from which to extract the numerical integration routine |

Value

numeric giving the measure of the complement of $[0, z]$.

Note

The list par must contain different arguments depending on the model. For the Brown–Resnick model, the user must supply the conditionally negative definite matrix Lambda following the parametrization in Engelke *et al.* (2015) or the covariance matrix Sigma, following Wadsworth and Tawn (2014). For the Husler–Reiss model, the user provides the mean and covariance matrix, m and Sigma. For the extremal student, the covariance matrix Sigma and the degrees of freedom df. For the logistic model, the strictly positive dependence parameter alpha.

Examples

```
## Not run:
# Extremal Student
Sigma <- stats::rWishart(n = 1, df = 20, Sigma = diag(10))[, , 1]
expme(z = rep(1, ncol(Sigma)), par = list(Sigma = cov2cor(Sigma), df = 3), model = "xstud")
# Brown-Resnick model
D <- 5L
loc <- cbind(runif(D), runif(D))
di <- as.matrix(dist(rbind(c(0, ncol(loc)), loc)))
semivario <- function(d, alpha = 1.5, lambda = 1) {
  (d / lambda)^alpha
}
Vmat <- semivario(di)
Lambda <- Vmat[-1, -1] / 2
expme(z = rep(1, ncol(Lambda)), par = list(Lambda = Lambda), model = "br", method = "mvPot")
Sigma <- outer(Vmat[-1, 1], Vmat[1, -1], "+") - Vmat[-1, -1]
expme(z = rep(1, ncol(Lambda)), par = list(Lambda = Lambda), model = "br", method = "mvPot")

## End(Not run)
```

Description

Density function, distribution function, quantile function and random generation for the extended generalized Pareto distribution (GPD) with scale and shape parameters.

Arguments

| | |
|-----------|--|
| q | vector of quantiles |
| x | vector of observations |
| p | vector of probabilities |
| n | sample size |
| prob | mixture probability for model type 4 |
| kappa | shape parameter for type 1, 3 and 4 |
| delta | additional parameter for type 2, 3 and 4 |
| sigma | scale parameter |
| xi | shape parameter |
| type | integer between 0 to 5 giving the model choice |
| step | function of step size for discretization with default 0, corresponding to continuous quantiles |
| log | logical; should the log-density be returned (default to FALSE)? |
| unifsamp | sample of uniform; if provided, the data will be used in place of new uniform random variates |
| censoring | numeric vector of length 2 containing the lower and upper bound for censoring |

Details

The extended generalized Pareto families proposed in Naveau *et al.* (2016) retain the tail index of the distribution while being compliant with the theoretical behavior of extreme low rainfall. There are five proposals, the first one being equivalent to the GP distribution.

- type 0 corresponds to uniform carrier, $G(u) = u$.
- type 1 corresponds to a three parameters family, with carrier $G(u) = u^\kappa$.
- type 2 corresponds to a three parameters family, with carrier $G(u) = 1 - V_\delta((1 - u)^\delta)$.
- type 3 corresponds to a four parameters family, with carrier

$$G(u) = 1 - V_\delta((1 - u)^\delta))^{\kappa/2}$$

- type 4 corresponds to a five parameter model (a mixture of type 2, with $G(u) = pu^\kappa + (1 - p) * u^\delta$

Usage

```
pextgp(q, prob=NA, kappa=NA, delta=NA, sigma=NA, xi=NA, type=1)
dextgp(x, prob=NA, kappa=NA, delta=NA, sigma=NA, xi=NA, type=1, log=FALSE)
qextgp(p, prob=NA, kappa=NA, delta=NA, sigma=NA, xi=NA, type=1)
rextgp(n, prob=NA, kappa=NA, delta=NA, sigma=NA, xi=NA, type=1, unifsamp=NULL, censoring=c(0,Inf))
```

Author(s)

Raphael Huser and Philippe Naveau

References

Naveau, P., R. Huser, P. Ribereau, and A. Hannart (2016), Modeling jointly low, moderate, and heavy rainfall intensities without a threshold selection, *Water Resour. Res.*, 52, 2753-2769, doi : 10.1002/2015WR018552.

extgp.G

Carrier distribution for the extended GP distributions of Naveau et al.

Description

Density, distribution function, quantile function and random number generation for the carrier distributions of the extended Generalized Pareto distributions.

Arguments

| | |
|-----------|--|
| u | vector of observations (dextgp.G), probabilities (qextgp.G) or quantiles (pextgp.G), in [0, 1] |
| prob | mixture probability for model type 4 |
| kappa | shape parameter for type 1, 3 and 4 |
| delta | additional parameter for type 2, 3 and 4 |
| type | integer between 0 to 5 giving the model choice |
| log | logical; should the log-density be returned (default to FALSE)? |
| n | sample size |
| unifsamp | sample of uniform; if provided, the data will be used in place of new uniform random variates |
| censoring | numeric vector of length 2 containing the lower and upper bound for censoring |
| direct | logical; which method to use for sampling in model of type 4? |

Usage

```
pextgp.G(u, type=1, prob, kappa, delta)
dextgp.G(u, type=1, prob=NA, kappa=NA, delta=NA, log=FALSE)
qextgp.G(u, type=1, prob=NA, kappa=NA, delta=NA)
rextgp.G(n, prob=NA, kappa=NA, delta=NA, type=1, unifsamp=NULL, direct=FALSE, censoring=c(0,1))
```

Author(s)

Raphael Huser and Philippe Naveau

See Also

[extgp](#)

fit.egp

Parameter stability plot and maximum likelihood routine for extended GP models

Description

The function `tstab.egp` provides classical threshold stability plot for (κ, σ, ξ) . The fitted parameter values are displayed with pointwise normal 95% confidence intervals. The function returns an invisible list with parameter estimates and standard errors, and p-values for the Wald test that $\kappa = 1$. The plot is for the modified scale (as in the generalised Pareto model) and as such it is possible that the modified scale be negative. `tstab.egp` can also be used to fit the model to multiple thresholds.

Usage

```
fit.egp(
  xdat,
  thresh = 0,
  model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
            "logist"),
  start = NULL,
  method = c("Nelder", "nlminb", "BFGS"),
  fpar = NULL,
  show = FALSE,
  ...
)
```

Arguments

| | |
|---------------------|--|
| <code>xdat</code> | vector of observations, greater than the threshold |
| <code>thresh</code> | threshold value |
| <code>model</code> | a string indicating which extended family to fit |
| <code>start</code> | optional named list of initial values, with κ , <i>sigma</i> or <i>xi</i> . |
| <code>method</code> | the method to be used. See Details . Can be abbreviated. |
| <code>fpar</code> | a named list with fixed parameters, either scale or shape |
| <code>show</code> | logical; if TRUE, print the results of the optimization |
| <code>...</code> | additional parameters, for backward compatibility purposes |

Details

`fit.egp` is a numerical optimization routine to fit the extended generalised Pareto models of Papasathopoulos and Tawn (2013), using maximum likelihood estimation.

Value

`fit.egp` outputs the list returned by `optim`, which contains the parameter values, the hessian and in addition the standard errors

`tstab.egp` returns a plot(s) of the parameters fit over the range of provided thresholds, with point-wise normal confidence intervals; the function also returns an invisible list containing notably the matrix of point estimates (`par`) and standard errors (`se`).

Author(s)

Leo Belzile

References

Papastathopoulos, I. and J. Tawn (2013). Extended generalised Pareto models for tail estimation, *Journal of Statistical Planning and Inference* **143**(3), 131–143.

Examples

```
xdat <- mev::rgp(
  n = 100,
  loc = 0,
  scale = 1,
  shape = 0.5)
fitted <- fit.egp(
  xdat = xdat,
  thresh = 1,
  model = "pt-gamma",
  show = TRUE)
thresh <- mev::qgp(seq(0.1, 0.5, by = 0.05), 0, 1, 0.5)
tstab.egp(
  xdat = xdat,
  thresh = thresh,
  model = "pt-gamma")
xdat <- regp(
  n = 100,
  scale = 1,
  shape = 0.1,
  kappa = 0.5,
  model = "pt-power"
)
fit.egp(
  xdat = xdat,
  model = "pt-power",
  show = TRUE,
  fpar = list(kappa = 1),
  method = "Nelder"
)
```

fit.extgp

Fit an extended generalized Pareto distribution of Naveau et al.

Description

This is a wrapper function to obtain PWM or MLE estimates for the extended GP models of Naveau et al. (2016) for rainfall intensities. The function calculates confidence intervals by means of nonparametric percentile bootstrap and returns histograms and QQ plots of the fitted distributions. The function handles both censoring and rounding.

Usage

```
fit.extgp(
  data,
  model = 1,
  method = c("mle", "pwm"),
  init,
  censoring = c(0, Inf),
  rounded = 0,
  confint = FALSE,
  R = 1000,
  ncpus = 1,
  plots = TRUE
)
```

Arguments

| | |
|-----------|--|
| data | data vector. |
| model | integer ranging from 0 to 4 indicating the model to select (see extgp). |
| method | string; either 'mle' for maximum likelihood, or 'pwm' for probability weighted moments, or both. |
| init | vector of initial values, comprising of p , κ , δ , σ , ξ (in that order) for the optimization. All parameters may not appear depending on model. |
| censoring | numeric vector of length 2 containing the lower and upper bound for censoring; censoring=c(0, Inf) is equivalent to no censoring. |
| rounded | numeric giving the instrumental precision (and rounding of the data), with default of 0. |
| confint | logical; should confidence interval be returned (percentile bootstrap). |
| R | integer; number of bootstrap replications. |
| ncpus | integer; number of CPUs for parallel calculations (default: 1). |
| plots | logical; whether to produce histogram and density plots. |

Details

The different models include the following transformations:

- model 0 corresponds to uniform carrier, $G(u) = u$.
- model 1 corresponds to a three parameters family, with carrier $G(u) = u^\kappa$.
- model 2 corresponds to a three parameters family, with carrier $G(u) = 1 - V_\delta((1 - u)^\delta)$.
- model 3 corresponds to a four parameters family, with carrier

$$G(u) = 1 - V_\delta((1 - u)^\delta))^{\kappa/2}$$

- model 4 corresponds to a five parameter model (a mixture of type 2, with $G(u) = pu^\kappa + (1 - p) * u^\delta$

Author(s)

Raphael Huser and Philippe Naveau

References

Naveau, P., R. Huser, P. Ribereau, and A. Hannart (2016), Modeling jointly low, moderate, and heavy rainfall intensities without a threshold selection, *Water Resour. Res.*, 52, 2753-2769, doi : 10.1002/2015WR018552.

See Also

[egp.fit](#), [egp](#), [extgp](#)

Examples

```
## Not run:
data(rain, package = "isnev")
fit.extgp(
  rain[rain > 0],
  model = 1,
  method = 'mle',
  init = c(0.9, fit.gpd(rain)$est),
  rounded = 0.1,
  confint = TRUE,
  R = 20
)

## End(Not run)
```

| | |
|---------|---|
| fit.gev | <i>Maximum likelihood estimation for the generalized extreme value distribution</i> |
|---------|---|

Description

This function returns an object of class `mev_gev`, with default methods for printing and quantile-quantile plots. The default starting values are the solution of the probability weighted moments.

Usage

```
fit.gev(
  xdat,
  start = NULL,
  method = c("nlminb", "BFGS"),
  show = FALSE,
  fpar = NULL,
  warnSE = FALSE
)
```

Arguments

| | |
|---------------------|---|
| <code>xdat</code> | a numeric vector of data to be fitted. |
| <code>start</code> | named list of starting values |
| <code>method</code> | string indicating the outer optimization routine for the augmented Lagrangian. One of <code>nlminb</code> or <code>BFGS</code> . |
| <code>show</code> | logical; if TRUE (the default), print details of the fit. |
| <code>fpar</code> | a named list with optional fixed components <code>loc</code> , <code>scale</code> and <code>shape</code> |
| <code>warnSE</code> | logical; if TRUE, a warning is printed if the standard errors cannot be returned from the observed information matrix when the shape is less than -0.5. |

Value

a list containing the following components:

- `estimate` a vector containing the maximum likelihood estimates.
- `std.err` a vector containing the standard errors.
- `vcov` the variance covariance matrix, obtained as the numerical inverse of the observed information matrix.
- `method` the method used to fit the parameter.
- `nllh` the negative log-likelihood evaluated at the parameter estimate.
- `convergence` components taken from the list returned by [auglag](#). Values other than 0 indicate that the algorithm likely did not converge.
- `counts` components taken from the list returned by [auglag](#).
- `xdat` vector of data

Examples

```
xdat <- mev::rgev(n = 100)
fit.gev(xdat, show = TRUE)
# Example with fixed parameter
fit.gev(xdat, show = TRUE, fpar = list(shape = 0))
```

fit.gevblock

Optimization for the GEV likelihood for blocks

Description

Given a matrix of n ordered samples of m order statistics from a postulated GEV, fit the parameters of the latter based on the marginal likelihood of the first $m-1$ order statistics using maximum likelihood.

Usage

```
fit.gevblock(
  xdat,
  marginal = FALSE,
  constraint = TRUE,
  rounding = 0,
  lb = NULL,
  start = NULL,
  vcov = FALSE
)
```

Arguments

| | |
|------------|---|
| xdat | matrix of observations of size n by m , ordered by rows |
| marginal | logical; if TRUE, use marginal likelihood of lower order statistics |
| constraint | logical; if TRUE, add support constraint |
| rounding | double; indicate the amount of rounding around value; default to zero |
| lb | lower bound; any point below lb is left-censored |
| start | vector of length 3 for starting values for GEV; default to NULL |
| vcov | logical; if TRUE, return as attribute the estimate of the covariance matrix of the parameters given by the inverse observed information matrix. |

Details

One can set constraint to TRUE to add a support constraint to the optimization to ensure that all values of xdat are in the support of the resulting distribution (only for the marginal likelihood).

Value

(constrained) maximum likelihood estimator of location, scale and shape parameters

Examples

```
set.seed(2026)
xdat <- build.blocks(mev::rgev(n = 200, shape = 0.1), m = 4)
fit.gevblock(xdat, marginal = TRUE)
fit.gevblock(round(xdat, 1), marginal = TRUE, lb = NULL, rounding = 0.1)
fit.gevblock(round(xdat, 1), marginal = TRUE, lb = -2, rounding = 0.1)
fit.gevblock(xdat, marginal = TRUE, lb = -2)
fit.gevblock(xdat)
fit.gevblock(round(xdat, 1), lb = NULL, rounding = 0.1)
fit.gevblock(round(xdat, 1), lb = -2, rounding = 0.1)
fit.gevblock(xdat, lb = -2)
```

fit.gpd

Maximum likelihood estimation for the generalized Pareto distribution

Description

Numerical optimization of the generalized Pareto distribution for data exceeding threshold. This function returns an object of class `mev_gpd`, with default methods for printing and quantile-quantile plots.

Usage

```
fit.gpd(
  xdat,
  threshold = 0,
  method = c("Grimshaw", "auglag", "nlm", "optim", "ismev", "zs", "zhang", "obre", "pwm"),
  show = FALSE,
  MCMC = NULL,
  k = 4,
  tol = 1e-08,
  fpar = NULL,
  warnSE = FALSE,
  returnsamp = TRUE,
  ...
)
```

Arguments

| | |
|------------------------|---|
| <code>xdat</code> | a numeric vector of data to be fitted. |
| <code>threshold</code> | the chosen threshold. |
| <code>method</code> | the method to be used. See Details . Can be abbreviated. |
| <code>show</code> | logical; if TRUE (the default), print details of the fit. |
| <code>MCMC</code> | NULL for frequentist estimates, otherwise a boolean or a list with parameters passed. If TRUE, runs a Metropolis-Hastings sampler to get posterior mean estimates. Can be used to pass arguments <code>niter</code> , <code>burnin</code> and <code>thin</code> to the sampler as a list. |

| | |
|------------|---|
| k | bound on the influence function (method = "obre"); the constant k is a robustness parameter (higher bounds are more efficient, low bounds are more robust). Default to 4, must be larger than $\sqrt{2}$. |
| tol | numerical tolerance for OBRE weights iterations (method = "obre"). Default to 1e-8. |
| fpar | a named list with fixed parameters, either scale or shape |
| warnSE | logical; if TRUE, a warning is printed if the standard errors cannot be returned from the observed information matrix when the shape is less than -0.5. |
| returnsamp | logical; if TRUE, the object returned contains the sample vector of exceedances, which is needed for plots. This argument is useful for cases where the vector of observations takes up a lot of memory to avoid needless copies. |
| ... | additional parameters for backward compatibility |

Details

The default method is 'Grimshaw', which maximizes the profile likelihood for the ratio scale/shape. Other options include 'obre' for optimal B -robust estimator of the parameter of Dupuis (1998), vanilla maximization of the log-likelihood using constrained optimization routine 'augLag', 1-dimensional optimization of the profile likelihood using `nlm` and `optim`. Method 'isnev' performs the two-dimensional optimization routine `gpd.fit` from the `isnev` library, with in addition the algebraic gradient. The approximate Bayesian methods ('zs' and 'zhang') are extracted respectively from Zhang and Stephens (2009) and Zhang (2010) and consists of a approximate posterior mean calculated via importance sampling assuming a GPD prior is placed on the parameter of the profile likelihood.

Value

If method is neither 'zs' nor 'zhang', a list containing the following components:

- estimate a vector containing the scale and shape parameters (optimized and fixed).
- std.err a vector containing the standard errors. For method = "obre", these are Huber's robust standard errors.
- vcov the variance covariance matrix, obtained as the numerical inverse of the observed information matrix. For method = "obre", this is the sandwich Godambe matrix inverse.
- threshold the threshold.
- method the method used to fit the parameter. See details.
- nllh the negative log-likelihood evaluated at the parameter estimate.
- nat number of points lying above the threshold.
- pat proportion of points lying above the threshold.
- convergence components taken from the list returned by `optim`. Values other than 0 indicate that the algorithm likely did not converge (in particular 1 and 50).
- counts components taken from the list returned by `optim`.
- exceedances excess over the threshold.

Additionally, if `method = "obre"`, a vector of OBRE weights.

Otherwise, a list containing

- `threshold` the threshold.
- `method` the method used to fit the parameter. See **Details**.
- `nat` number of points lying above the threshold.
- `pat` proportion of points lying above the threshold.
- `approx.mean` a vector containing the approximate posterior mean estimates.

and in addition if `MCMC` is neither `FALSE`, nor `NULL`

- `post.mean` a vector containing the posterior mean estimates.
- `post.se` a vector containing the posterior standard error estimates.
- `accept.rate` proportion of points lying above the threshold.
- `niter` length of resulting Markov Chain
- `burnin` amount of discarded iterations at start, capped at 10000.
- `thin` thinning integer parameter describing

Note

Some of the internal functions (which are hidden from the user) allow for modelling of the parameters using covariates. This is not currently implemented within `gp.fit`, but users can call internal functions should they wish to use these features.

Author(s)

Scott D. Grimshaw for the Grimshaw option. Paul J. Northrop and Claire L. Coleman for the methods `optim`, `nlm` and `ismev`. J. Zhang and Michael A. Stephens (2009) and Zhang (2010) for the `zs` and `zhang` approximate methods and L. Belzile for methods `auglag` and `obre`, the wrapper and MCMC samplers.

If `show = TRUE`, the optimal B robust estimated weights for the largest observations are printed alongside with the p -value of the latter, obtained from the empirical distribution of the weights. This diagnostic can be used to guide threshold selection: small weights for the r -largest order statistics indicate that the robust fit is driven by the lower tail and that the threshold should perhaps be increased.

References

- Davison, A.C. (1984). Modelling excesses over high thresholds, with an application, in *Statistical extremes and applications*, J. Tiago de Oliveira (editor), D. Reidel Publishing Co., 461–482.
- Grimshaw, S.D. (1993). Computing Maximum Likelihood Estimates for the Generalized Pareto Distribution, *Technometrics*, **35**(2), 185–191.
- Northrop, P.J. and C. L. Coleman (2014). Improved threshold diagnostic plots for extreme value analyses, *Extremes*, **17**(2), 289–303.
- Zhang, J. (2010). Improving on estimation for the generalized Pareto distribution, *Technometrics* **52**(3), 335–339.

Zhang, J. and M. A. Stephens (2009). A new and efficient estimation method for the generalized Pareto distribution. *Technometrics* **51**(3), 316–325.

Dupuis, D.J. (1998). Exceedances over High Thresholds: A Guide to Threshold Selection, *Extremes*, **1**(3), 251–261.

See Also

[fpot](#) and [gpd.fit](#)

Examples

```
data(eskrain)
fit.gpd(eskrain, threshold = 35, method = 'Grimshaw', show = TRUE)
fit.gpd(eskrain, threshold = 30, method = 'zs', show = TRUE)
```

fit.pp

Maximum likelihood estimation of the point process of extremes

Description

Data above threshold is modelled using the limiting point process of extremes.

Usage

```
fit.pp(
  xdat,
  threshold = 0,
  npp = 1,
  np = NULL,
  method = c("nlminb", "BFGS"),
  start = NULL,
  show = FALSE,
  fpar = NULL,
  warnSE = FALSE
)
```

Arguments

| | |
|-----------|---|
| xdat | a numeric vector of data to be fitted. |
| threshold | the chosen threshold. |
| npp | number of observation per period. See Details |
| np | number of periods of data, if xdat only contains exceedances. |
| method | the method to be used. See Details . Can be abbreviated. |
| start | named list of starting values |
| show | logical; if TRUE (the default), print details of the fit. |
| fpar | a named list with optional fixed components loc, scale and shape |
| warnSE | logical; if TRUE, a warning is printed if the standard errors cannot be returned from the observed information matrix when the shape is less than -0.5. |

Details

The parameter `npp` controls the frequency of observations. If data are recorded on a daily basis, using a value of `npp = 365.25` yields location and scale parameters that correspond to those of the generalized extreme value distribution fitted to block maxima.

Value

a list containing the following components:

- `estimate` a vector containing all parameters (optimized and fixed).
- `std.err` a vector containing the standard errors.
- `vcov` the variance covariance matrix, obtained as the numerical inverse of the observed information matrix.
- `threshold` the threshold.
- `method` the method used to fit the parameter. See details.
- `nllh` the negative log-likelihood evaluated at the parameter estimate.
- `nat` number of points lying above the threshold.
- `pat` proportion of points lying above the threshold.
- `convergence` components taken from the list returned by `optim`. Values other than 0 indicate that the algorithm likely did not converge (in particular 1 and 50).
- `counts` components taken from the list returned by `optim`.

References

Coles, S. (2001), An introduction to statistical modelling of extreme values. Springer : London, 208p.

Examples

```
data(eskrain)
pp_mle <- fit.pp(eskrain, threshold = 30, np = 6201)
plot(pp_mle)
```

fit.rho

Estimator of the second order tail index parameter

Description

Estimator of the second order tail index parameter

Usage

```
fit.rho(xdat, k, method = c("fagh", "dk", "ghp", "gbw"), ...)
```

Arguments

| | |
|--------|---|
| xdat | vector of positive observations |
| k | number of highest order statistics to use for estimation |
| method | string for the estimator |
| ... | additional arguments passed to individual routines currently ignored. |

Examples

```
# Example with rho = -0.2
n <- 1000
xdat <- mev::rgp(n = n, shape = 0.2)
kmin <- floor(n^0.995)
kmax <- ceiling(n^0.999)
rho_est <- fit.rho(
  xdat = xdat,
  k = n - kmin:kmax)
rho_med <- mean(rho_est$rho)
```

| | |
|-----------|---|
| fit.rlarg | <i>Maximum likelihood estimates of point process for the r-largest observations</i> |
|-----------|---|

Description

This uses a constrained optimization routine to return the maximum likelihood estimate based on an n by r matrix of observations. Observations should be ordered, i.e., the r -largest should be in the last column.

Usage

```
fit.rlarg(
  xdat,
  start = NULL,
  method = c("nlminb", "BFGS"),
  show = FALSE,
  fpar = NULL,
  warnSE = FALSE
)
```

Arguments

| | |
|--------|---|
| xdat | a matrix of size n by r |
| start | named list of starting values |
| method | the method to be used. See Details . Can be abbreviated. |
| show | logical; if TRUE (the default), print details of the fit. |
| fpar | a named list with fixed parameters, either scale or shape |
| warnSE | logical; if TRUE, a warning is printed if the standard errors cannot be returned from the observed information matrix when the shape is less than -0.5. |

Value

a list containing the following components:

- `estimate` a vector containing all the maximum likelihood estimates.
- `std.err` a vector containing the standard errors.
- `vcov` the variance covariance matrix, obtained as the numerical inverse of the observed information matrix.
- `method` the method used to fit the parameter.
- `nllh` the negative log-likelihood evaluated at the parameter estimate.
- `convergence` components taken from the list returned by `auglag`. Values other than 0 indicate that the algorithm likely did not converge.
- `counts` components taken from the list returned by `auglag`.
- `xdat` an n by r matrix of data

Examples

```
xdat <- rrlarg(n = 10, loc = 0, scale = 1, shape = 0.1, r = 4)
fit.rlarg(xdat)
```

fit.shape

Shape parameter estimates

Description

Wrapper to estimate the tail index or shape parameter of an extreme value distribution. Each function has similar sets of arguments, a vector or scalar number of order statistics k and a vector of positive observations `xdat`. The `method` argument allows users to choose between different indicators, including the Hill estimator (`hill`, for positive observations and shape only), the moment estimator of Dekkers and de Haan (`mom` or `dekkers`), the de Vries estimator of de Haan and Peng (`vries`), the generalized jackknife estimator of Gomes et al. (`genjack`), the Beirlant, Vynckier and Teugels generalized quantile estimator (`bvt` or `genquant`), the Pickands estimator (`pickands`), the extreme U -statistics estimator of Oorschot, Segers and Zhou (`osz`), or the exponential regression model of Beirlant et al. (`erm`).

Usage

```
fit.shape(
  xdat,
  k,
  method = c("hill", "rbm", "osz", "vries", "genjack", "mom", "dekkers", "genquant",
            "pickands", "erm"),
  ...
)
```

Arguments

| | |
|--------|---|
| xdat | vector of positive observations of length n |
| k | number of largest order statistics |
| method | estimation method. |
| ... | additional parameters passed to functions |

Value

a data frame with the number of order statistics k and the shape parameter estimate `shape`, or a single numeric value if k is a scalar.

| | |
|----------|---|
| fit.wgpd | <i>Maximum likelihood estimation for weighted generalized Pareto distribution</i> |
|----------|---|

Description

Weighted maximum likelihood estimation, with user-specified vector of weights.

Usage

```
fit.wgpd(xdat, threshold = 0, weightfun = Stein_weights, start = NULL, ...)
```

Arguments

| | |
|-----------|--|
| xdat | vector of observations |
| threshold | numeric, value of the threshold |
| weightfun | function whose first argument is the length of the weight vector |
| start | optional vector of scale and shape parameters for the optimization routine, defaults to NULL |
| ... | additional arguments passed to the weighting function <code>weightfun</code> |

Value

a list with components

- `estimate` a vector containing the scale and shape parameters (optimized and fixed).
- `std.err` a vector containing the standard errors.
- `vcov` the variance covariance matrix, obtained as the numerical inverse of the observed information matrix.
- `threshold` the threshold.
- `method` the method used to fit the parameter. See details.
- `nllh` the negative log-likelihood evaluated at the parameter estimate.
- `nat` number of points lying above the threshold.

- pat proportion of points lying above the threshold.
- convergence logical indicator of convergence.
- weights vector of weights for exceedances.
- exceedances excess over the threshold, sorted in decreasing order.

frwind

French wind data

Description

Daily mean wind speed (in km/h) at four stations in the south of France, namely Cap Cepet (S1), Lyon St-Exupery (S2), Marseille Marignane (S3) and Montelimar (S4). The data includes observations from January 1976 until April 2023; days containing missing values are omitted.

Format

A data frame with 17209 observations and 8 variables:

date date of measurement

S1 wind speed (in km/h) at Cap Cepet

S2 wind speed (in km/h) at Lyon Saint-Exupery

S3 wind speed (in km/h) at Marseille Marignane

S4 wind speed (in km/h) at Montelimar

H2 humidity (in percentage) at Lyon Saint-Exupery

T2 mean temperature (in degree Celcius) at Lyon Saint-Exupery

The metadata attribute includes latitude and longitude (in degrees, minutes, seconds), altitude (in m), station name and station id.

Source

European Climate Assessment and Dataset project <https://www.ecad.eu/>

References

Klein Tank, A.M.G. and Coauthors, 2002. Daily dataset of 20th-century surface air temperature and precipitation series for the European Climate Assessment. *Int. J. of Climatol.*, 22, 1441-1453.

Examples

```
data(frwind, package = "mev")
head(frwind)
attr(frwind, which = "metadata")
```

| | |
|-------------|------------------------|
| geomagnetic | <i>Magnetic storms</i> |
|-------------|------------------------|

Description

Absolute magnitude of 373 geomagnetic storms lasting more than 48h with absolute magnitude (dst) larger than 100 in 1957-2014.

Format

a vector of size 373

Note

For a detailed article presenting the derivation of the Dst index, see <http://wdc.kugi.kyoto-u.ac.jp/dstdir/dst2/onDst>

Source

Aki Vehtari

References

World Data Center for Geomagnetism, Kyoto, M. Nose, T. Iyemori, M. Sugiura, T. Kamei (2015), *Geomagnetic Dst index*, <doi:10.17593/14515-74000>.

| | |
|-----|---|
| gev | <i>Generalized extreme value distribution</i> |
|-----|---|

Description

Likelihood, score function and information matrix, bias, approximate ancillary statistics and sample space derivative for the generalized extreme value distribution

Arguments

| | |
|--------|---|
| par | vector of loc, scale and shape |
| dat | sample vector |
| method | string indicating whether to use the expected ('exp') or the observed ('obs' - the default) information matrix. |
| V | vector calculated by gev.vfun |
| n | sample size |
| p | vector of probabilities |

Usage

```
gev.ll(par, dat)
gev.ll.optim(par, dat)
gev.score(par, dat)
gev.infomat(par, dat, method = c('obs', 'exp'))
gev.retlev(par, p)
gev.bias(par, n)
gev.Fscore(par, dat, method=c('obs', 'exp'))
gev.Vfun(par, dat)
gev.phi(par, dat, V)
gev.dphi(par, dat, V)
```

Functions

- `gev.ll`: log likelihood
- `gev.ll.optim`: negative log likelihood parametrized in terms of location, $\log(\text{scale})$ and shape in order to perform unconstrained optimization
- `gev.score`: score vector
- `gev.infomat`: observed or expected information matrix
- `gev.retlev`: return level, corresponding to the $(1 - p)$ th quantile
- `gev.bias`: Cox-Snell first order bias
- `gev.Fscore`: Firth's modified score equation
- `gev.Vfun`: vector implementing conditioning on approximate ancillary statistics for the TEM
- `gev.phi`: canonical parameter in the local exponential family approximation
- `gev.dphi`: derivative matrix of the canonical parameter in the local exponential family approximation

References

- Firth, D. (1993). Bias reduction of maximum likelihood estimates, *Biometrika*, **80**(1), 27–38.
- Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*, Springer, 209 p.
- Cox, D. R. and E. J. Snell (1968). A general definition of residuals, *Journal of the Royal Statistical Society: Series B (Methodological)*, **30**, 248–275.
- Cordeiro, G. M. and R. Klein (1994). Bias correction in ARMA models, *Statistics and Probability Letters*, **19**(3), 169–176.

 gev.abias

Asymptotic bias of block maxima for fixed sample sizes

Description

Asymptotic bias of block maxima for fixed sample sizes

Usage

```
gev.abias(shape, rho)
```

Arguments

| | |
|-------|--------------------------------------|
| shape | shape parameter |
| rho | second-order parameter, non-positive |

Value

a vector of length three containing the bias for location, scale and shape (in this order)

References

Dombry, C. and A. Ferreira (2017). Maximum likelihood estimators based on the block maxima method. <https://arxiv.org/abs/1705.00465>

 gev.bcor

Bias correction for GEV distribution

Description

Bias corrected estimates for the generalized extreme value distribution using Firth's modified score function or implicit bias subtraction.

Usage

```
gev.bcor(par, dat, corr = c("subtract", "firth"), method = c("obs", "exp"))
```

Arguments

| | |
|--------|---|
| par | parameter vector (scale, shape) |
| dat | sample of observations |
| corr | string indicating which correction to employ either subtract or firth |
| method | string indicating whether to use the expected ('exp') or the observed ('obs' — the default) information matrix. Used only if corr='firth' |

Details

Method subtractsolves

$$\tilde{\theta} = \hat{\theta} + b(\tilde{\theta})$$

for $\tilde{\theta}$, using the first order term in the bias expansion as given by `gev.bias`.

The alternative is to use Firth's modified score and find the root of

$$U(\tilde{\theta}) - i(\tilde{\theta})b(\tilde{\theta}),$$

where U is the score vector, b is the first order bias and i is either the observed or Fisher information.

The routine uses the MLE (bias-corrected) as starting values and proceeds to find the solution using a root finding algorithm. Since the bias-correction is not valid for $\xi < -1/3$, any solution that is unbounded will return a vector of NA as the solution does not exist then.

Value

vector of bias-corrected parameters

Examples

```
set.seed(1)
dat <- mev::rgev(n=40, loc = 1, scale=1, shape=-0.2)
par <- mev::fit.gev(dat)$estimate
gev.bcor(par, dat, 'subtract')
gev.bcor(par, dat, 'firth') #observed information
gev.bcor(par, dat, 'firth','exp')
```

gev.boot

Bootstrap approximation for generalized extreme value parameters

Description

Given an object of class `mev_gev`, returns a matrix of parameter values to mimic the estimation uncertainty.

Usage

```
gev.boot(object, B = 1000L, method = c("post", "norm"))
```

Arguments

| | |
|--------|---|
| object | object of class <code>mev_gev</code> |
| B | number of pairs to sample |
| method | string; one of 'norm' for the normal approximation or 'post' (default) for posterior sampling |

Details

Two options are available: a normal approximation to the location, scale and shape based on the maximum likelihood estimates and the observed information matrix. This method uses forward sampling to simulate from a trivariate normal distribution that satisfies the support and positivity constraints

The second approximation uses the ratio-of-uniforms method to obtain samples from the posterior distribution with uninformative priors, thus mimicking the joint distribution of maximum likelihood. The benefit of the latter is that it is more reliable in small samples and when the shape is negative.

Value

a matrix of size B by 3 whose columns contain scale and shape parameters

Examples

```
set.seed(2025)
xdat <- rgev(100, loc = 0, scale = 2, shape = -0.1)
fgev <- fit.gev(xdat)
pairs(gev.boot(fgev, method = "post"))
pairs(gev.boot(fgev, method = "norm"))
```

| | |
|---------|--|
| gev.mle | <i>Generalized extreme value maximum likelihood estimates for various quantities of interest</i> |
|---------|--|

Description

This function calls the `fit.gev` routine on the sample of block maxima and returns maximum likelihood estimates for all quantities of interest, including location, scale and shape parameters, quantiles and mean and quantiles of maxima of N blocks.

Usage

```
gev.mle(
  xdat,
  args = c("loc", "scale", "shape", "quant", "Nmean", "Nquant"),
  N,
  p,
  q
)
```

Arguments

| | |
|------|---|
| xdat | sample vector of maxima |
| args | vector of strings indicating which arguments to return the maximum likelihood values for. |

| | |
|---|---|
| N | size of block over which to take maxima. Required only for args Nmean and Nquant. |
| p | tail probability. Required only for arg quant. |
| q | level of quantile for maxima of N exceedances. Required only for args Nquant. |

Value

named vector with maximum likelihood estimated parameter values for arguments args

Examples

```
dat <- mev::rgev(n = 100, shape = 0.2)
gev.mle(xdat = dat, N = 100, p = 0.01, q = 0.5)
```

| | |
|---------|---|
| gev.Nyr | <i>N-year return levels, median and mean estimate</i> |
|---------|---|

Description

N-year return levels, median and mean estimate

Usage

```
gev.Nyr(par, nobs, N, type = c("retlev", "median", "mean"), p = 1/N)
```

Arguments

| | |
|------|---|
| par | vector of location, scale and shape parameters for the GEV distribution |
| nobs | integer number of observation on which the fit is based |
| N | integer number of observations for return level. See Details |
| type | string indicating the statistic to be calculated (can be abbreviated). |
| p | probability indicating the return level, corresponding to the quantile at $1-1/p$ |

Details

If there are n_y observations per year, the L-year return level is obtained by taking N equal to $n_y L$.

Value

a list with components

- est point estimate
- var variance estimate based on delta-method
- type statistic

 gev.pll

Profile log-likelihood for the generalized extreme value distribution

Description

This function calculates the profile likelihood along with two small-sample corrections based on Severini's (1999) empirical covariance and the Fraser and Reid tangent exponential model approximation.

Usage

```
gev.pll(
  psi,
  param = c("loc", "scale", "shape", "quant", "Nmean", "Nquant"),
  mod = "profile",
  dat,
  N = NULL,
  p = NULL,
  q = NULL,
  correction = TRUE,
  plot = TRUE,
  ...
)
```

Arguments

| | |
|------------|--|
| psi | parameter vector over which to profile (unidimensional) |
| param | string indicating the parameter to profile over |
| mod | string indicating the model, one of profile, tem or modif. See Details . |
| dat | sample vector |
| N | size of block over which to take maxima. Required only for param Nmean and Nquant. |
| p | tail probability. Required only for param quant. |
| q | probability level of quantile. Required only for param Nquant. |
| correction | logical indicating whether to use spline.corr to smooth the tem approximation. |
| plot | logical; should the profile likelihood be displayed? Default to TRUE |
| ... | additional arguments such as output from call to Vfun if mode='tem'. |

Details

The two additional mod available are tem, the tangent exponential model (TEM) approximation and modif for the penalized profile likelihood based on p^* approximation proposed by Severini. For the latter, the penalization is based on the TEM or an empirical covariance adjustment term.

Value

a list with components

- `mle`: maximum likelihood estimate
- `psi.max`: maximum profile likelihood estimate
- `param`: string indicating the parameter to profile over
- `std.error`: standard error of `psi.max`
- `psi`: vector of parameter ψ given in `psi`
- `pll`: values of the profile log likelihood at `psi`
- `maxpll`: value of maximum profile log likelihood

In addition, if `mod` includes `tem`

- `normal`: maximum likelihood estimate and standard error of the interest parameter ψ
- `r`: values of likelihood root corresponding to ψ
- `q`: vector of likelihood modifications
- `rstar`: modified likelihood root vector
- `rstar.old`: uncorrected modified likelihood root vector
- `tem.psimax`: maximum of the tangent exponential model likelihood

In addition, if `mod` includes `modif`

- `tem.mle`: maximum of tangent exponential modified profile log likelihood
- `tem.profl1`: values of the modified profile log likelihood at `psi`
- `tem.maxpll`: value of maximum modified profile log likelihood
- `empcov.mle`: maximum of Severini's empirical covariance modified profile log likelihood
- `empcov.profl1`: values of the modified profile log likelihood at `psi`
- `empcov.maxpll`: value of maximum modified profile log likelihood

References

Fraser, D. A. S., Reid, N. and Wu, J. (1999), A simple general formula for tail probabilities for frequentist and Bayesian inference. *Biometrika*, **86**(2), 249–264.

Severini, T. (2000) Likelihood Methods in Statistics. Oxford University Press. ISBN 9780198506508.

Brazzale, A. R., Davison, A. C. and Reid, N. (2007) Applied asymptotics: case studies in small-sample statistics. Cambridge University Press, Cambridge. ISBN 978-0-521-84703-2

Examples

```
## Not run:
set.seed(123)
dat <- rgev(n = 100, loc = 0, scale = 2, shape = 0.3)
gev.pll(psi = seq(0,0.5, length = 50), param = 'shape', dat = dat)
gev.pll(psi = seq(-1.5, 1.5, length = 50), param = 'loc', dat = dat)
gev.pll(psi = seq(10, 40, length = 50), param = 'quant', dat = dat, p = 0.01)
```

```

gev.pll(psi = seq(12, 100, length = 50), param = 'Nmean', N = 100, dat = dat)
gev.pll(psi = seq(12, 90, length = 50), param = 'Nquant', N = 100, dat = dat, q = 0.5)

## End(Not run)

```

 gev.tem

Tangent exponential model approximation for the GEV distribution

Description

The function `gev.tem` provides a tangent exponential model (TEM) approximation for higher order likelihood inference for a scalar parameter for the generalized extreme value distribution. Options include location scale and shape parameters as well as value-at-risk (or return levels). The function attempts to find good values for `psi` that will cover the range of options, but the fail may fit and return an error.

Usage

```

gev.tem(
  param = c("loc", "scale", "shape", "quant", "Nmean", "Nquant"),
  dat,
  psi = NULL,
  p = NULL,
  q = 0.5,
  N = NULL,
  n.psi = 50,
  plot = TRUE,
  correction = TRUE
)

```

Arguments

| | |
|-------------------------|---|
| <code>param</code> | parameter over which to profile |
| <code>dat</code> | sample vector for the GEV distribution |
| <code>psi</code> | scalar or ordered vector of values for the interest parameter. If <code>NULL</code> (default), a grid of values centered at the MLE is selected |
| <code>p</code> | tail probability for the (1-p)th quantile (return levels). Required only if <code>param = 'retlev'</code> |
| <code>q</code> | probability level of quantile. Required only for <code>param Nquant</code> . |
| <code>N</code> | size of block over which to take maxima. Required only for <code>param Nmean</code> and <code>Nquant</code> . |
| <code>n.psi</code> | number of values of <code>psi</code> at which the likelihood is computed, if <code>psi</code> is not supplied (<code>NULL</code>). Odd values are more prone to give rise to numerical instabilities near the MLE. If <code>psi</code> is a vector of length 2 and <code>n.psi</code> is greater than 2, these are taken to be endpoints of the sequence. |
| <code>plot</code> | logical indicating whether <code>plot.fr</code> should be called upon exit |
| <code>correction</code> | logical indicating whether spline.corr should be called. |

Value

an invisible object of class `fr` (see `tem` in package `hoa`) with elements

- `normal`: maximum likelihood estimate and standard error of the interest parameter ψ
- `par.hat`: maximum likelihood estimates
- `par.hat.se`: standard errors of maximum likelihood estimates
- `th.rest`: estimated maximum profile likelihood at $(\psi, \hat{\lambda})$
- `r`: values of likelihood root corresponding to ψ
- `psi`: vector of interest parameter
- `q`: vector of likelihood modifications
- `rstar`: modified likelihood root vector
- `rstar.old`: uncorrected modified likelihood root vector
- `param`: parameter

Author(s)

Leo Belzile

Examples

```
## Not run:
set.seed(1234)
dat <- rgev(n = 40, loc = 0, scale = 2, shape = -0.1)
gev.tem('shape', dat = dat, plot = TRUE)
gev.tem('quant', dat = dat, p = 0.01, plot = TRUE)
gev.tem('scale', psi = seq(1, 4, by = 0.1), dat = dat, plot = TRUE)
dat <- rgev(n = 40, loc = 0, scale = 2, shape = 0.2)
gev.tem('loc', dat = dat, plot = TRUE)
gev.tem('Nmean', dat = dat, p = 0.01, N=100, plot = TRUE)
gev.tem('Nquant', dat = dat, q = 0.5, N=100, plot = TRUE)

## End(Not run)
```

gevdist

Generalized extreme value distribution

Description

Density function, distribution function, quantile function and random number generation for the generalized extreme value distribution.

Usage

```
qgev(p, loc = 0, scale = 1, shape = 0, lower.tail = TRUE, log.p = FALSE)
```

```
rgev(n, loc = 0, scale = 1, shape = 0)
```

```
dgev(x, loc = 0, scale = 1, shape = 0, log = FALSE)
```

```
pgev(q, loc = 0, scale = 1, shape = 0, lower.tail = TRUE, log.p = FALSE)
```

Arguments

| | |
|------------|--|
| p | vector of probabilities |
| loc | scalar or vector of location parameters whose length matches that of the input |
| scale | scalar or vector of positive scale parameters whose length matches that of the input |
| shape | scalar shape parameter |
| lower.tail | logical; if TRUE (default), returns the distribution function, otherwise the survival function |
| n | scalar number of observations |
| x, q | vector of quantiles |
| log, log.p | logical; if TRUE, probabilities p are given as $\log(p)$. |

Details

The distribution function of a GEV distribution with parameters $\text{loc} = \mu$, $\text{scale} = \sigma$ and $\text{shape} = \xi$ is

$$F(x) = \exp\{-[1 + \xi(x - \mu)/\sigma]^{-1/\xi}\}$$

for $1 + \xi(x - \mu)/\sigma > 0$. If $\xi = 0$ the distribution function is defined as the limit as ξ tends to zero.

The quantile function, when evaluated at zero or one, returns the lower and upper endpoint, whether the latter is finite or not.

Author(s)

Leo Belzile, with code adapted from Paul Northrop

References

- Jenkinson, A. F. (1955) The frequency distribution of the annual maximum (or minimum) of meteorological elements. *Quart. J. R. Met. Soc.*, **81**, 158-171. Chapter 3: [doi:10.1002/qj.49708134804](https://doi.org/10.1002/qj.49708134804)
- Coles, S. G. (2001) *An Introduction to Statistical Modeling of Extreme Values*, Springer-Verlag, London. [doi:10.1007/9781447136750_3](https://doi.org/10.1007/9781447136750_3)

| | |
|------|---|
| gevN | <i>Generalized extreme value distribution (quantile/mean of N-block maxima parametrization)</i> |
|------|---|

Description

Likelihood, score function and information matrix, approximate ancillary statistics and sample space derivative for the generalized extreme value distribution parametrized in terms of the quantiles/mean of N-block maxima parametrization z , scale and shape.

Arguments

| | |
|-----|--|
| par | vector of loc, quantile/mean of N-block maximum and shape |
| dat | sample vector |
| V | vector calculated by gevN.Vfun |
| q | probability, corresponding to q th quantile of the N-block maximum |
| qty | string indicating whether to calculate the q quantile or the mean |

Usage

```

gevN.ll(par, dat, N, q, qty = c('mean', 'quantile'))
gevN.ll.optim(par, dat, N, q = 0.5, qty = c('mean', 'quantile'))
gevN.score(par, dat, N, q = 0.5, qty = c('mean', 'quantile'))
gevN.infomat(par, dat, qty = c('mean', 'quantile'), method = c('obs', 'exp'), N, q = 0.5, nobs = length(dat))
gevN.Vfun(par, dat, N, q = 0.5, qty = c('mean', 'quantile'))
gevN.phi(par, dat, N, q = 0.5, qty = c('mean', 'quantile'), V)
gevN.dphi(par, dat, N, q = 0.5, qty = c('mean', 'quantile'), V)

```

Functions

- gevN.ll: log likelihood
- gevN.score: score vector
- gevN.infomat: expected and observed information matrix
- gevN.Vfun: vector implementing conditioning on approximate ancillary statistics for the TEM
- gevN.phi: canonical parameter in the local exponential family approximation
- gevN.dphi: derivative matrix of the canonical parameter in the local exponential family approximation

Author(s)

Leo Belzile

gevr *Generalized extreme value distribution (return level parametrization)*

Description

Likelihood, score function and information matrix, approximate ancillary statistics and sample space derivative for the generalized extreme value distribution parametrized in terms of the return level z , scale and shape.

Arguments

| | |
|--------|---|
| par | vector of retlev, scale and shape |
| dat | sample vector |
| p | tail probability, corresponding to $(1 - p)$ th quantile for z |
| method | string indicating whether to use the expected ('exp') or the observed ('obs' - the default) information matrix. |
| nobs | number of observations |
| V | vector calculated by <code>gevr.Vfun</code> |

Usage

```
gevr.ll(par, dat, p)
gevr.ll.optim(par, dat, p)
gevr.score(par, dat, p)
gevr.infomat(par, dat, p, method = c('obs', 'exp'), nobs = length(dat))
gevr.Vfun(par, dat, p)
gevr.phi(par, dat, p, V)
gevr.dphi(par, dat, p, V)
```

Functions

- `gevr.ll`: log likelihood
- `gevr.ll.optim`: negative log likelihood parametrized in terms of return levels, $\log(\text{scale})$ and shape in order to perform unconstrained optimization
- `gevr.score`: score vector
- `gevr.infomat`: observed information matrix
- `gevr.Vfun`: vector implementing conditioning on approximate ancillary statistics for the TEM
- `gevr.phi`: canonical parameter in the local exponential family approximation
- `gevr.dphi`: derivative matrix of the canonical parameter in the local exponential family approximation

Author(s)

Leo Belzile

Author(s)

Leo Belzile

References

- Firth, D. (1993). Bias reduction of maximum likelihood estimates, *Biometrika*, **80**(1), 27–38.
- Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*, Springer, 209 p.
- Cox, D. R. and E. J. Snell (1968). A general definition of residuals, *Journal of the Royal Statistical Society: Series B (Methodological)*, **30**, 248–275.
- Cordeiro, G. M. and R. Klein (1994). Bias correction in ARMA models, *Statistics and Probability Letters*, **19**(3), 169–176.
- Giles, D. E., Feng, H. and R. T. Godwin (2016). Bias-corrected maximum likelihood estimation of the parameters of the generalized Pareto distribution, *Communications in Statistics - Theory and Methods*, **45**(8), 2465–2483.

gpd.abias

*Asymptotic bias of threshold exceedances for k order statistics***Description**

The formula given in de Haan and Ferreira, 2007 (Springer). Note that the latter differs from that found in Drees, Ferreira and de Haan.

Usage

```
gpd.abias(shape, rho)
```

Arguments

| | |
|-------|--------------------------------------|
| shape | shape parameter |
| rho | second-order parameter, non-positive |

Value

a vector of length containing the bias for scale and shape (in this order)

References

- Dombry, C. and A. Ferreira (2017). Maximum likelihood estimators based on the block maxima method. <https://arxiv.org/abs/1705.00465>

gpd.bcor

*Bias correction for GP distribution***Description**

Bias corrected estimates for the generalized Pareto distribution using Firth's modified score function or implicit bias subtraction.

Usage

```
gpd.bcor(par, dat, corr = c("subtract", "firth"), method = c("obs", "exp"))
```

Arguments

| | |
|--------|---|
| par | parameter vector (scale, shape) |
| dat | sample of observations |
| corr | string indicating which correction to employ either subtract or firth |
| method | string indicating whether to use the expected ('exp') or the observed ('obs' — the default) information matrix. Used only if corr='firth' |

Details

Method subtract solves

$$\tilde{\theta} = \hat{\theta} + b(\tilde{\theta})$$

for $\tilde{\theta}$, using the first order term in the bias expansion as given by [gpd.bias](#).

The alternative is to use Firth's modified score and find the root of

$$U(\tilde{\theta}) - i(\tilde{\theta})b(\tilde{\theta}),$$

where U is the score vector, b is the first order bias and i is either the observed or Fisher information.

The routine uses the MLE as starting value and proceeds to find the solution using a root finding algorithm. Since the bias-correction is not valid for $\xi < -1/3$, any solution that is unbounded will return a vector of NA as the bias correction does not exist then.

Value

vector of bias-corrected parameters

Examples

```
set.seed(1)
dat <- rgp(n=40, scale=1, shape=-0.2)
par <- gp.fit(dat, threshold=0, show=FALSE)$estimate
gpd.bcor(par,dat, 'subtract')
gpd.bcor(par,dat, 'firth') #observed information
gpd.bcor(par,dat, 'firth','exp')
```

`gpd.boot`*Bootstrap approximation for generalized Pareto parameters*

Description

Given an object of class `mev_gpd`, returns a matrix of parameter values to mimic the estimation uncertainty.

Usage

```
gpd.boot(object, B = 1000L, method = c("post", "norm"))
```

Arguments

| | |
|---------------------|---|
| <code>object</code> | object of class <code>mev_gpd</code> |
| <code>B</code> | number of pairs to sample |
| <code>method</code> | string; one of 'norm' for the normal approximation or 'post' (default) for posterior sampling |

Details

Two options are available: a normal approximation to the scale and shape based on the maximum likelihood estimates and the observed information matrix. This method uses forward sampling to simulate from a bivariate normal distribution that satisfies the support and positivity constraints

The second approximation uses the ratio-of-uniforms method to obtain samples from the posterior distribution with uninformative priors, thus mimicking the joint distribution of maximum likelihood. The benefit of the latter is that it is more reliable in small samples and when the shape is negative.

Value

a matrix of size `B` by 2 whose columns contain scale and shape parameters

Examples

```
set.seed(2025)
xdat <- rgev(100, loc = 0, scale = 2, shape = -0.1)
fgp <- fit.gpd(xdat)
plot(
  gpd.boot(fgp, method = "post")
)
points(
  gpd.boot(fgp, method = "norm"),
  col = 2,
  pch = 20
)
```

gpd.lmom

Estimation of generalized Pareto parameters via L-moments

Description

Given a sample of exceedances, compute the first four L-moments and use either the first two to obtain the scale and shape (default), or else use L-skewness and L-scale to compute the scale and shape of the generalized Pareto distribution

Usage

```
gpd.lmom(xdat, thresh, sorted = FALSE, Lskew = FALSE)
```

Arguments

| | |
|--------|--|
| xdat | [numeric] vector of observations |
| thresh | [numeric] optional threshold argument |
| sorted | [logical] if TRUE, observations are sorted in increasing order |
| Lskew | [logical]; if TRUE, shape is obtained from L-skewness rather than first two moments. |

Value

a vector of length two with the scale and shape estimates

gpd.mle

Generalized Pareto maximum likelihood estimates for various quantities of interest

Description

This function calls the `fit.gpd` routine on the sample of excesses and returns maximum likelihood estimates for all quantities of interest, including scale and shape parameters, quantiles and value-at-risk, expected shortfall and mean and quantiles of maxima of N threshold exceedances

Usage

```
gpd.mle(
  xdat,
  args = c("scale", "shape", "quant", "VaR", "ES", "Nmean", "Nquant"),
  m,
  N,
  p,
  q
)
```

Arguments

| | |
|------|---|
| xdat | sample vector of excesses |
| args | vector of strings indicating which arguments to return the maximum likelihood values for |
| m | number of observations of interest for return levels. Required only for args values 'VaR' or 'ES' |
| N | size of block over which to take maxima. Required only for args Nmean and Nquant. |
| p | tail probability, equivalent to $1/m$. Required only for args quant. |
| q | level of quantile for N-block maxima. Required only for args Nquant. |

Value

named vector with maximum likelihood values for arguments args

Examples

```
xdat <- mev::rgp(n = 30, shape = 0.2)
gpd.mle(xdat = xdat, N = 100, p = 0.01, q = 0.5, m = 100)
```

gpd.pll

Profile log-likelihood for the generalized Pareto distribution

Description

This function calculates the (modified) profile likelihood based on the p^* formula. There are two small-sample corrections that use a proxy for $\ell_{\lambda; \hat{\lambda}}$, which are based on Severini's (1999) empirical covariance and the Fraser and Reid tangent exponential model approximation.

Usage

```
gpd.pll(
  psi,
  param = c("scale", "shape", "quant", "retlev", "VaR", "ES", "Nmean", "Nquant"),
  mod = "profile",
  mle = NULL,
  dat,
  m = NULL,
  N = NULL,
  p = NULL,
  q = NULL,
  correction = TRUE,
  thresh = NULL,
  plot = TRUE,
  ...
)
```

Arguments

| | |
|------------|---|
| psi | parameter vector over which to profile (unidimensional) |
| param | string indicating the parameter to profile over |
| mod | string indicating the model. See Details . |
| mle | maximum likelihood estimate in (ψ, ξ) parametrization if $\psi \neq \xi$ and (σ, ξ) otherwise (optional). |
| dat | sample vector of excesses, unless thresh is provided (in which case user provides original data) |
| m | number of observations of interest for return levels. Required only for args values 'VaR' or 'ES' |
| N | size of block over which to take maxima. Required only for args Nmean and Nquant. |
| p | tail probability, equivalent to $1/m$. Required only for args quant. |
| q | level of quantile for N-block maxima. Required only for args Nquant. |
| correction | logical indicating whether to use spline.corr to smooth the tem approximation. |
| thresh | numerical threshold above which to fit the generalized Pareto distribution |
| plot | logical; should the profile likelihood be displayed? Default to TRUE |
| ... | additional arguments such as output from call to Vfun if mode='tem'. |

Details

The three mod available are profile (the default), tem, the tangent exponential model (TEM) approximation and modif for the penalized profile likelihood based on p^* approximation proposed by Severini. For the latter, the penalization is based on the TEM or an empirical covariance adjustment term.

Value

a list with components

- mle: maximum likelihood estimate
- psi.max: maximum profile likelihood estimate
- param: string indicating the parameter to profile over
- std.error: standard error of psi.max
- psi: vector of parameter ψ given in psi
- pll: values of the profile log likelihood at psi
- maxpll: value of maximum profile log likelihood
- family: a string indicating "gpd"
- thresh: value of the threshold, by default zero

In addition, if mod includes tem

- `normal`: maximum likelihood estimate and standard error of the interest parameter ψ
- `r`: values of likelihood root corresponding to ψ
- `q`: vector of likelihood modifications
- `rstar`: modified likelihood root vector
- `rstar.old`: uncorrected modified likelihood root vector
- `tem.psimax`: maximum of the tangent exponential model likelihood

In addition, if `mod` includes `modif`

- `tem.mle`: maximum of tangent exponential modified profile log likelihood
- `tem.profl1`: values of the modified profile log likelihood at `psi`
- `tem.maxpll`: value of maximum modified profile log likelihood
- `empcov.mle`: maximum of Severini's empirical covariance modified profile log likelihood
- `empcov.profl1`: values of the modified profile log likelihood at `psi`
- `empcov.maxpll`: value of maximum modified profile log likelihood

Examples

```
## Not run:
dat <- rgp(n = 100, scale = 2, shape = 0.3)
gpd.pll(psi = seq(-0.5, 1, by=0.01), param = 'shape', dat = dat)
gpd.pll(psi = seq(0.1, 5, by=0.1), param = 'scale', dat = dat)
gpd.pll(psi = seq(20, 35, by=0.1), param = 'quant', dat = dat, p = 0.01)
gpd.pll(psi = seq(20, 80, by=0.1), param = 'ES', dat = dat, m = 100)
gpd.pll(psi = seq(15, 100, by=1), param = 'Nmean', N = 100, dat = dat)
gpd.pll(psi = seq(15, 90, by=1), param = 'Nquant', N = 100, dat = dat, q = 0.5)

## End(Not run)
```

gpd.tem

Tangent exponential model approximation for the GP distribution

Description

The function `gpd.tem` provides a tangent exponential model (TEM) approximation for higher order likelihood inference for a scalar parameter for the generalized Pareto distribution. Options include scale and shape parameters as well as value-at-risk (also referred to as quantiles, or return levels) and expected shortfall. The function attempts to find good values for `psi` that will cover the range of options, but the fit may fail and return an error. In such cases, the user can try to find good grid of starting values and provide them to the routine.

Usage

```

gpd.tem(
  dat,
  param = c("scale", "shape", "quant", "VaR", "retlev", "ES", "Nmean", "Nquant"),
  psi = NULL,
  m = NULL,
  thresh = 0,
  n.psi = 50,
  N = NULL,
  p = NULL,
  q = NULL,
  plot = FALSE,
  correction = TRUE,
  ...
)

```

Arguments

| | |
|-------------------------|---|
| <code>dat</code> | sample vector for the GP distribution |
| <code>param</code> | parameter over which to profile |
| <code>psi</code> | scalar or ordered vector of values for the interest parameter. If NULL (default), a grid of values centered at the MLE is selected. If <code>psi</code> is of length 2 and <code>n.psi</code> >2, it is assumed to be the minimal and maximal values at which to evaluate the profile log likelihood. |
| <code>m</code> | number of observations of interest for return levels. See Details . Required only for <code>param = 'VaR'</code> or <code>param = 'ES'</code> . |
| <code>thresh</code> | threshold value corresponding to the lower bound of the support or the location parameter of the generalized Pareto distribution. |
| <code>n.psi</code> | number of values of <code>psi</code> at which the likelihood is computed, if <code>psi</code> is not supplied (NULL). Odd values are more prone to give rise to numerical instabilities near the MLE |
| <code>N</code> | size of block over which to take maxima. Required only for args <code>Nmean</code> and <code>Nquant</code> . |
| <code>p</code> | tail probability, equivalent to $1/m$. Required only for args <code>quant</code> . |
| <code>q</code> | level of quantile for N-block maxima. Required only for args <code>Nquant</code> . |
| <code>plot</code> | logical indicating whether <code>plot.fr</code> should be called upon exit |
| <code>correction</code> | logical indicating whether spline.corr should be called. |
| <code>...</code> | additional arguments, for backward compatibility |

Details

As of version 1.11, this function is a wrapper around `gpd.pll`.

The interpretation for `m` is as follows: if there are on average m_y observations per year above the threshold, then $m = Tm_y$ corresponds to T -year return level.

Value

an invisible object of class `fr` (see `tem` in package `hoa`) with elements

- `normal`: maximum likelihood estimate and standard error of the interest parameter ψ
- `par.hat`: maximum likelihood estimates
- `par.hat.se`: standard errors of maximum likelihood estimates
- `th.rest`: estimated maximum profile likelihood at $(\psi, \hat{\lambda})$
- `r`: values of likelihood root corresponding to ψ
- `psi`: vector of interest parameter
- `q`: vector of likelihood modifications
- `rstar`: modified likelihood root vector
- `rstar.old`: uncorrected modified likelihood root vector
- `param`: parameter

Author(s)

Leo Belzile

Examples

```
set.seed(123)
dat <- rgp(n = 40, scale = 1, shape = -0.1)
#with plots
m1 <- gpd.tem(param = 'shape', n.psi = 50, dat = dat, plot = TRUE)
## Not run:
m2 <- gpd.tem(param = 'scale', n.psi = 50, dat = dat)
m3 <- gpd.tem(param = 'VaR', n.psi = 50, dat = dat, m = 100)
#Providing psi
psi <- c(seq(2, 5, length = 15), seq(5, 35, length = 45))
m4 <- gpd.tem(param = 'ES', dat = dat, m = 100, psi = psi, correction = FALSE)
mev:::plot.fr(m4, which = c(2, 4))
plot(fr4 <- spline.corr(m4))
confint(m1)
confint(m4, parm = 2, warn = FALSE)
m5 <- gpd.tem(param = 'Nmean', dat = dat, N = 100, psi = psi, correction = FALSE)
m6 <- gpd.tem(param = 'Nquant', dat = dat, N = 100, q = 0.7, correction = FALSE)

## End(Not run)
```


- `gpde.Vfun`: vector implementing conditioning on approximate ancillary statistics for the TEM
- `gpde.phi`: canonical parameter in the local exponential family approximation
- `gpde.dphi`: derivative matrix of the canonical parameter in the local exponential family approximation

Author(s)

Leo Belzile

gpdist

Generalized Pareto distribution

Description

Density function, distribution function, quantile function and random number generation for the generalized Pareto distribution.

Usage

```
pgp(q, loc = 0, scale = 1, shape = 0, lower.tail = TRUE, log.p = FALSE)
```

```
dgp(x, loc = 0, scale = 1, shape = 0, log = FALSE)
```

```
qgp(p, loc = 0, scale = 1, shape = 0, lower.tail = TRUE)
```

```
rgp(n, loc = 0, scale = 1, shape = 0)
```

Arguments

| | |
|-------------------------|---|
| <code>loc</code> | location parameter. |
| <code>scale</code> | scale parameter, strictly positive. |
| <code>shape</code> | shape parameter. |
| <code>lower.tail</code> | logical; if TRUE (default), the lower tail probability $\Pr(X \leq x)$ is returned. |
| <code>log.p, log</code> | logical; if FALSE (default), values are returned on the probability scale. |
| <code>x, q</code> | vector of quantiles |
| <code>p</code> | vector of probabilities |
| <code>n</code> | scalar number of observations |

References

Coles, S. G. (2001) *An Introduction to Statistical Modeling of Extreme Values*, Springer-Verlag, London. doi:[10.1007/9781447136750_3](https://doi.org/10.1007/9781447136750_3)

| | |
|------|---|
| gpdN | <i>Generalized Pareto distribution (mean of maximum of N exceedances parametrization)</i> |
|------|---|

Description

Likelihood, score function and information matrix, approximate ancillary statistics and sample space derivative for the generalized Pareto distribution parametrized in terms of average maximum of N exceedances.

The parameter N corresponds to the number of threshold exceedances of interest over which the maxima is taken. z is the corresponding expected value of this block maxima. Note that the actual parametrization is in terms of excess expected mean, meaning expected mean minus threshold.

Arguments

| | |
|-----|--|
| par | vector of length 2 containing z and ξ , respectively the mean excess of the maxima of N exceedances above the threshold and the shape parameter. |
| dat | sample vector |
| N | block size for threshold exceedances. |
| tol | numerical tolerance for the exponential model |
| V | vector calculated by <code>gpdN.Vfun</code> |

Details

The observed information matrix was calculated from the Hessian using symbolic calculus in Sage.

Usage

```
gpdN.ll(par, dat, N, tol=1e-5)
gpdN.score(par, dat, N)
gpdN.infomat(par, dat, N, method = c('obs', 'exp'), nobs = length(dat))
gpdN.Vfun(par, dat, N)
gpdN.phi(par, dat, N, V)
gpdN.dphi(par, dat, N, V)
```

Functions

- `gpdN.ll`: log likelihood
- `gpdN.score`: score vector
- `gpdN.infomat`: observed information matrix for GP parametrized in terms of mean of the maximum of N exceedances and shape
- `gpdN.Vfun`: vector implementing conditioning on approximate ancillary statistics for the TEM
- `gpdN.phi`: canonical parameter in the local exponential family approximation
- `gpdN.dphi`: derivative matrix of the canonical parameter in the local exponential family approximation

Author(s)

Leo Belzile

gpdr

*Generalized Pareto distribution (return level parametrization)***Description**

Likelihood, score function and information matrix, approximate ancillary statistics and sample space derivative for the generalized Pareto distribution parametrized in terms of return levels.

Arguments

| | |
|--------|--|
| par | vector of length 2 containing y_m and ξ , respectively the m -year return level and the shape parameter. |
| dat | sample vector |
| m | number of observations of interest for return levels. See Details |
| tol | numerical tolerance for the exponential model |
| method | string indicating whether to use the expected ('exp') or the observed ('obs' - the default) information matrix. |
| nobs | number of observations |
| V | vector calculated by <code>gpdr.Vfun</code> |

Details

The observed information matrix was calculated from the Hessian using symbolic calculus in Sage.

The interpretation for m is as follows: if there are on average m_y observations per year above the threshold, then $m = Tm_y$ corresponds to T -year return level.

Usage

```
gpdr.ll(par, dat, m, tol=1e-5)
gpdr.ll.optim(par, dat, m, tol=1e-5)
gpdr.score(par, dat, m)
gpdr.infomat(par, dat, m, method = c('obs', 'exp'), nobs = length(dat))
gpdr.Vfun(par, dat, m)
gpdr.phi(par, V, dat, m)
gpdr.dphi(par, V, dat, m)
```

Functions

- `gpd. ll`: log likelihood
- `gpd. ll. optim`: negative log likelihood parametrized in terms of $\log(\text{scale})$ and shape in order to perform unconstrained optimization
- `gpd. score`: score vector
- `gpd. infomat`: observed information matrix for GPD parametrized in terms of rate of m -year return level and shape
- `gpd. Vfun`: vector implementing conditioning on approximate ancillary statistics for the TEM
- `gpd. phi`: canonical parameter in the local exponential family approximation
- `gpd. dphi`: derivative matrix of the canonical parameter in the local exponential family approximation

Author(s)

Leo Belzile

ibvpot

Interpret bivariate threshold exceedance models

Description

This is an adaptation of the `evir` package `interpret.gpdbiv` function. `interpret.fbvpot` deals with the output of a call to `fbvpot` from the **evd** and to handle families other than the logistic distribution. The likelihood derivation comes from expression 2.10 in Smith et al. (1997).

Usage

```
ibvpot(fitted, q, silent = FALSE)
```

Arguments

| | |
|---------------------|--|
| <code>fitted</code> | the output of <code>fbvpot</code> or a list. See Details. |
| <code>q</code> | a vector of quantiles to consider, on the data scale. Must be greater than the thresholds. |
| <code>silent</code> | boolean; whether to print the interpretation of the result. Default to FALSE. |

Details

The list `fitted` must contain

- `model` a string; see `bvevd` from package `evd` for options
- `param` a named vector containing the parameters of the model, as well as parameters `scale1`, `shape1`, `scale2` and `shape2`, corresponding to marginal GPD parameters.
- `threshold` a vector of length 2 containing the two thresholds.
- `pat` the proportion of observations above the corresponding threshold

Value

an invisible numeric vector containing marginal, joint and conditional exceedance probabilities.

Author(s)

Leo Belzile, adapting original S code by Alexander McNeil

References

Smith, Tawn and Coles (1997), Markov chain models for threshold exceedances. *Biometrika*, **84**(2), 249–268.

See Also

`interpret.gpdbiv` in package `evir`

Examples

```
if (requireNamespace("evd", quietly = TRUE)) {  
  y <- rgp(1000,1,1,1)  
  x <- y*rmevspec(n=1000,d=2,sigma=cbind(c(0,0.5),c(0.5,0)), model='hr')  
  mod <- evd::fbvpot(x, threshold = c(1,1), model = 'hr', likelihood = 'censored')  
  ibvpot(mod, c(20,20))  
}
```

leedspollution

Leeds air pollution

Description

Daily maximum data (hourly for PM10) on air pollution for the Leeds Centre station in Yorkshire and Humberside station. The data goes from January 1st, 1993, until December 31st, 2024. Data show seasonality and there are some outliers. From December 2nd, 2008 onwards, particulate matters (PM10 and PM2.5) are measured using a tapered element oscillating microbalance (TEOM) and Filter Dynamics Measurement System (FDMS). The data for PM2.5 is missing before the change of instrumentation. A total of 231 daily measurements with only missing values were removed during preprocessing.

Usage

leedspollution

Format

A data frame with 11455 rows and 8 variables:

date [character] a date with format yyy-mm-dd
 O3 [integer] ozone (in nanograms per cubic meter)
 NO [integer] nitrogen oxyde (in nanograms per cubic meter)
 CO [double] carbon monoxyde (in micrograms per cubic meter)
 NO2 nitrogen dioxyde (in nanograms per cubic meter)
 SO2 sulphur dioxide (in nanograms per cubic meter)
 PM10 [integer] particulate matter 10, (in nanograms per cubic meter)
 PM2.5 [integer] particulate matter 2.5, (in nanograms per cubic meter)

Source

Crown 2025 copyright Defra via uk-air.defra.gov.uk, licenced under the Open Government Licence (OGL).

 maiquetia

Maiquetia Daily Rainfall

Description

Daily cumulated rainfall (in mm) at Maiquetia airport, Venezuela. The observations cover the period from January 1961 to December 1999. The original series had missing days in February 1996 (during which there were 2 days with 1hr each of light rain) and January 1998 (no rain). These were replaced by zeros.

Format

a vector of size 14244 containing daily rainfall (in mm),

Source

J.R. Cordova and M. González, accessed 25.11.2018 from <<https://rss.onlinelibrary.wiley.com/hub/journal/14679876/series-c-datasets>>

References

Coles, S. and L.R. Pericchi (2003). Anticipating Catastrophes through Extreme Value Modelling, *Applied Statistics*, **52**(4), 405-416.
 Coles, S., Pericchi L.R. and S. Sisson (2003). A fully probabilistic approach to extreme rainfall modeling, *Journal of Hydrology*, **273**, 35-50.

Examples

```
## Not run:
data(maiquetia, package = "mev")
day <- seq.Date(from = as.Date("1961-01-01"), to = as.Date("1999-12-31"), by = "day")
nzrain <- maiquetia[substr(day, 1, 4) < 1999 & maiquetia > 0]
fit.gpd(nzrain, threshold = 30, show = TRUE)

## End(Not run)
```

maxstable

Transform arguments using max stability

Description

Given a vector of location, scale and shape parameters, compute the corresponding parameters for block of size m assuming a generalized extreme value distribution.

Usage

```
maxstable(pars, m = 1L, inverse = FALSE)
```

Arguments

| | |
|---------|--|
| pars | vector of location, scale and shape parameters |
| m | [integer] block size |
| inverse | [logical] whether to compute the parameters for the inverse relationship (defaults to FALSE) |

Examples

```
maxstable(pars = maxstable(pars = c(1,2,0), m = 10), m = 10, inv = TRUE)
maxstable(pars = maxstable(pars = c(1,2,0.1), m = 5), m = 1/5)
```

mgp.c11

Censored likelihood for multivariate peaks over threshold models

Description

Censored likelihoods for various parametric limiting models over region determined by

$$\{y \in F : \max_{j=1}^D \sigma_j \frac{y_j^\xi - 1}{\xi_j} + \mu_j > u\};$$

where μ is loc, σ is scale and ξ is shape.

Usage

```
mgp.cll(
  dat,
  thresh,
  mthresh = thresh,
  loc,
  scale,
  shape,
  par,
  model = c("log", "neglog", "br", "xstud"),
  likt = c("mgp", "pois", "binom"),
  lambdau = 1,
  ...
)
```

Arguments

| | |
|----------------------|---|
| <code>dat</code> | matrix of observations |
| <code>thresh</code> | functional threshold for the maximum |
| <code>mthresh</code> | vector of individuals thresholds under which observations are censored |
| <code>loc</code> | vector of location parameter for the marginal generalized Pareto distribution |
| <code>scale</code> | vector of scale parameter for the marginal generalized Pareto distribution |
| <code>shape</code> | vector of shape parameter for the marginal generalized Pareto distribution |
| <code>par</code> | list of parameters: alpha for the logistic model, Lambda for the Brown–Resnick model or else Sigma and df for the extremal Student. |
| <code>model</code> | string indicating the model family, one of "log", "neglog", "br" or "xstud" |
| <code>likt</code> | string indicating the type of likelihood, with an additional contribution for the non-exceeding components: one of "mgp", "binom" and "pois". |
| <code>lambdau</code> | vector of marginal rate of marginal threshold exceedance. |
| <code>...</code> | additional arguments (see Details) |

Details

Optional arguments can be passed to the function via ...

- `censored` matrix of booleans and NA indicating whether observations `dat` fall below the mthreshold `mthresh`
- `c1` cluster instance created by `makeCluster` (default to NULL)
- `ncors` number of cores for parallel computing of the likelihood
- `numAbovePerRow` number of observations above `mthresh` (non-missing) per row
- `numAbovePerCol` number of observations above `mthresh` (non-missing) per column
- `mmax` maximum per column
- `B1` number of replicates for quasi Monte Carlo integral for the exponent measure

- B2 number of replicates for quasi Monte Carlo integral for the censored intensity contribution
- genvec1 generating vector for the quasi Monte Carlo routine (exponent measure), associated with B1
- genvec2 generating vector for the quasi Monte Carlo routine (individual obs contrib), associated with B2

Value

the value of the log-likelihood with attributes expme, giving the exponent measure

Note

The location and scale parameters are not identifiable unless one of them is fixed.

mgp.ll

Likelihood for multivariate peaks over threshold models

Description

Likelihood for the various parametric limiting models over region determined by

$$\{y \in F : \max_{j=1}^D \sigma_j \frac{y_j^\xi - 1}{\xi_j} + \mu_j > u\};$$

where μ is loc, σ is scale and ξ is shape.

Usage

```
mgp.ll(  
  dat,  
  thresh,  
  loc,  
  scale,  
  shape,  
  par,  
  model = c("log", "br", "xstud"),  
  likt = c("mgp", "pois", "binom"),  
  lambdau = 1,  
  ...  
)
```

Arguments

| | |
|--------|---|
| dat | matrix of observations |
| thresh | functional threshold for the maximum |
| loc | vector of location parameter for the marginal generalized Pareto distribution |

| | |
|---------|---|
| scale | vector of scale parameter for the marginal generalized Pareto distribution |
| shape | vector of shape parameter for the marginal generalized Pareto distribution |
| par | list of parameters: alpha for the logistic model, Lambda for the Brown–Resnick model or else Sigma and df for the extremal Student. |
| model | string indicating the model family, one of "log", "neglog", "br" or "xstud" |
| likt | string indicating the type of likelihood, with an additional contribution for the non-exceeding components: one of "mgp", "binom" and "pois". |
| lambdau | vector of marginal rate of marginal threshold exceedance. |
| ... | additional arguments (see Details) |

Details

Optional arguments can be passed to the function via ...

- c1 cluster instance created by makeCluster (default to NULL)
- ncors number of cores for parallel computing of the likelihood
- mmax maximum per column
- B1 number of replicates for quasi Monte Carlo integral for the exponent measure
- genvec1 generating vector for the quasi Monte Carlo routine (exponent measure), associated with B1

Value

the value of the log-likelihood with attributes expme, giving the exponent measure

Note

The location and scale parameters are not identifiable unless one of them is fixed.

nidd

River Nidd Flow

Description

The data consists of exceedances over the threshold 65 cubic meter per second of the River Nidd at Hunsingore Weir, for 35 years of data between 1934 and 1969.

Format

a vector of size 154

Source

Natural Environment Research Council (1975). *Flood Studies Report*, volume 4. pp. 235–236.

References

Davison, A.C. and R.L. Smith (1990). Models for Exceedances over High Thresholds (with discussion), *Journal of the Royal Statistical Society. Series B (Methodological)*, **52**(3), 393–442.

See Also

nidd.thresh from the evir package

nutrients

Nutrient data

Description

Interview component of survey 'What we eat in America'. These are extracted from the 2015–2016 National Health and Nutrition Examination Survey (NHANES, <https://wwwn.cdc.gov/nchs/nhanes/Default.aspx>) report and consist of the total nutrients for all food and beverage intake ingested over a 24 hours period.

Usage

nutrients

Format

A data frame with 9544 rows and 38 variables:

prot proteins (in grams)

carb carbohydrate (in gram)

sugr total sugars (in gram)

fibe dietary fibers (in grams)

tfat total fat (in grams)

sfat saturated fat (in grams)

mfat monounsaturated fat (in grams)

pfat polyunsaturated fat (in grams)

chol cholesterol (in milligrams)

atoc vitamin E as alpha-tocopherol (in milligrams)

ret retinol (in micrograms)

vara Vitamin A as retinol activity equivalents (in micrograms).

acar alpha-carotene (in micrograms)

bcar beta-carotene (in micrograms)

cryp beta-cryptoxanthin (in micrograms)

lyco lycopene (in micrograms)

lz lutein and zeaxanthin (in micrograms).
vb1 thiamin (vitamin B1, in milligrams)
vb2 riboflavin (vitamin B2, in milligrams)
niac niacin (in milligrams)
vb6 vitamin B5 (in milligrams)
fo1a total folate (in micrograms)
fa folic acid (in micrograms)
ff food folate (in micrograms)
ch1 total choline (in milligrams)
vb12 vitamin B12 (in micrograms)
vc vitamin C (in milligrams)
vd vitamin D (comprising D2 and D3, in micrograms)
vk vitamin K (in micrograms)
calc calcium (in milligrams)
phos phosphorus (in milligrams)
magn magnesium (in milligrams)
iron iron (in milligrams)
zinc zinc (in milligrams)
copp copper (in milligrams)
sodi sodium (in milligrams)
pota potassium (in milligrams)
sele selenium (in micrograms)

Details

Note that the sample design oversampled specific population targets and that only respondents are provided. The website contains more information about sampling weights. There are multiple missing records.

Note

These data are subject to a data user agreement, available at <https://www.cdc.gov/nchs/policy/data-user-agreement.html>

Source

National Center for Health Statistics, now available from the Wayback Machine via https://web.archive.org/web/20201029113801/https://www.cdc.gov/Nchs/Nhanes/2015-2016/DR1TOT_I.XPT

pandemics *Deaths from pandemics*

Description

The data base contains estimated records of the number of deaths from pandemics.

Usage

pandemics

Format

A data frame with 72 rows and 8 variables:

event name of the event
 startyear start year of the event
 endyear end year of the event
 lower lower bound on estimated deaths (in thousands)
 average average estimated deaths (in thousands)
 upper upper bound on estimated deaths (in thousands)
 saverage scaled average of estimated deaths (in thousands)
 population estimated population at risk (in thousands)

Source

Cirillo, P. and N.N. Taleb (2020). *Tail risk of contagious diseases*. Nat. Phys. **16**, 606–613 (2020).
 <doi:10.1038/s41567-020-0921-x>

penultimate *Smith's penultimate approximations*

Description

The function takes as arguments the distribution and density functions. There are two options: method='bm' yields block maxima and method='pot' threshold exceedances. For method='bm', the user should provide in such case the block sizes via the argument m, whereas if method='pot', a vector of threshold values should be provided. The other argument (thresh or m depending on the method) is ignored.

Usage

```
penultimate(family, method = c("bm", "pot"), thresh, qllev, m, ...)
```

Arguments

| | |
|--------|--|
| family | the name of the parametric family. Will be used to obtain <code>dfamily</code> , <code>pfamily</code> , <code>qfamily</code> |
| method | either block maxima ('bm') or peaks-over-threshold ('pot') are supported |
| thresh | vector of thresholds for method 'pot' |
| qllev | vector of quantile levels for method 'pot', e.g., 0.9, 0.95, ... Ignored if argument thresh is provided. |
| m | vector of block sizes for method 'bm' |
| ... | additional arguments passed to <code>densF</code> and <code>distF</code> |

Details

Alternatively, the user can provide functions `densF`, `quantF` and `distF` for the density, quantile function and distribution functions, respectively. The user can also supply the derivative of the density function, `ddensF`. If the latter is missing, it will be approximated using finite-differences.

For method = "pot", the function computes the reciprocal hazard and its derivative on the log scale to avoid numerical overflow. Thus, the density function should have argument `log` and the distribution function arguments `log.p` and `lower.tail`, respectively.

Value

a data frame containing

- `loc`: location parameters (method='bm')
- `scale`: scale parameters
- `shape`: shape parameters
- `thresh`: thresholds (if method='pot'), percentile corresponding to threshold (if method='pot')
- `m`: block sizes (if method='bm')

Author(s)

Leo Belzile

References

Smith, R.L. (1987). Approximations in extreme value theory. *Technical report 205*, Center for Stochastic Process, University of North Carolina, 1–34.

Examples

```
# Threshold exceedance for Normal variables
quants <- seq(1, 5, by = 0.02)
penult <- penultimate(
  family = "norm",
  method = 'pot',
  thresh = quants,
  ddensF = function(x){-x*dnorm(x)}, # optional argument
```

```

)
plot(x = quants,
     y = penult$shape,
     type = 'l',
     xlab = 'quantile',
     ylab = 'Penultimate shape',
     ylim = c(-0.5, 0))
# Block maxima for Gamma variables
# User must provide arguments for shape (or rate), for which there is no default
m <- seq(30, 3650, by = 30)
penult <- penultimate(family = 'gamma', method = 'bm', m = m, shape = 0.1)
plot(x = m,
     y = penult$shape,
     type = 'l',
     xlab = 'quantile',
     ylab = 'penultimate shape')

# Comparing density of GEV approximation with true density of maxima
m <- 100 # block of size 100
p <- penultimate(
  family = 'norm',
  ddensF = function(x){-x*dnorm(x)},
  method = 'bm',
  m = m)
x <- seq(1, 5, by = 0.01)
plot(
  x = x,
  y = m * dnorm(x) * exp((m-1) * pnorm(x, log.p = TRUE)),
  type = 'l',
  ylab = 'density',
  main = 'Distribution of the maxima of\n 100 standard normal variates')
lines(x, mev::dgev(x, loc = p$loc, scale = p$scale, shape = 0), col = 2)
lines(x, mev::dgev(x, loc = p$loc, scale = p$scale, shape = p$shape), col = 4)
legend(
  x = 'topright',
  lty = c(1, 1, 1),
  col = c(1, 2, 4),
  legend = c('exact', 'ultimate', 'penultimate'),
  bty = 'n')

```

plot.eprof

Plot of (modified) profile likelihood

Description

The function plots the (modified) profile likelihood and the tangent exponential profile likelihood

Usage

```
## S3 method for class 'eprof'
plot(x, ...)
```

Arguments

x an object of class eprof returned by `gpd.pll` or `gev.pll`.
 ... further arguments to plot.

Value

a graph of the (modified) profile likelihoods

References

Brazzale, A. R., Davison, A. C. and Reid, N. (2007). *Applied Asymptotics: Case Studies in Small-Sample Statistics*. Cambridge University Press, Cambridge.
 Severini, T. A. (2000). *Likelihood Methods in Statistics*. Oxford University Press, Oxford.

plot.fr

Plot of tangent exponential model profile likelihood

Description

This function is adapted from the `plot.fr` function from the `hoa` package bundle. It differs from the latter mostly in the placement of legends.

Usage

```
## S3 method for class 'fr'
plot(x, ...)
```

Arguments

x an object of class fr returned by `gpd.tem` or `gev.tem`.
 ... further arguments to plot currently ignored. Providing a numeric vector which allows for custom selection of the plots. A logical `all`. See **Details**.

Details

Plots produced depend on the integers provided in `which`. 1 displays the Wald pivot, the likelihood root r , the modified likelihood root r_{star} and the likelihood modification q as functions of the parameter ψ . 2 gives the renormalized profile log likelihood and adjusted form, with the maximum likelihood having ordinate value of zero. 3 provides the significance function, a transformation of 1. Lastly, 4 plots the correction factor as a function of the likelihood root; it is a diagnostic plot aimed for detecting failure of the asymptotic approximation, often due to poor numerics in a neighborhood of $r=0$; the function should be smooth. The function `spline.corr` is designed to handle this by correcting numerically unstable estimates, replacing outliers and missing values with the fitted values from the fit.

Value

graphs depending on argument which

References

Brazzale, A. R., Davison, A. C. and Reid, N. (2007). *Applied Asymptotics: Case Studies in Small-Sample Statistics*. Cambridge University Press, Cambridge.

plot.mev_thselect_wadsworth

Sequential analysis diagnostic plots for threshold selection

Description

Function to produce diagnostic plots and test statistics for the threshold diagnostics exploiting structure of maximum likelihood estimators based on the non-homogeneous Poisson process likelihood or the coefficient of tail dependence

Usage

```
## S3 method for class 'mev_thselect_wadsworth'
plot(x, type = c("wn", "ps"), ...)

thselect.wseq(
  xdat,
  thresh,
  qllev,
  model = c("nhpp", "taildep", "rtaildep"),
  npp = 1,
  nsim = 1000L,
  level = 0.95,
  plot = FALSE,
  ...
)
```

Arguments

| | |
|--------|---|
| x | object returned by a call to thselect.wseq |
| type | string giving the plots to produce |
| ... | additional parameters passed to internal routine |
| xdat | a numeric vector or matrix of data to be fitted. |
| thresh | vector of candidate thresholds. |
| qllev | vector of probabilities for empirical quantiles used in place of the threshold, used if argument thresh is missing. |

| | |
|-------|--|
| model | string specifying whether the univariate or multivariate diagnostic should be used. Either nhpp for the univariate model, or exp (invexp) for the bivariate exponential model with rate (inverse rate) parametrization. See details. |
| npp | number of observations per period for the non-homogeneous point process model. Default to 1. |
| nsim | number of Monte Carlo simulations used to assess the null distribution of the test statistic |
| level | confidence level of intervals, defaults to 0.95 |
| plot | logical; if TRUE, calls the plot routine |

Details

The function is a wrapper for the univariate (non-homogeneous Poisson process model) and exponential dependence model applied to the minimum component (tail dependence coefficient). For the latter, the user can select either the rate ("taildep" or inverse rate parameter ("rtaildep"). The inverse rate parametrization works better for uniformity of the p-value distribution under the likelihood ratio test for the changepoint.

For the coefficient of tail dependence, users must provide pairwise minimum of marginally exponentially distributed margins (see example)

Value

an object of class invisible list with components

- thresh0: threshold selected by the likelihood ratio procedure
- thresh: vector of candidate thresholds
- coef: maximum likelihood estimates from all thresholds
- vcov: joint asymptotic covariance matrix for shape ξ or coefficient of tail dependence η , or it's reciprocal.
- wn: values of the white noise process
- stat: value of the likelihood ratio test statistic for the changepoint test
- pval: P -value of the likelihood ratio test
- mle: maximum likelihood estimates for the selected threshold
- model: model fitted, either nhpp, exp or invexp
- nsim: number of Monte Carlo simulations for changepoint test
- xdat: vector of observations

Author(s)

Jennifer L. Wadsworth, Léo Belzile

References

Wadsworth, J.L. (2016). Exploiting Structure of Maximum Likelihood Estimators for Extreme Value Threshold Selection, *Technometrics*, **58**(1), 116-126, <http://dx.doi.org/10.1080/00401706.2014.998345>.

Examples

```
## Not run:
set.seed(123)
xdat <- abs(rnorm(5000))
thresh <- quantile(xdat, seq(0, 0.9, by = 0.1))
(diag <- thselect.wseq(
  xdat = xdat,
  thresh = thresh,
  plot = TRUE,
  type = "ps"))
# Multivariate example, with coefficient of tail dependence
xbvn <- rmnorm(n = 6000L,
               mu = rep(0, 2),
               Sigma = cbind(c(1, 0.7), c(0.7, 1)))
thselect.wseq(
  xdat = xbvn,
  qllev = seq(0, 0.9, length.out = 30),
  model = 'taildep',
  plot = TRUE)

## End(Not run)
```

plot.mev_tstab_mrl *Mean residual life parameter stability plot*

Description

Mean residual life parameter stability plot

Usage

```
## S3 method for class 'mev_tstab_mrl'
plot(
  x,
  xlab = c("thresh", "nexc"),
  level = 0.95,
  type = c("band", "ptwise"),
  ...
)
```

Arguments

| | |
|-------|--|
| x | object resulting from a call to <code>tstab.mrl</code> |
| xlab | [string]; whether to plot mean residual life plot as a function of threshold value of number of exceedances |
| level | [numeric] level of Wald confidence intervals |
| type | [string] whether to plot pointwise confidence intervals using segments ("ptwise") or using dashed lines ("band") |
| ... | additional arguments, currently ignored |

Value

NULL; use to produce plots

pp *Poisson process of extremes.*

Description

Likelihood, score function and information matrix for the Poisson process likelihood.

Arguments

| | |
|--------|---|
| par | vector of loc, scale and shape |
| dat | sample vector |
| u | threshold |
| method | string indicating whether to use the expected ('exp') or the observed ('obs' - the default) information matrix. |
| np | number of periods of observations. This is a <i>post hoc</i> adjustment for the intensity so that the parameters of the model coincide with those of a generalized extreme value distribution with block size $\text{length}(\text{dat})/\text{np}$. |
| nobs | number of observations for the expected information matrix. Default to $\text{length}(\text{dat})$ if dat is provided. |

Usage

```
pp.ll(par, dat)
pp.ll(par, dat, u, np)
pp.score(par, dat)
pp.informat(par, dat, method = c('obs', 'exp'))
```

Functions

- pp.ll: log likelihood
- pp.score: score vector
- pp.informat: observed or expected information matrix

Author(s)

Leo Belzile

References

- Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*, Springer, 209 p.
- Wadsworth, J.L. (2016). Exploiting Structure of Maximum Likelihood Estimators for Extreme Value Threshold Selection, *Technometrics*, **58**(1), 116-126, <http://dx.doi.org/10.1080/00401706.2014.998345>.
- Sharkey, P. and J.A. Tawn (2017). A Poisson process reparameterisation for Bayesian inference for extremes, *Extremes*, **20**(2), 239-263, <http://dx.doi.org/10.1007/s10687-016-0280-2>.

qqplot.blocksize *Diagnostic plots for max-stability based on blocks of GEV samples*

Description

Given a sample of ordered GEV draws, calculate the ingredients of diagnostic quantile-quantile plots using the bootstrap

Usage

```
qqplot.blocksize(
  xdat,
  type = c("max", "range", "all"),
  B = 1000L,
  marginal = FALSE,
  rounding = 0,
  lb = NULL,
  plot = TRUE,
  level = 0.95,
  np = NULL
)
```

Arguments

| | |
|----------|---|
| xdat | n by m matrix of GEV observations, ordered by row from smallest to largest |
| type | string; the statistic to return. Either the maximum of each row (max), the standardized difference between the penultimate and largest value (spacing), the ratio of maximum to spacing (ratio) or the whole sample (all) |
| B | number of bootstrap samples |
| marginal | logical; if TRUE, estimates are based on the marginal likelihood of the $m - 1$ smallest order statistics of the sample |
| rounding | amount of rounding |
| lb | lower bound for left-censoring, default to NULL in absence |
| plot | logical; if TRUE (default), returns a quantile-quantile plot |
| level | confidence level for confidence and tolerance intervals |
| np | number of points at which to evaluate quantile-quantile plots. Must be either NULL, or a vector of integer of the same length as type (otherwise it is recycled). |

Value

a list with elements for building quantile-quantile plots, including

- plots list of plots with elements x, y, a list confint with matrices simultaneous and pointwise, type of value and distribution (currently only uniform)
- mle: maximum likelihood estimate of the location, scale, and shape

- param B by 3 matrix of bootstrap parameter estimates
- type vector of string with statistics
- bootstrap type of bootstrap, only parametric for now
- n number of rows of xdat
- m number of columns of xdat for comparison
- marginal logical; if TRUE, uses the marginal likelihood of the $m - 1$ smallest order statistics per block for estimation
- icens logical; if TRUE, data treated as rounded (interval-censored)
- lcens logical; if TRUE, data are left-censored below lb
- lb lower bound for left-censoring
- rounding double δ indicating the amount of rounding, assuming $\delta/2$ on either size of the reported value
- xdat matrix of original observations

Examples

```
xdat <- build.blocks(mev::rgev(n = 50), m = 2)
## Not run:
qqplot.blocksize(xdat, type = "max", marginal = TRUE, B = 100)

## End(Not run)
```

| | |
|-------------|--|
| qqplot.unif | <i>Pointwise and simultaneous binomial confidence intervals for uniform via simulation</i> |
|-------------|--|

Description

Given a vector of draws transformed using the probability integral transform scale to what should be uniform positions, produce plots with pointwise and simultaneous confidence intervals for uniformity.

Usage

```
qqplot.unif(xdat, K = 100, B = 1000, level = 0.95, plot = TRUE)
```

Arguments

| | |
|-------|---|
| xdat | vector of N postulated uniform samples, obtained by applying the ECDF |
| K | number of evaluation points for the plotting positions |
| B | number of Monte Carlo samples |
| level | vector of pointwise and simultaneous confidence levels, recycled if necessary |
| plot | logical; if TRUE, produce a plot of the empirical distribution function |

References

Sailynoja, T., Burkner, P.C. and Vehtari, A. (2022). Graphical test for discrete uniformity and its applications in goodness-of-fit evaluation and multiple sample comparison, *Statistics and Computing*, 32, doi:10.1007/s11222022100906

Examples

```
xdat <- runif(200)
qqplot.unif(xdat)
```

qweissman

Weissman's quantile estimator

Description

Given a small probability of exceedance p , the number of exceedances k out of n observation above the threshold u (thresh) (corresponding typically to the $(k + 1)$ th order statistic, compute the tail quantile at level $Q(1 - p)$ using the estimator of Weissman (1978) under the assumption of Pareto tail (positive shape ξ), viz.

$$Q(1 - p) = u \left(\frac{k}{pn} \right)^\xi .$$

Usage

```
qweissman(
  p,
  k,
  n,
  thresh,
  shape,
  confint = c("none", "bbw1", "bbw2", "bbw3"),
  level = 0.95
)
```

Arguments

| | |
|----------------------|---|
| <code>p</code> | tail probability, must be larger than the proportion of exceedances k/n . |
| <code>k</code> | vector of the number of exceedances above thresh |
| <code>n</code> | integer, total sample size |
| <code>thresh</code> | vector of thresholds |
| <code>shape</code> | vector of positive shape parameters |
| <code>confint</code> | string indicating the type of confidence interval. |
| <code>level</code> | level of confidence intervals, default to 0.95. |

Value

a vector of tail quantiles if `confint = "none"` (default), or a data frame with columns `quantile`, `lower` and `upper` for the point estimates and confidence intervals of the quantiles.

Note

The confidence interval estimators are those for Hill estimator derived in Buidentag, Beirlant and de Wet (2020) in equations 23 (`bbw1`), 28 (`bbw2`) and 31 (`bbw3`, saddlepoint approximation) under the assumption of zero asymptotic bias.

References

Weissman, I. (1978). Estimation of Parameters and Larger Quantiles Based on the k Largest Observations. *Journal of the American Statistical Association*, 73(364), 812–815. <doi:10.2307/2286285>.

Buitendag S, Beirlant J and de Wet T. (2020) Confidence intervals for extreme Pareto-type quantiles. *Scandinavian Journal of Statistics*, 47, 36–55. <doi:10.1111/sjos.12396>.

Examples

```
set.seed(2025)
p <- 1/100
xdat <- rgp(n = 1000, loc = 2, scale = 2, shape = 0.4)
hill <- shape.hill(xdat, k = seq(20L, 100L, by = 10L))
thresh <- sort(xdat, decreasing = TRUE)[hill$k+1]
qweissman(
  p = 1/100,
  k = hill$k,
  n = length(xdat),
  thresh = thresh,
  shape = hill$shape)
# Compare with true quantile
qgp(1/100, loc = 2, scale = 2, shape = 0.4, lower.tail = FALSE)
```

rdir

Random variate generation for Dirichlet distribution on S_d

Description

A function to sample Dirichlet random variables, based on the representation as ratios of Gamma. Note that the RNG will generate on the full simplex and the sum to one constraint is respected here

Usage

```
rdir(n, alpha, normalize = TRUE)
```

Arguments

| | |
|-----------|--|
| n | sample size |
| alpha | vector of parameter |
| normalize | boolean. If FALSE, the function returns Gamma variates with parameter alpha. |

Value

sample of dimension d (size of alpha) from the Dirichlet distribution.

Examples

```
rdir(n=100, alpha=c(0.5,0.5,2),TRUE)
rdir(n=100, alpha=c(3,1,2),FALSE)
```

| | |
|--------|---|
| rgparp | <i>Simulation from generalized R-Pareto processes</i> |
|--------|---|

Description

The generalized R-Pareto process is supported on $(loc - scale / shape, Inf)$ if $shape > 0$, or $(-Inf, loc - scale / shape)$ for negative shape parameters, conditional on $(X - r(loc)) / r(scale) > 0$. The standard Pareto process corresponds to $scale = loc = rep(1, d)$.

Usage

```
rgparp(
  n,
  shape = 1,
  thresh = 1,
  risk = c("mean", "sum", "site", "max", "min", "l2"),
  siteindex = NULL,
  d,
  loc,
  scale,
  param,
  sigma,
  model = c("log", "neglog", "bilog", "negbilog", "hr", "br", "xstud", "smith",
    "schlather", "ct", "sdir", "dirmix"),
  weights,
  vario,
  coord = NULL,
  ...
)
```

Arguments

| | |
|-----------|--|
| n | number of observations |
| shape | shape parameter of the generalized Pareto variable |
| thresh | univariate threshold for the exceedances of risk functional |
| risk | string indicating the risk functional. |
| siteindex | integer between 1 and d specifying the index of the site or variable |
| d | dimension of sample |
| loc | location vector |
| scale | scale vector |
| param | parameter vector for the logistic, bilogistic, negative bilogistic and extremal Dirichlet (Coles and Tawn) model. Parameter matrix for the Dirichlet mixture. Degree of freedoms for extremal student model. See Details . |
| sigma | covariance matrix for Brown-Resnick and extremal Student-t distributions. Symmetric matrix of squared coefficients λ^2 for the Husler-Reiss model, with zero diagonal elements. |
| model | for multivariate extreme value distributions, users can choose between 1-parameter logistic and negative logistic, asymmetric logistic and negative logistic, bilogistic, Husler-Reiss, extremal Dirichlet model (Coles and Tawn) or the Dirichlet mixture. Spatial models include the Brown-Resnick, Smith, Schlather and extremal Student max-stable processes. Max linear models are also supported |
| weights | vector of length m for the m mixture components that sum to one. For the "maxlin" model, weights should be a matrix with d columns that represent the weight of the components and whose column sum to one (if provided, this argument overrides asy). |
| vario | semivariogram function whose first argument must be distance. Used only if provided in conjunction with coord and if sigma is missing |
| coord | d by k matrix of coordinates, used as input in the variogram vario or as parameter for the Smith model. If grid is TRUE, unique entries should be supplied. |
| ... | additional arguments for the vario function |

Value

an n by d sample from the generalized R-Pareto process, with attributes `accept.rate` if the procedure uses rejection sampling.

Examples

```
rgparp(n = 10, risk = 'site', siteindex = 2, d = 3, param = 2.5,
       model = 'log', scale = c(1, 2, 3), loc = c(2, 3, 4), shape = 0.5)
rgparp(n = 10, risk = 'max', d = 4, param = c(0.2, 0.1, 0.9, 0.5),
       scale = 1:4, loc = 1:4, model = 'bilog')
rgparp(n = 10, risk = 'sum', d = 3, param = c(0.8, 1.2, 0.6, -0.5),
       scale = 1:3, loc = 1:3, model = 'sdir')
vario <- function(x, scale = 0.5, alpha = 0.8){ scale*x^alpha }
grid.coord <- as.matrix(expand.grid(runif(4), runif(4)))
```

```
rgparp(n = 10, risk = 'max', vario = vario, coord = grid.coord,
       model = 'br', scale = runif(16), loc = rnorm(16))
```

rho.dk

*Second order tail index estimator of Drees and Kaufmann***Description**

Estimator of the second order regular variation parameter $\rho \leq 0$ parameter for heavy-tailed data proposed by Drees and Kaufmann (1998)

Usage

```
rho.dk(xdat, k, tau = 0.5)
```

Arguments

| | |
|------|--|
| xdat | vector of positive observations |
| k | number of highest order statistics to use for estimation |
| tau | tuning parameter $\tau \in (0, 1)$ |

References

Drees, H. and E. Kaufmann (1998). Selecting the optimal sample fraction in univariate extreme value estimation, *Stochastic Processes and their Applications*, 75(2), 149-172, <doi:10.1016/S0304-4149(98)00017-9>.

rho.fagh

Second order tail index estimator of Fraga Alves et al. Estimator of the second order regular variation parameter $\rho \leq 0$ parameter for heavy-tailed data proposed by Fraga Alves et al. (2003)

Description

Second order tail index estimator of Fraga Alves et al. Estimator of the second order regular variation parameter $\rho \leq 0$ parameter for heavy-tailed data proposed by Fraga Alves et al. (2003)

Usage

```
rho.fagh(xdat, k, tau = 0)
```

Arguments

| | |
|------|---|
| xdat | vector of positive observations |
| k | number of highest order statistics to use for estimation |
| tau | scalar real tuning parameter. Default values is 0, which is typically chosen whenever $\rho \geq -1$. The choice $\tau = 1$ otherwise. |

References

Fraga Alves, M.I., Gomes, M. Ivette, and de Haan, Laurens (2003). A new class of semi-parametric estimators of the second order parameter. *Portugaliae Mathematica*. Nova Serie 60(2), 193-213. <<http://eudml.org/doc/50867>>.

Examples

```
# Example with rho = -0.2
n <- 1000
xdat <- mev::rgp(n = n, shape = 0.2)
kmin <- floor(n^0.995)
kmax <- ceiling(n^0.999)
rho_est <- rho.fagh(
  xdat = xdat,
  k = n - kmin:kmax)
rho_med <- mean(rho_est$rho)
```

rho.gbw

Second order tail index estimator of Goegebeur et al. (2008)

Description

Estimator of the second order regular variation parameter $\rho \leq 0$ parameter for heavy-tailed data based on ratio of kernel goodness-of-fit statistics.

Usage

```
rho.gbw(xdat, k)
```

Arguments

| | |
|------|--|
| xdat | vector of positive observations |
| k | number of highest order statistics to use for estimation |

References

Goegebeur, Y., J. Beirlant and T. de Wet (2008). Linking Pareto-tail kernel goodness-of-fit statistics with tail index at optimal threshold and second order estimation. *REVSTAT-Statistical Journal*, 6(1), 51-69. <[doi:10.57805/revstat.v6i1.57](https://doi.org/10.57805/revstat.v6i1.57)>

| | |
|---------|--|
| rho.ghp | <i>Second order tail index estimator of Gomes et al.</i> |
|---------|--|

Description

Estimator of the second order regular variation parameter $\rho \leq 0$ parameter for heavy-tailed data proposed by Gomes et al. (2003)

Usage

```
rho.ghp(xdat, k, alpha = 2)
```

Arguments

| | |
|-------|--|
| xdat | vector of positive observations |
| k | number of highest order statistics to use for estimation |
| alpha | positive scalar tuning parameter |

References

Gomes, M.I., de Haan, L. & Peng, L. (2002). Semi-parametric Estimation of the Second Order Parameter in Statistics of Extremes. *Extremes* 5, 387–414. <doi:10.1023/A:1025128326588>

| | |
|-------|---|
| rlarg | <i>Distribution of the r-largest observations</i> |
|-------|---|

Description

Likelihood, score function and information matrix for the r-largest observations likelihood.

Arguments

| | |
|--------|---|
| par | vector of loc, scale and shape |
| dat | an n by r sample matrix, ordered from largest to smallest in each row |
| method | string indicating whether to use the expected ('exp') or the observed ('obs' - the default) information matrix. |
| nobs | number of observations for the expected information matrix. Default to nrow(dat) if dat is provided. |
| r | number of order statistics kept. Default to ncol(dat) |

Usage

```
rlarg.ll(par, dat, u, np)
rlarg.score(par, dat)
rlarg.infomat(par, dat, method = c('obs', 'exp'), nobs = nrow(dat), r = ncol(dat))
```

Functions

- `rlarg.ll`: log likelihood
- `rlarg.score`: score vector
- `rlarg.informat`: observed or expected information matrix

Author(s)

Leo Belzile

References

- Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*, Springer, 209 p.
- Smith, R.L. (1986). Extreme value theory based on the r largest annual events, *Journal of Hydrology*, **86**(1-2), 27–43, [http://dx.doi.org/10.1016/0022-1694\(86\)90004-1](http://dx.doi.org/10.1016/0022-1694(86)90004-1).

rmar1

Simulation from first-order max-autoregressive processes

Description

Generate data from stationary sequences for extremes for non-negative shapes, following Tavares (1977) for the Gumbel case.

Usage

```
rmar1(n, theta, shape = 0)
```

Arguments

| | |
|--------------------|---|
| <code>n</code> | sample size |
| <code>theta</code> | extremal index, a value in $(0,1]$ |
| <code>shape</code> | non-negative shape parameter of the GEV |

Details

The models are parametrized in terms of extremal index $\theta \in (0, 1]$.

When $\text{shape} = 0$, the stationary process has unit Gumbel margins. When $\text{shape} > 0$, the margins have Frechet margins with distribution $F(x) = \exp(-x^{-1/\xi})$.

Value

a vector of length n drawn from the stationary distribution.

References

- Valadares Tavares, L. (1977). The Exact Distribution of Extremes of a Non-Gaussian Process. *Stochastic Processes and Their Applications* (2): 151-56. doi:10.1016/03044149(77)900266
- Davis, Richard A., and Sidney I. Resnick (1989). Basic Properties and Prediction of Max-ARMA Processes, *Advances in Applied Probability*, 21 (4): 781–803. doi:10.2307/1427767.

Examples

```
X1 <- rmar1(n = 1000, theta = 0.5)
X2 <- rmar1(n = 1000, theta = 0.2, shape = 0.2)
par(mfrow = c(1, 2))
plot(X1)
plot(X2)
xacf(X1, qllev = 0.9)
xacf(X2, qllev = 0.9)
```

 rmev

Exact simulations of multivariate extreme value distributions

Description

Implementation of the random number generators for multivariate extreme-value distributions and max-stable processes based on the two algorithms described in Dombry, Engelke and Oesting (2016).

Usage

```
rmev(
  n,
  d,
  param,
  asy,
  sigma,
  model = c("log", "alog", "neglog", "aneglog", "bilog", "negbilog", "hr", "br", "xstud",
    "smith", "schlather", "ct", "sdir", "dirmix", "pairbeta", "pairexp", "wdirbs",
    "wexpbs", "maxlin"),
  alg = c("ef", "sm"),
  weights = NULL,
  vario = NULL,
  coord = NULL,
  grid = FALSE,
  dist = NULL,
  ...
)
```

Arguments

| | |
|---------|--|
| n | number of observations |
| d | dimension of sample |
| param | parameter vector for the logistic, bilogistic, negative bilogistic and extremal Dirichlet (Coles and Tawn) model. Parameter matrix for the Dirichlet mixture. Degree of freedoms for extremal student model. See Details . |
| asy | list of asymmetry parameters, as in function <code>rmvevd</code> from package <code>evd</code> , of $2^d - 1$ vectors of size corresponding to the power set of d , with sum to one constraints. |
| sigma | covariance matrix for Brown-Resnick and extremal Student-t distributions. Symmetric matrix of squared coefficients λ^2 for the Husler-Reiss model, with zero diagonal elements. |
| model | for multivariate extreme value distributions, users can choose between 1-parameter logistic and negative logistic, asymmetric logistic and negative logistic, bilogistic, Husler-Reiss, extremal Dirichlet model (Coles and Tawn) or the Dirichlet mixture. Spatial models include the Brown-Resnick, Smith, Schlather and extremal Student max-stable processes. Max linear models are also supported |
| alg | algorithm, either simulation via extremal function ('ef') or via the spectral measure ('sm'). Default to ef. |
| weights | vector of length m for the m mixture components that sum to one. For the "maxlin" model, weights should be a matrix with d columns that represent the weight of the components and whose column sum to one (if provided, this argument overrides <code>asy</code>). |
| vario | semivariogram function whose first argument must be distance. Used only if provided in conjunction with <code>coord</code> and if <code>sigma</code> is missing |
| coord | d by k matrix of coordinates, used as input in the variogram <code>vario</code> or as parameter for the Smith model. If <code>grid</code> is TRUE, unique entries should be supplied. |
| grid | Logical. TRUE if the coordinates are two-dimensional grid points (spatial models). |
| dist | symmetric matrix of pairwise distances. Default to NULL. |
| ... | additional arguments for the <code>vario</code> function |

Details

The vector `param` differs depending on the model

- `log`: one dimensional parameter greater than 1
- `alog`: $2^d - d - 1$ dimensional parameter for dep. Values are recycled if needed.
- `neglog`: one dimensional positive parameter
- `aneglog`: $2^d - d - 1$ dimensional parameter for dep. Values are recycled if needed.
- `bilog`: d -dimensional vector of parameters in $[0, 1]$
- `negbilog`: d -dimensional vector of negative parameters
- `ct`, `dir`, `negdir`, `sdir`: d -dimensional vector of positive (a)symmetry parameters. For `dir` and `negdir`, a $d + 1$ vector consisting of the d Dirichlet parameters and the last entry is an index of regular variation in $(-\min(\alpha_1, \dots, \alpha_d), 1]$ treated as shape parameter

- `xstud`: one dimensional parameter corresponding to degrees of freedom α
- `dirmix`: d by m -dimensional matrix of positive (a)symmetry parameters
- `pairbeta`, `pairexp`: $d(d-1)/2+1$ vector of parameters, containing the concentration parameter and the coefficients of the pairwise beta, in lexicographical order e.g., $\beta_{12}, \beta_{13}, \dots$
- `wdirbs`, `wexpbs`: $2d$ vector of d concentration parameters followed by the d Dirichlet parameters

Stephenson points out that the multivariate asymmetric negative logistic model given in e.g. Coles and Tawn (1991) is not a valid distribution function in dimension $d > 3$ unless additional constraints are imposed on the parameter values. The implementation in `mev` uses the same construction as the asymmetric logistic distribution (see the vignette). As such it does not match the bivariate implementation of `rbvevd`.

The dependence parameter of the `evd` package for the Husler-Reiss distribution can be recovered taking for the Brown–Resnick model $2/r = \sqrt{(2\gamma(h))}$ where h is the lag vector between sites and $r = 1/\lambda$ for the Husler–Reiss.

Value

an n by d exact sample from the corresponding multivariate extreme value model

Warning

As of version 1.8 (August 16, 2016), there is a distinction between models `hr` and `br`. The latter is meant to be used in conjunction with variograms. The parametrization differs between the two models.

The family of scaled Dirichlet is now parametrized by a parameter in $-\min(\alpha)$ appended to the the d vector `param` containing the parameter α of the Dirichlet model. Arguments `model='dir'` and `model='negdir'` are still supported internally, but not listed in the options.

Author(s)

Leo Belzile

References

Dombry, Engelke and Oesting (2016). Exact simulation of max-stable processes, *Biometrika*, **103**(2), 303–317.

See Also

[rmevspec](#), [rmvevd](#), [rbvevd](#)

Examples

```
set.seed(1)
rmev(n=100, d=3, param=2.5, model='log', alg='ef')
rmev(n=100, d=4, param=c(0.2,0.1,0.9,0.5), model='bilog', alg='sm')
## Spatial example using power variogram
##NEW: Semi-variogram must take distance as argument
```

```

semivario <- function(x, scale, alpha){ scale*x^alpha }
#grid specification
grid.coord <- as.matrix(expand.grid(runif(4), runif(4)))
rmev(n=100, vario=semivario, coord=grid.coord, model='br', scale = 0.5, alpha = 1)
#using the Brown-Resnick model with a covariance matrix
vario2cov <- function(coord, semivario,...){
  sapply(1:nrow(coord), function(i) sapply(1:nrow(coord), function(j)
    semivario(sqrt(sum((coord[i,]^2)), ...)) +
    semivario(sqrt(sum((coord[j,]^2)), ...)) -
    semivario(sqrt(sum((coord[i,]-coord[j,])^2)), ...)))
}
rmev(n=100, sigma=vario2cov(grid.coord, semivario = semivario, scale = 0.5, alpha = 1), model='br')
# asymmetric logistic model - see function 'rmvevd' from package 'evd '
asy <- list(0, 0, 0, 0, c(0,0), c(0,0), c(0,0), c(0,0), c(0,0), c(0,0),
  c(.2,.1,.2), c(.1,.1,.2), c(.3,.4,.1), c(.2,.2,.2), c(.4,.6,.2,.5))
rmev(n=1, d=4, param=0.3, asy=asy, model="alog")
#Example with a grid (generating an array)
rmev(n=10, sigma=cbind(c(2,1), c(1,3)), coord=cbind(runif(4), runif(4)), model='smith', grid=TRUE)
## Example with Dirichlet mixture
alpha.mat <- cbind(c(2,1,1),c(1,2,1),c(1,1,2))
rmev(n=100, param=alpha.mat, weights=rep(1/3,3), model='dirmix')
rmev(n=10, param=c(0.1,1,2,3), d=3, model='pairbeta')

```

rmevspec

Random samples from spectral distributions of multivariate extreme value models.

Description

Generate from Q_i , the spectral measure of a given multivariate extreme value model based on the L1 norm.

Usage

```

rmevspec(
  n,
  d,
  param,
  sigma,
  model = c("log", "neglog", "bilog", "negbilog", "hr", "br", "xstud", "smith",
    "schlather", "ct", "sdir", "dirmix", "pairbeta", "pairexp", "wdirbs", "wexpbs",
    "maxlin"),
  weights = NULL,
  vario = NULL,
  coord = NULL,
  grid = FALSE,
  dist = NULL,
  ...
)

```

Arguments

| | |
|---------|--|
| n | number of observations |
| d | dimension of sample |
| param | parameter vector for the logistic, bilogistic, negative bilogistic and extremal Dirichlet (Coles and Tawn) model. Parameter matrix for the Dirichlet mixture. Degree of freedoms for extremal student model. See Details . |
| sigma | covariance matrix for Brown-Resnick and extremal Student-t distributions. Symmetric matrix of squared coefficients λ^2 for the Husler-Reiss model, with zero diagonal elements. |
| model | for multivariate extreme value distributions, users can choose between 1-parameter logistic and negative logistic, asymmetric logistic and negative logistic, bilogistic, Husler-Reiss, extremal Dirichlet model (Coles and Tawn) or the Dirichlet mixture. Spatial models include the Brown-Resnick, Smith, Schlather and extremal Student max-stable processes. Max linear models are also supported |
| weights | vector of length m for the m mixture components that sum to one. For the "maxlin" model, weights should be a matrix with d columns that represent the weight of the components and whose column sum to one (if provided, this argument overrides asy). |
| vario | semivariogram function whose first argument must be distance. Used only if provided in conjunction with coord and if sigma is missing |
| coord | d by k matrix of coordinates, used as input in the variogram vario or as parameter for the Smith model. If grid is TRUE, unique entries should be supplied. |
| grid | Logical. TRUE if the coordinates are two-dimensional grid points (spatial models). |
| dist | symmetric matrix of pairwise distances. Default to NULL. |
| ... | additional arguments for the vario function |

Details

The vector param differs depending on the model

- log: one dimensional parameter greater than 1
- neglog: one dimensional positive parameter
- bilog: d-dimensional vector of parameters in $[0, 1]$
- negbilog: d-dimensional vector of negative parameters
- ct, dir, negdir: d-dimensional vector of positive (a)symmetry parameters. Alternatively, a $d + 1$ vector consisting of the d Dirichlet parameters and the last entry is an index of regular variation in $(0, 1]$ treated as scale
- xstud: one dimensional parameter corresponding to degrees of freedom alpha
- dirmix: d by m-dimensional matrix of positive (a)symmetry parameters
- pairbeta, pairexp: $d(d-1)/2+1$ vector of parameters, containing the concentration parameter and the coefficients of the pairwise beta, in lexicographical order e.g., $\beta_{1,2}, \beta_{1,3}, \dots$
- wdirbs, wexpbs: 2d vector of d concentration parameters followed by the d Dirichlet parameters

Value

an n by d exact sample from the corresponding multivariate extreme value model

Note

This functionality can be useful to generate for example Pareto processes with marginal exceedances.

Author(s)

Leo Belzile

References

Dombry, Engelke and Oesting (2016). Exact simulation of max-stable processes, *Biometrika*, **103**(2), 303–317.

Boldi (2009). A note on the representation of parametric models for multivariate extremes. *Extremes* **12**, 211–218.

Examples

```
set.seed(1)
rmevspec(n=100, d=3, param=2.5, model='log')
rmevspec(n=100, d=3, param=2.5, model='neglog')
rmevspec(n=100, d=4, param=c(0.2,0.1,0.9,0.5), model='bilog')
rmevspec(n=100, d=2, param=c(0.8,1.2), model='ct') #Dirichlet model
rmevspec(n=100, d=2, param=c(0.8,1.2,0.5), model='sdir') #with additional scale parameter
#Variogram gamma(h) = scale*||h||^alpha
#NEW: Variogram must take distance as argument
vario <- function(x, scale=0.5, alpha=0.8){ scale*x^alpha }
#grid specification
grid.coord <- as.matrix(expand.grid(runif(4), runif(4)))
rmevspec(n=100, vario=vario, coord=grid.coord, model='br')
## Example with Dirichlet mixture
alpha.mat <- cbind(c(2,1,1),c(1,2,1),c(1,1,2))
rmevspec(n=100, param=alpha.mat, weights=rep(1/3,3), model='dirmix')
```

 rmnorm

Multivariate Normal distribution sampler

Description

Sampler derived using the eigendecomposition of the covariance matrix Σ . The function uses the Armadillo random normal generator

Usage

```
rmnorm(n, mu, Sigma)
```

Arguments

| | |
|-------|--|
| n | sample size |
| mu | mean vector. Will set the dimension |
| Sigma | a square covariance matrix, of same dimension as mu. No sanity check is performed to validate that the matrix is positive definite, so use at own risk |

Value

an n sample from a multivariate Normal distribution

Examples

```
rmnorm(n = 10, mu = c(0,2), Sigma = diag(2))
```

| | |
|-------|---|
| rparp | <i>Simulation from R-Pareto processes</i> |
|-------|---|

Description

Simulation from R-Pareto processes

Usage

```
rparp(
  n,
  shape = 1,
  risk = c("sum", "site", "max", "min", "l2"),
  siteindex = NULL,
  d,
  param,
  sigma,
  model = c("log", "neglog", "bilog", "negbilog", "hr", "br", "xstud", "smith",
    "schlather", "ct", "sdir", "dirmix"),
  weights,
  vario,
  coord = NULL,
  ...
)
```

Arguments

| | |
|-----------|--|
| n | number of observations |
| shape | shape tail index of Pareto variable |
| risk | string indicating risk functional. |
| siteindex | integer between 1 and d specifying the index of the site or variable |

| | |
|---------|--|
| d | dimension of sample |
| param | parameter vector for the logistic, bilogistic, negative bilogistic and extremal Dirichlet (Coles and Tawn) model. Parameter matrix for the Dirichlet mixture. Degree of freedoms for extremal student model. See Details . |
| sigma | covariance matrix for Brown-Resnick and extremal Student-t distributions. Symmetric matrix of squared coefficients λ^2 for the Husler-Reiss model, with zero diagonal elements. |
| model | for multivariate extreme value distributions, users can choose between 1-parameter logistic and negative logistic, asymmetric logistic and negative logistic, bilogistic, Husler-Reiss, extremal Dirichlet model (Coles and Tawn) or the Dirichlet mixture. Spatial models include the Brown-Resnick, Smith, Schlather and extremal Student max-stable processes. Max linear models are also supported |
| weights | vector of length m for the m mixture components that sum to one. For the "maxlin" model, weights should be a matrix with d columns that represent the weight of the components and whose column sum to one (if provided, this argument overrides asy). |
| vario | semivariogram function whose first argument must be distance. Used only if provided in conjunction with coord and if sigma is missing |
| coord | d by k matrix of coordinates, used as input in the variogram vario or as parameter for the Smith model. If grid is TRUE, unique entries should be supplied. |
| ... | additional arguments for the vario function |

Details

For riskf=max and riskf=min, the procedure uses rejection sampling based on Pareto variates sampled from sum and may be slow if d is large.

Value

an n by d sample from the R-Pareto process, with attributes `accept.rate` if the procedure uses rejection sampling.

Examples

```
rparp(n=10, risk = 'site', siteindex=2, d=3, param=2.5, model='log')
rparp(n=10, risk = 'min', d=3, param=2.5, model='neglog')
rparp(n=10, risk = 'max', d=4, param=c(0.2,0.1,0.9,0.5), model='bilog')
rparp(n=10, risk = 'sum', d=3, param=c(0.8,1.2,0.6, -0.5), model='sdir')
vario <- function(x, scale=0.5, alpha=0.8){ scale*x^alpha }
grid.coord <- as.matrix(expand.grid(runif(4), runif(4)))
rparp(n=10, risk = 'max', vario=vario, coord=grid.coord, model='br')
```

rparpcs

*Simulation from Pareto processes using composition sampling***Description**

The algorithm performs forward sampling by simulating first from a mixture, then sample angles conditional on them being less than (max) or greater than (min) one. The resulting sample from the angular distribution is then multiplied by Pareto variates with tail index shape.

Usage

```
rparpcs(
  n,
  model = c("log", "neglog", "br", "xstud"),
  risk = c("max", "min"),
  param = NULL,
  d,
  Lambda = NULL,
  Sigma = NULL,
  df = NULL,
  shape = 1,
  ...
)
```

Arguments

| | |
|--------|--|
| n | sample size. |
| model | string indicating the model family. |
| risk | string indicating the risk functional. Only max and min are currently supported. |
| param | parameter value for the logistic or negative logistic model |
| d | dimension of the multivariate model, only needed for logistic or negative logistic models |
| Lambda | parameter matrix for the Brown–Resnick model. See Details . |
| Sigma | correlation matrix if model = 'xstud', otherwise the covariance matrix formed from the stationary Brown-Resnick process. |
| df | degrees of freedom for extremal Student process. |
| shape | tail index of the Pareto variates (reciprocal shape parameter). Must be strictly positive. |
| ... | additional parameters, currently ignored |

Details

For the moment, only exchangeable models and models based on elliptical processes are handled.

The parametrization of the Brown–Resnick is in terms of the matrix Lambda, which is formed by evaluating the semivariogram γ at sites s_i, s_j , meaning that $\Lambda_{i,j} = \gamma(s_i, s_j)/2$.

The argument Sigma is ignored for the Brown-Resnick model if Lambda is provided by the user.

Value

an n by d matrix of samples, where $d = \text{ncol}(\text{Sigma})$, with attributes `mixt.weights`.

Author(s)

Leo Belzile

See Also

[rparp](#) for general simulation of Pareto processes based on an accept-reject algorithm.

Examples

```
## Not run:
##Brown-Resnick, Wadsworth and Tawn (2014) parametrization
D <- 20L
coord <- cbind(runif(D), runif(D))
semivario <- function(d, alpha = 1.5, lambda = 1){0.5 * (d/lambda)^alpha}
Lambda <- semivario(as.matrix(dist(coord))) / 2
rparpcs(n = 10, Lambda = Lambda, model = 'br', shape = 0.1)
##Extremal Student
Sigma <- stats::rWishart(n = 1, df = 20, Sigma = diag(10))[, , 1]
rparpcs(n = 10, Sigma = cov2cor(Sigma), df = 3, model = 'xstud')

## End(Not run)
```

rparpcshr

Simulation of generalized Huesler-Reiss Pareto vectors via composition sampling

Description

Sample from the generalized Pareto process associated to Huesler-Reiss spectral profiles. For the Huesler-Reiss Pareto vectors, the matrix `Sigma` is utilized to build Q viz.

$$Q = \Sigma^{-1} - \frac{\Sigma^{-1} \mathbf{1}_d \mathbf{1}_d^\top \Sigma^{-1}}{\mathbf{1}_d^\top \Sigma^{-1} \mathbf{1}_d}.$$

The location vector `m` and `Sigma` are the parameters of the underlying log-Gaussian process.

Usage

```
rparpcshr(n, u, alpha, Sigma, m)
```

Arguments

| | |
|-------|---|
| n | sample size |
| u | vector of marginal location parameters (must be strictly positive) |
| alpha | vector of shape parameters (must be strictly positive). |
| Sigma | covariance matrix of process, used to define Q . See Details . |
| m | location vector of Gaussian distribution. |

Value

n by d matrix of observations

References

Ho, Z. W. O and C. Dombry (2019), Simple models for multivariate regular variations and the Huesler-Reiss Pareto distribution, *Journal of Multivariate Analysis* (**173**), p. 525-550, doi:[10.1016/j.jmva.2019.04.008](https://doi.org/10.1016/j.jmva.2019.04.008)

Examples

```
D <- 20L
coord <- cbind(runif(D), runif(D))
di <- as.matrix(dist(rbind(c(0, ncol(coord)), coord)))
semivario <- function(d, alpha = 1.5, lambda = 1){(d/lambda)^alpha}
Vmat <- semivario(di)
Sigma <- outer(Vmat[-1, 1], Vmat[1, -1], '+') - Vmat[-1, -1]
m <- Vmat[-1,1]
## Not run:
samp <- rparpcshr(n = 100, u = c(rep(1, 10), rep(2, 10)),
  alpha = seq(0.1, 1, length = 20), Sigma = Sigma, m = m)

## End(Not run)
```

rrlarg

Simulate r-largest observations from point process of extremes

Description

Simulate the r-largest observations from a Poisson point process with intensity

$$\Lambda(x) = (1 + \xi(x - \mu)/\sigma)^{-1/\xi}$$

.

Usage

```
rrlarg(n, r, loc = 0, scale = 1, shape = 0)
```

Arguments

| | |
|-------|----------------------------------|
| n | sample size |
| r | number of observations per block |
| loc | location parameter |
| scale | scale parameter |
| shape | shape parameter |

Value

an n by r matrix of samples from the point process, ordered from largest to smallest in each row.

| | |
|-----------|--|
| shape.erm | <i>Exponential regression estimator of the shape</i> |
|-----------|--|

Description

This function implements the exponential regression estimator of the shape parameter for the case of Pareto tails with positive shape index.

Usage

```
shape.erm(xdat, k, method = c("bdgm", "fh"), bounds = NULL)
```

Arguments

| | |
|--------|---|
| xdat | vector of observations |
| k | vector of integer, the number of largest observations to consider |
| method | string; one of bdgm for the approach of Beirlant, Dierckx, Goegebeur and Matthys (1999) or fh for Feuerverger and Hall (1999) |
| bounds | vector of length 2 giving the bounds for rho, the second order parameter. Default to $\rho \in [-5, -0.5]$ |

Details

The second-order parameter is difficult to pin down, and while values within $[-1, 0)$ are most logical under Hall model, the model parameters become unidentifiable when $\rho \rightarrow 0$. The default constraint restrict $-5 < \rho < -0.5$, with the upper bound changed to -0.25 for sample of sizes larger than 5000 observations. Users can set the value of the bounds for ρ via argument bounds. The optimization is initialized at the Hill estimator.

Value

a data frame with columns

- k number of exceedances
- shape estimate of the shape parameter
- rho estimate of the second-order regular variation index
- scale estimate of the scale parameter

References

Feuerverger, A. and P. Hall (1999), Estimating a tail exponent by modelling departure from a Pareto distribution, *The Annals of Statistics* 27(2), 760-781. <doi:10.1214/aos/1018031215>

Beirlant, J., Dierckx, G., Goegebeur, Y. G. Matthys (1999). Tail Index Estimation and an Exponential Regression Model. *Extremes*, 2, 177–200 (1999). <doi:10.1023/A:1009975020370>

shape.genjack

Generalized jackknife shape estimator

Description

Generalized jackknife shape estimator

Usage

```
shape.genjack(xdat, k)
```

Arguments

xdat vector of positive observations

k vector of order statistics; if missing, a vector going from 10 to sample size minus one.

Value

a data frame with the number of order statistics k and the shape parameter estimate shape, or a single numeric value if k is a scalar integer.

References

Gomes, I.M., João Martins, M. and Neves, M. (2000) Alternatives to a Semi-Parametric Estimator of Parameters of Rare Events-The Jackknife Methodology. *Extremes*, 3, 207–229. [doi:10.1023/A:1011470010228](https://doi.org/10.1023/A:1011470010228)

| | |
|----------------|---|
| shape.genquant | <i>Beirlant et al. generalized quantile shape estimator</i> |
|----------------|---|

Description

This estimator estimates the real shape parameter based on generalized quantile plots based on mean excess functions, generalized median excesses or trimmed mean excesses.

Usage

```
shape.genquant(
  xdat,
  k,
  type = c("genmean", "genmed", "trimmean"),
  weight,
  p = 0.5
)
```

Arguments

| | |
|--------|--|
| xdat | n vector of observations |
| k | number of upper order statistics |
| type | string indicating the estimator choice, one of genmean, genmed and trimmean. |
| weight | weight a kernel function on $[0, 1]$ |
| p | number between zero and one giving the proportion of order statistics for the second threshold |

References

Beirlant, J., Vynckier P. and J.L. Teugels (1996). *Excess functions and estimation of the extreme-value index*. Bernoulli, 2(4), 293-318.

| | |
|------------|--|
| shape.hill | <i>Hill's estimator of the shape parameter</i> |
|------------|--|

Description

Given a sample of positive observations, calculate the tail index or shape parameter. The shape estimate returned is positive.

Usage

```
shape.hill(xdat, k)
```

Arguments

xdat vector of positive observations
 k vector of order statistics; if missing, a vector going from 10 to sample size minus one.

Value

a data frame with the number of order statistics k and the shape parameter estimate `shape`, or a single numeric value if k is a scalar integer.

References

Hill, B.M. (1975). *A simple general approach to inference about the tail of a distribution*. *Annals of Statistics*, **3**, 1163-1173.

Examples

```
xdat <- mev::rgp(n = 200, loc = 1, scale = 0.5, shape = 0.5)
shape.hill(xdat)
```

shape.lthill *Lower-trimmed Hill shape estimator*

Description

Given a sample of Pareto-tailed samples (positive tail index), compute the lower-trimmed Hill estimator. If $k_0 = k$, the estimator reduces to Hill's estimator for the shape index

Usage

```
shape.lthill(xdat, k, k0 = k, sorted = FALSE, ...)
```

Arguments

xdat [numeric] vector of positive observations
 k [integer] number of order statistics for the threshold
 k0 [integer] vector of number of largest order statistics, no greater than k
 sorted [logical] if TRUE, data are assumed to be sorted in decreasing order.
 ... additional arguments for other routines (notably `vectorize`)

Value

a scalar with the shape parameter estimate if k_0 is a scalar, otherwise a data frame with columns k_0 for the number of exceedances and `shape` for the tail index.

References

Bladt, M., Albrecher, H. & Beirlant, J. (2020) *Threshold selection and trimming in extremes*. *Extremes*, 23, 629-665 . doi:[10.1007/s10687020003850](https://doi.org/10.1007/s10687020003850)

Examples

```
# Pareto sample
n <- 200
xdat <- 10/(1 - runif(n)) - 10
shape.lthill(xdat = xdat, k = 100, k0 = 5:100)
```

shape.moment

Dekkers and de Haan moment estimator for the shape

Description

Given a sample of exceedances, compute the moment estimator of the real shape parameter.

Usage

```
shape.moment(xdat, k)
```

Arguments

xdat vector of positive observations of length n
k number of largest order statistics

Value

a data frame with the number of order statistics k and the shape parameter estimate `shape`, or a single numeric value if k is a scalar.

References

Dekkers, A.L.M. and de Haan, L. (1989). *On the estimation of the extreme-value index and large quantile estimation.*, *Annals of Statistics*, **17**, 1795-1833.

| | |
|-----------|---|
| shape.osz | <i>Extreme U-statistic Pickands shape estimator</i> |
|-----------|---|

Description

Given a random sample of n exceedances, the estimator returns an estimator of the shape parameter or extreme value index using a kernel of order 3, based on k largest exceedances of `xdat`. Note that the method does not allow for ties.

Usage

```
shape.osz(xdat, k, ...)
```

Arguments

| | |
|-------------------|---|
| <code>xdat</code> | vector of observations of length n |
| <code>k</code> | number of largest order statistics $3 \leq k < n$. |
| <code>...</code> | additional arguments for backward compatibility |

Details

The calculations are based on the recursions provided in Lemma 4.3 of Oorschot et al.

References

Oorschot, J, J. Segers and C. Zhou (2023), Tail inference using extreme U-statistics, Electronic Journal of Statistics 17(1): 1113-1159. doi:[10.1214/23EJS2129](https://doi.org/10.1214/23EJS2129)

Examples

```
xdat <- rgp(n = 1000, shape = 0.2)
shape.osz(xdat, k = 10)
```

| | |
|----------------|----------------------------------|
| shape.pickands | <i>Pickand's shape estimator</i> |
|----------------|----------------------------------|

Description

Given a sample of size n of positive exceedances, compute the real shape parameter ξ based on the k largest order statistics.

Usage

```
shape.pickands(xdat, k)
```

Arguments

xdat vector of positive observations of length n
 k number of largest order statistics

Value

a data frame with the number of order statistics k and the shape parameter estimate `shape`, or a single numeric value if k is a scalar.

References

Pickands, III, J. (1975). *Statistical inference using extreme order statistics*. *Annals of Statistics*, **3**, 119-131.

| | |
|-----------|---|
| shape.rbm | <i>Random block maxima shape estimator of Wager</i> |
|-----------|---|

Description

Computes the shape estimator for varying k up to sample size of maximum k_{\max} largest observations

The `plot` S3 routine returns plot of the shape estimator along with the value (and 95% Wald-based confidence interval) at the selected threshold, or a plot of the empirical Bayes risk.

Usage

```
shape.rbm(xdat, k = 10:floor(length(xdat)/2), ...)
```

```
## S3 method for class 'mev_shape_rbm'  
plot(x, type = c("shape", "risk"), log = TRUE, ...)
```

Arguments

xdat [vector] sample exceedances
 k [int] vector of integers for which to compute the estimator
 ... additional arguments, currently ignored
 x object of class `mev_shape_rbm` returned by `shape.rbm`
 type [string] type of plot, either "shape" for the tail index or "risk" for the empirical Bayes risk
 log [logical] if TRUE (default), the x-axis for the number of exceedances is displayed on the log scale.

Value

a list with elements

k number of exceedances

shape tail index estimate, strictly positive

risk empirical Bayes estimate of risk

thresh threshold given by the smallest order statistic considered in the sample

References

Wager, S. (2014). Subsampling extremes: From block maxima to smooth tail estimation, *Journal of Multivariate Analysis*, 130, 335-353, doi:10.1016/j.jmva.2014.06.010

| | |
|----------------|---|
| shape.trimhill | <i>Trimmed Hill estimator for the shape parameter</i> |
|----------------|---|

Description

Given a sample of Pareto-tailed samples (positive tail index), compute the trimmed Hill estimator. If $k_0 = k$, the estimator reduces to Hill's estimator for the shape index

Usage

```
shape.trimhill(xdat, k, k0, sorted = FALSE)
```

Arguments

| | |
|--------|--|
| xdat | [numeric] vector of positive observations |
| k | [integer] number of order statistics for the threshold |
| k0 | [integer] number of largest order statistics, strictly less than k |
| sorted | [logical] if TRUE, data are assumed to be sorted in decreasing order |

Value

a scalar with the shape parameter estimate

References

Bhattacharya, S., Kallitsis, M. and S. Stoev, (2019) Data-adaptive trimming of the Hill estimator and detection of outliers in the extremes of heavy-tailed data. *Electronic Journal of Statistics* 13, 1872–1925

| | |
|-------------|---------------------------------|
| shape.vries | <i>de Vries shape estimator</i> |
|-------------|---------------------------------|

Description

Given a sample of exceedances, compute the moment estimator of the positive shape parameter using the ratio of log ratio of exceedance and it's square.

Usage

```
shape.vries(xdat, k)
```

Arguments

| | |
|------|---|
| xdat | vector of positive observations of length n |
| k | number of largest order statistics |

Value

a data frame with the number of order statistics k and the shape parameter estimate `shape`, or a single numeric value if k is a scalar.

References

de Haan, L. and Peng, L. (1998). *Comparison of tail index estimators*, *Statistica Neerlandica* 52, 60-70.

de Haan, L. and Peng, L. (1998). Comparison of tail index estimators. *Statistica Neerlandica*, 52: 60-70. [doi:10.1111/14679574.00068](https://doi.org/10.1111/14679574.00068)

| | |
|--------|---|
| spunif | <i>Semi-parametric marginal transformation to uniform</i> |
|--------|---|

Description

The function `spunif` transforms a matrix or vector of data x to the pseudo-uniform scale using a semiparametric transform. Data below the threshold are transformed to pseudo-uniforms using a rank transform, while data above the threshold are assumed to follow a generalized Pareto distribution. The parameters of the latter are estimated using maximum likelihood if either `scale = NULL` or `shape = NULL`.

Usage

```
spunif(x, thresh, scale = NULL, shape = NULL)
```

Arguments

| | |
|--------|--|
| x | matrix or vector of data |
| thresh | vector of marginal thresholds |
| scale | vector of marginal scale parameters for the generalized Pareto |
| shape | vector of marginal shape parameters for the generalized Pareto |

Value

a matrix or vector of the same dimension as x, with pseudo-uniform observations

Author(s)

Leo Belzile

Examples

```
x <- rmev(1000, d = 3, param = 2, model = 'log')
thresh <- apply(x, 2, quantile, 0.95)
spunif(x, thresh)
```

taildep

Coefficient of tail correlation and tail dependence

Description

For data with unit Pareto margins, the coefficient of tail dependence η is defined via

$$\Pr(\min(X) > x) = L(x)x^{-1/\eta},$$

where $L(x)$ is a slowly varying function. Ignoring the latter, several estimators of η can be defined. In unit Pareto margins, η is a nonnegative shape parameter that can be estimated by fitting a generalized Pareto distribution above a high threshold. In exponential margins, η is a scale parameter and the maximum likelihood estimator of the latter is the Hill estimator. Both methods are based on peaks-over-threshold and the user can choose between pointwise confidence obtained through a likelihood ratio test statistic ("lrt") or the Wald statistic ("wald").

Usage

```
taildep(
  xdat,
  qllev = NULL,
  nq = 40,
  qlim = c(0.8, 0.99),
  depmeas = c("eta", "chi"),
  estimator = list(eta = c("emp", "betacop", "gpd", "hill"), chi = c("emp", "betacop")),
  confint = c("wald", "lrt"),
  level = 0.95,
```

```

trunc = TRUE,
margtrans = c("emp", "none"),
ties.method = "random",
plot = TRUE,
...
)

```

Arguments

| | |
|-------------|---|
| xdat | an n by d matrix of multivariate observations |
| qllev | vector of percentiles between 0 and 1 |
| nq | number of quantiles of the structural variable at which to form a grid; only used if $u = \text{NULL}$. |
| qlim | limits for the sequence u of the structural variable |
| depmeas | dependence measure, either of "eta" or "chi" |
| estimator | named list giving the estimation method for eta and chi. Default to "emp" for both. |
| confint | string indicating the type of confidence interval for η , one of "wald" or "lrt" |
| level | confidence level requested (default to 0.95). |
| trunc | logical indicating whether the estimates and confidence intervals should be truncated in $[0, 1]$ |
| margtrans | marginal transformation; if "none", data are assumed to be in uniform margins |
| ties.method | string indicating the type of method for rank; see rank for a list of options. Default to "random" |
| plot | logical; should graphs be plotted? |
| ... | additional arguments passed to plot; current support for main, xlab, ylab, add and further pch, lty, type, col for points; additional arguments for confidence intervals are handled via cipch, cilty, citype, cicol. |

Details

The most common approach for estimation is the empirical survival copula, by evaluating the proportion of sample minima with uniform margins that exceed a given x . An alternative estimator uses a smoothed estimator of the survival copula using Bernstein polynomial, resulting in the so-called betacop estimator. Approximate pointwise confidence intervals for the latter are obtained by assuming the proportion of points is binomial.

The coefficient of tail correlation χ is

$$\chi = \lim_{u \rightarrow 1} \frac{\Pr(F_1(X_1) > u, \dots, F_D(X_D) > u)}{1 - u}.$$

Asymptotically independent vectors have $\chi = 0$. The estimator uses an estimator of the survival copula

Value

a named list with elements

- qllev: a K vector of percentile levels
- eta: a K by 3 matrix with point estimates, lower and upper confidence intervals
- chi: a K by 3 matrix with point estimates, lower and upper confidence intervals

Note

As of version 1.15, the percentiles used are from the minimum variable. This ensures that, regardless of the number of variables, there is no error message returned because the quantile levels are too low for there to be observations

See Also

[chiplot](#) for bivariate empirical estimates of χ and $\bar{\chi}$.

Examples

```
## Not run:
set.seed(765)
# Max-stable model
dat <- rmev(n = 1000, d = 4, param = 0.7, model = "log")
taildep(dat, confint = 'wald')

## End(Not run)
```

| | |
|----------------|--|
| test.blocksize | <i>Likelihood ratio test for max-stability</i> |
|----------------|--|

Description

Given a matrix of block maxima split into blocks of size m, calculate test statistics and return p-values based on the asymptotic chi-square distribution.

Usage

```
test.blocksize(xdat, rounding = 0, alternative = c(1L, 2L, 3L), lb = NULL)
```

Arguments

| | |
|-------------|--|
| xdat | n by m matrix of observations assumed to arise from a GEV, ordered by row |
| rounding | double, a positive number indicating the amount of censoring (e.g., 0.1 or 1) |
| alternative | integer; 1 for a single shape parameter with parameters obtained from max-stability, 2 for a common shape parameter, but with free location and scale. |
| lb | lower bound for left-censoring; default to none (NULL) |

Value

a data frame containing likelihood ratio statistics (stat), the degrees of freedom, a vector of p-values pval and the name of the alternative.

Examples

```
samp <- build.blocks(mev::rgev(50, scale = 10), m = 4)
test.blocksize(xdat = round(samp, 0), rounding = 1, lb = -5)
test.blocksize(xdat = round(samp, 0), rounding = 1)
test.blocksize(xdat = samp)
test.blocksize(xdat = samp, lb = -5, alternative = 1L)
```

| | |
|--------------|---|
| test.maxstab | <i>P-P plot for testing max stability</i> |
|--------------|---|

Description

The diagnostic, proposed by Gabda, Towe, Wadsworth and Tawn, relies on the fact that, for max-stable vectors on the unit Gumbel scale, the distribution of the maxima is Gumbel distribution with a location parameter equal to the exponent measure. One can thus consider tuples of size m and estimate the location parameter via maximum likelihood and transforming observations to the standard Gumbel scale. Replicates are then pooled and empirical quantiles are defined. The number of combinations of m vectors can be prohibitively large, hence only n_{\max} randomly selected tuples are selected from all possible combinations. The confidence intervals are obtained by a nonparametric bootstrap, by resampling observations with replacement observations for the selected tuples and re-estimating the location parameter. The procedure can be computationally intensive as a result.

Usage

```
test.maxstab(
  xdat,
  m = prod(dim(dat)[-1]),
  nmax = 500L,
  B = 1000L,
  ties.method = "random",
  plot = TRUE,
  ...
)
```

Arguments

| | |
|------|--|
| xdat | matrix or array of max-stable observations, typically block maxima. The first dimension should consist of replicates |
| m | integer indicating how many tuples should be aggregated. |
| nmax | maximum number of pairs. Default to 500L. |

B number of nonparametric bootstrap replications. Default to 1000L.
 ties.method string indicating the method for [rank](#). Default to "random".
 plot logical indicating whether a graph should be produced (default to TRUE).
 ... additional arguments for backward compatibility

Value

a Tukey probability-probability plot with 95% confidence intervals obtained using a nonparametric bootstrap

References

Gabda, D.; Towe, R. Wadsworth, J. and J. Tawn, Discussion of "Statistical Modeling of Spatial Extremes" by A. C. Davison, S. A. Padoan and M. Ribatet. *Statist. Sci.* **27** (2012), no. 2, 189–192.

Examples

```
## Not run:
xdat <- mev::rmev(n = 250, d = 100, param = 0.5, model = "log")
test.maxstab(xdat, m = 100)
xdat <- rmnorm(n = 250, Sigma = diag(0.5, 10) + matrix(0.5, 10, 10), mu = rep(0, 10))
test.maxstab(xdat, m = 2, nmax = 100)
test.maxstab(xdat, m = ncol(xdat))

## End(Not run)
```

| | |
|-----------------|---|
| test.scoreindep | <i>Ramos and Ledford test of independence</i> |
|-----------------|---|

Description

The Ramos and Ledford (2005) score test of independence is a modification of tests by Tawn (1988) and Ledford and Tawn (1996) for a logistic model parameter $\alpha = 1$; the latter two have scores with zero expectation, but the variance of the score are infinite, which produces non-regularity and yield test, once suitably normalized, that converge slowly to their asymptotic null distribution. The test, designed for bivariate samples, transforms observations to have unit Frechet margins and considers a bivariate censored likelihood approach for the logistic distribution.

Usage

```
test.scoreindep(xdat, p, test = c("ledford", "tawn"))
```

Arguments

xdat a n by 2 matrix of observations
 p probability level for the marginal threshold
 test string; if tawn, only censor observations in the upper quadrant when both variables are large as in Tawn (1988), otherwise censor marginally for ledford as in Ledford and Tawn (1996).

Value

a list with elements

stat value of the score test statistic

pval asymptotic p-value

test test argument

Examples

```
samp <- rmev(n = 1000, d = 2,
  param = 0.99, model = "log")
(test.scoreindep(samp, p = 0.9))
```

| | |
|--------|--------------------------------------|
| thames | <i>Thames river flow at Kingston</i> |
|--------|--------------------------------------|

Description

Time series of annual maximum daily peak flow (in meter per seconds) of the Thames River at Kingston. The 1894 record was modified as the previous value reported of 1064 cubic meter per second was considered to be an overestimate of the true flow.

Usage

thames

Format

A data frame with 142 rows and 3 variables:

date date of measurement

flow double maximum daily river flow (in cubic meter per seconds)

flag logical; if TRUE, the value represents the instantaneous annual maximum, otherwise the natural annual max mean daily flow

Source

Acknowledgement: Data from the UK National River Flow Archive, <https://nrfa.ceh.ac.uk/data/station/info/39001>, extracted March 2026

thselect.alrs

Automatic L-moment ratio selection method

Description

Given a sample of observations, calculate the L-skewness and L-kurtosis over a set of candidate thresholds. For each threshold candidate, we find the L-skewness that minimizes the sum of squared distance between the theoretical L-skewness and L-kurtosis of the generalized Pareto distribution,

$$\min_{\tau_3} (t_3 - \tau_3)^2 + [t_4 - \tau_3(1 + 5\tau_3)/(5 + \tau_3)]^2.$$

The function returns the threshold with the minimum distance.

Usage

```
thselect.alrs(xdat, thresh, plot = FALSE)
```

Arguments

| | |
|--------|--|
| xdat | [numeric] vector of observations |
| thresh | [numeric] vector of candidate thresholds. If missing, 20 sample quantiles starting at the 0.25 quantile in increments of 3.75 percent. |
| plot | [logical] if TRUE, return a plot of the sample L-kurtosis against the L-skewness, along with the theoretical generalized Pareto curve. |

Value

scalar for the chosen numeric threshold

References

Silva Lomba, J., Fraga Alves, M.I. (2020). *L-moments for automatic threshold selection in extreme value analysis*. *Stoch Environ Res Risk Assess*, 34, 465–491. doi:[10.1007/s0047702001789x](https://doi.org/10.1007/s0047702001789x)

thselect.bab

Lower truncated Hill threshold selection

Description

Given a sample of positive data with Pareto tail, the algorithm computes the optimal number of order statistics that minimizes the variance of the average left truncated tail index estimator, and uses the relationship to the Hill estimator for the Hall class of distributions to derive the optimal number (minimizing the asymptotic mean squared error) of the Hill estimator. The default value for the second order regular variation index is taken to be $\rho = -1$.

Usage

```
thselect.bab(
  xdat,
  kmin = floor(0.2 * length(xdat)),
  kmax = length(xdat) - 1L,
  rho = -1,
  test = FALSE,
  nsim = 999L,
  level = 0.95
)
```

Arguments

| | |
|-------|---|
| xdat | [vector] positive vector of exceedances |
| kmin | [int] minimum number of exceedances |
| kmax | [int] maximum number of exceedances for the estimation of the shape parameter. |
| rho | [double] scalar for the second order regular variation index, a negative number. |
| test | [logical] if TRUE, computes the goodness-of-fit statistic for the model using Monte Carlo |
| nsim | [int] number of replications for Monte Carlo test, used only if test=TRUE. |
| level | [double] confidence level for test |

Value

a list with the number of order statistics for the Hill estimator, k_0 and the corresponding shape estimate shape, the average lower-trimmed Hill estimator shape.₁th and the number of order statistics upon which the latter is based, k_{0_1 th.

References

Bladt, M., Albrecher, H. & Beirlant, J. (2020) *Threshold selection and trimming in extremes*. *Extremes*, 23, 629-665 . doi:[10.1007/s10687020003850](https://doi.org/10.1007/s10687020003850)

 thselect.cbm

Threshold selection by shape mean square error minimization

Description

Use a semiparametric bootstrap to calculate the mean squared error of the shape parameter using maximum likelihood for different thresholds, and return the one that minimize the mean squared error.

Usage

```
thselect.cbm(xdat, thresh, B = 100)
```

Arguments

| | |
|--------|----------------------------------|
| xdat | vector of observations |
| thresh | vector of thresholds |
| B | number of bootstrap replications |

Value

an object of class `mev_thselect_cbm` containing

- `thresh`: ordered vector of candidate thresholds
- `thresh0`: selected threshold
- `shape`: shape parameter coefficient estimates at each threshold
- `nexc`: number of exceedances at each threshold
- `bias`: vector of bootstrap bias estimates
- `var`: vector of bootstrap variance estimates
- `mse`: vector of mean squared error bootstrap estimates

References

Caers, J., Beirlant, J. and Maes, M.A. (1999). Statistics for Modeling Heavy Tailed Distributions in Geology: Part I. Methodology. *Mathematical Geology*, 31, 391-410. <doi:10.1023/A:1007538624271>

Examples

```
set.seed(2025)
xdat <- rnorm(1000)
thresh <- qnorm(c(0.8, 0.9, 0.95))
thselect.cbm(xdat, thresh, B = 50)
```

thselect.cv

Threshold selection via coefficient of variation

Description

This function computes the empirical coefficient of variation and computes a weighted statistic comparing the squared distance with the theoretical coefficient variation corresponding to a specific shape parameter (estimated from the data using a moment estimator as the value minimizing the test statistic, or using maximum likelihood). The procedure stops if there are no more than 10 exceedances above the highest threshold.

Usage

```
thselect.cv(
  xdat,
  thresh,
  method = c("mle", "wcv", "cv"),
  nsim = 999L,
  nthresh = 10L,
  level = 0.05,
  lazy = FALSE,
  plot = FALSE
)
```

Arguments

| | |
|---------|--|
| xdat | [vector] vector of observations |
| thresh | [vector] vector of threshold. If missing, set to p^k for $k = 0$ to $k = \text{nthresh}$ |
| method | [string], either moment estimator for the (weighted) coefficient of variation (wcv and cv) or maximum likelihood (mle) |
| nsim | [integer] number of bootstrap replications |
| nthresh | [integer] number of thresholds, if thresh is not supplied by the user |
| level | [numeric] probability level for sequential testing procedure |
| lazy | [logical] compute the bootstrap p-value until the test stops rejecting at level level? Default to FALSE |
| plot | [logical] if TRUE, returns a plot of the p-value path |

Value

a list with elements

- thresh: value of threshold returned by the procedure, NA if the hypothesis is rejected at all thresholds
- thresh0: sorted vector of candidate thresholds
- cindex: index of selected threshold among thresh0 or NA if none returned
- pval: bootstrap p-values, with NA if lazy and the p-value exceeds level at lower thresholds
- shape: shape parameter estimates
- nexc: number of exceedances of each threshold thresh0
- method: estimation method for the shape parameter

Note

The authors suggest transformation of

$$Y = -1/(X + c) + 1/c,$$

where X are exceedances and $c = \sigma/\xi$ is the ratio of estimated scale and shape parameters. For heavy-tailed distributions with $\xi > 0.25$, this may be preferable, but must be conducted outside of the function.

References

del Castillo, J. and M. Padilla (2016). *Modelling extreme values by the residual coefficient of variation*, SORT, 40(2), pp. 303–320.

Examples

```
thselect.cv(
  xdat = rgp(1000),
  thresh = qgp(seq(0,0.9, by = 0.1)),
  nsim = 99,
  lazy = TRUE,
  plot = TRUE)
```

 thselect.egp

Threshold selection based on extended generalized Pareto models

Description

Fit an EGP model to data over a range of candidate thresholds `thresh` and perform likelihood-based tests of equality for $\kappa = c$, where $c = 1$ for all regular models and $c = 0$ for the 'gj-tnorm' and 'logist' models, for which the generalized Pareto special case corresponds to a value of κ occurring on the boundary of the parameter space.

Usage

```
thselect.egp(
  xdat,
  thresh,
  model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
    "logist"),
  type = c("wald", "lrt"),
  level = 0.95,
  transform = FALSE,
  plot = FALSE,
  ...
)
```

Arguments

| | |
|---------------------|---|
| <code>xdat</code> | vector of observations, greater than the threshold |
| <code>thresh</code> | threshold value |
| <code>model</code> | a string indicating which extended family to fit |
| <code>type</code> | choice of test statistic, either <code>wald</code> for Wald-based intervals, or <code>lrt</code> for profile likelihood ratio test. |
| <code>level</code> | [double] confidence interval level, default to 0.95. |

| | |
|-----------|---|
| transform | logical; if TRUE and type="wald", intervals for kappa are computed on the log-scale and back-transformed. |
| plot | [logical] if TRUE, return a plot of p-values against threshold |
| ... | additional arguments, passed to plotting routine |

Details

The threshold selection procedure returns chi-square statistics (`stat`) for Wald or profile likelihood ratio tests, along with p-values (`pval`) obtained from large sample distribution. The threshold returned is the lowest for which all further higher thresholds fail to reject the null hypothesis of $\kappa = c$, or equivalently of generalized Pareto tail.

Value

an invisible list of class `mev_thselect_egp` with elements

- `thresh`: vector of threshold candidates
- `thresh0`: selected threshold among candidates
- `coef`: vector of parameter estimates for κ
- `stat`: squared version of the test statistic
- `pval`: p-value obtained from the χ_1^2 approximation
- `level`: level of the confidence intervals
- `model`: string giving the EGP model family
- `type`: type of confidence interval

Examples

```
ths <- thselect.egp(
  xdat = rexp(1000),
  thresh = qexp(c(0.8, 0.9, 0.95)),
  model = "pt-power")
print(ths)
plot(ths)
```

thselect.expgqt

Generalized quantile threshold selection

Description

The methodology proposed by Beirlant, Vynckier and Teugels (1996) uses an asymptotic expansion of the mean squared error for Hill's estimator given a random sample with Pareto tails and positive shape, using an exponential regression. The value of k is selected to minimize the mean squared error given optimal weighting scheme. This depends on the order of regular variation ρ , which is obtained based on the slope of the difference in Hill estimators, suitably reweighted. The iterative procedure of Beirlant et al. alternates between parameter estimation until convergence. It returns

the generalized extreme value estimate, the Hill shape estimate, the number of higher order statistic, the parameter ρ and estimates of the standard error of the shape and the mean squared error, based on the ultimate parameter values. If the tail probability is provided, an estimate of the tail quantile at level $1 - p$ is also provided. Since the weights can become negative, there is no guarantee that the mean squared error estimate is positive, nor that the estimated value of ρ is nonpositive.

Usage

```
thselect.expgqt(
  xdat,
  maxiter = 10L,
  tol = 2,
  kmin = max(10, floor(length(xdat)/100)),
  kmax = floor(0.8 * length(xdat)),
  p = NULL,
  ...
)
```

Arguments

| | |
|---------|--|
| xdat | [vector] sample of exceedances |
| maxiter | [int] maximum number of iteration |
| tol | [double] tolerance for difference in value of k for the fixed point |
| kmin | [int] minimum number of exceedances for the estimator |
| kmax | [int] maximum number of exceedances for the estimator |
| p | [double] tail probability between 0 and 1/n (length of data). If provided, computes the tail quantile using the formula from Remark 2 of Beirlant and al. (2005) |
| ... | additional arguments, currently ignored |

Value

a list with components

- shape the exponential regression model shape estimator, based on the k_0 largest order statistics
- hill the Hill estimator of the shape, based on the k largest order statistics
- k_0 number of high order statistics for estimation of the shape using Hill's estimator
- rho estimate of the second order regular variation parameter
- mse mean squared error estimate of the shape parameter
- se standard error of the shape parameter
- convergence logical; if TRUE, indicates that the method converged to a fixed point within `tol` before reaching the maximum number of iterations `maxiter`
- p tail probability, if non-null.
- quantile tail quantile at probability level $1 - p$, if `p` is provided.

References

Beirlant, J., Vynckier, P., & Teugels, J. L. (1996). Excess Functions and Estimation of the Extreme-Value Index. *Bernoulli*, 2(4), 293–318. doi:10.2307/3318416

Beirlant, J., Dierckx, G., & Guillou, A. (2005). Estimation of the Extreme-Value Index and Generalized Quantile Plots. *Bernoulli*, 11(6), 949–970. <http://www.jstor.org/stable/25464774>

Examples

```
# Simulate Pareto data - log(xdat) is exponential with rate 2
xdat <- rgp(n = 200, loc = 1, scale = 0.5, shape = 0.5)
(thselect.expgqt(xdat))
```

| | |
|--------------|--|
| thselect.gbw | <i>Kernel-based threshold selection of Goegebeur, Beirlant and de Wet (2008)</i> |
|--------------|--|

Description

Kernel-based threshold selection of Goegebeur, Beirlant and de Wet (2008)

Usage

```
thselect.gbw(
  xdat,
  kmax,
  kernel = c("Jackson", "Lewis"),
  rho = c("gbw", "ghp", "fagh", "dk"),
  ...
)
```

Arguments

| | |
|--------|---|
| xdat | [vector] sample exceedances |
| kmax | [int] maximum number of exceedances considered |
| kernel | [string] kernel choice, one of Jackson or Lewis |
| rho | string for the estimator of the second order regular variation. Can also be a negative scalar |
| ... | additional arguments, for backward compatibility purposes |

Value

a list with elements

- k0: number of exceedances
- shape: Hill's shape estimate
- rho: second-order regular variation parameter estimate
- gof: goodness-of-fit statistic for the chosen threshold.

References

Goegebeur , Y., Beirlant , J., and de Wet , T. (2008). Linking Pareto-Tail Kernel Goodness-of-fit Statistics with Tail Index at Optimal Threshold and Second Order Estimation. REVSTAT-Statistical Journal, 6(1), 51–69. <doi:10.57805/revstat.v6i1.57>

Examples

```
xdat <- rgp(n = 1000, scale = 2, shape = 0.5)
(thselect.gbw(xdat, kmax = 500))
```

| | |
|---------------|--|
| thselect.goks | <i>Threshold selection based on weighted Kolmogorov-Smirnov distance</i> |
|---------------|--|

Description

Use a semiparametric bootstrap to calculate the null distribution of the weighted Kolmogorov-Smirnov difference between the generalized Pareto distribution and the empirical distribution

Usage

```
thselect.goks(xdat, thresh, test = TRUE, B = 100, eps = 0.5)
```

Arguments

| | |
|--------|--|
| xdat | vector of observations |
| thresh | vector of thresholds |
| test | logical; if TRUE, test goodness-of-fit via a parametric bootstrap from the fitted generalized Pareto distribution |
| B | number of bootstrap replications |
| eps | scalar between 0 and 0.5 giving the power of the number of exceedances. The default is Kolmogorov-Smirnov, and 0 returns Pickands (1975) method. |

Value

an object of class `mev_thselect_goks` containing

- `thresh`: ordered vector of candidate thresholds
- `thresh0`: selected threshold
- `coef`: scale and shape parameters
- `nexc`: number of exceedances at each threshold
- `stat`: vector of weighted Kolmogorov-Smirnov statistic
- `pval`: bootstrap p-value for the weighted Kolmogorov-Smirnov statistic at the selected threshold

References

Gonzalo, Jesus and Jose Olmo (2004). Which Extreme Values Are Really Extreme?, *Journal of Financial Econometrics*, 2(3), <doi:10.1093/jjfinec/nbh014>

Examples

```
set.seed(2025)
xdat <- rgp(n = 200, shape = 0.1)
thresh <- quantile(xdat, c(0.8,0.9,0.95))
thselect.goks(xdat, thresh, B = 50)
```

thselect.ksmd

Mahalanobis distance-based methodology

Description

Compute the Mahalanobis distance-based threshold method over a grid of thresholds by transforming data from generalized Pareto to unit exponential based on probability weighted moment estimates, then computing the first L-moment and the L-skewness. The latter are compared to the theoretical counterparts from a unit exponential sample of the same size, which is used to compute the Mahalanobis distance. The threshold returned is the one which minimizes the distance.

Usage

```
thselect.ksmd(xdat, thresh, approx = c("asymptotic", "mc"), nsim = 1000L)
```

Arguments

| | |
|--------|--|
| xdat | [numeric] vector of observations |
| thresh | [numeric] vector of candidate thresholds. If missing, 20 sample quantiles starting at the 0.25 quantile in increments of 3.75 percent. |
| approx | [string] method to use to obtain moments of first L-moment |
| nsim | [integer] number of replications for Monte Carlo approximation |

Value

a list with components

- thresh0: selected threshold returned by the procedure
- thresh: vector of candidate thresholds
- pval: scalar p -value for the chi-square approximation to the test statistic for the selected threshold
- dist: vector of Mahalanobis distance
- approx: type of approximation

References

Kiran, K. G. and Srivinas, V.V. (2021). *A Mahalanobis distance-based automatic threshold selection method for peaks over threshold model*. *Water Resources Research* 57. <doi:10.1029/2020WR027534>

| | |
|---------------|---|
| thselect.mdps | <i>Minimum distance threshold selection procedure</i> |
|---------------|---|

Description

Minimum distance threshold selection procedure

Usage

```
thselect.mdps(xdat)
```

Arguments

xdat vector of positive exceedances

Value

a list with components

k0 order statistic corresponding to threshold (number of exceedances)

shape Hill's estimator of the tail index based on k0 exceedances

thresh0 numerical value of the threshold, the n-k0+1 order statistic of the original sample

References

Clauset, A., Shalizi, C.R. and Newman, M.E.J. (2009). *Power-Law Distributions in Empirical Data*. *SIAM Review*. Society for Industrial and Applied Mathematics, **51**, 661-703, doi:10.1137/070710111

| | |
|--------------|---|
| thselect.mrl | <i>Automated mean residual life plots</i> |
|--------------|---|

Description

This function implements the automated proposal from Section 2.2 of Langousis et al. (2016) for mean residual life plots. It returns the threshold that minimize the weighted mean square error and moment estimators for the scale and shape parameter based on weighted least squares.

Usage

```
thselect.mrl(xdat, thresh, kmax, plot = TRUE, ...)
```

```
## S3 method for class 'mev_thselect_automrl'
plot(x, type = c("mrl", "mse"), ...)
```

Arguments

| | |
|--------|--|
| xdat | [numeric] vector of observations |
| thresh | [numeric] vector of thresholds; if missing, uses all order statistics from the 20th largest until kmax as candidates |
| kmax | [integer] maximum number of order statistics |
| plot | [logical] if TRUE (default), return a plot of the mean residual life plot with the fitted slope and the chosen threshold |
| ... | additional arguments, currently ignored |
| x | object of class mev_thselect_automrl |
| type | string indicating the response, either mean residual life or log of mean squared error |

Details

The procedure consists in estimating the usual mean residual life as a function of the threshold, and looking for an order statistic or threshold value above which the fit is more or less linear.

Value

a list containing

- thresh: candidate threshold vector
- thresh0: selected threshold
- scale: scale parameter estimate
- shape: shape parameter estimate
- mrl: empirical mean excess values
- xdat: ordered observations
- intercept: intercept for mean excess value at chosen threshold
- slope: slope for mean excess value at chosen threshold
- tmanual: logical; TRUE if the user passed a vector of thresholds

References

Langousis, A., A. Mamalakis, M. Puliga and R. Deidda (2016). *Threshold detection for the generalized Pareto distribution: Review of representative methods and application to the NOAA NCDC daily rainfall database*, Water Resources Research, **52**, 2659–2681.

Examples

```
thselect.mrl(rgp(n = 100))
```

| | |
|----------------|--|
| thselect.ncpgp | <i>Northop and Coleman piecewise generalized Pareto threshold selection diagnostic</i> |
|----------------|--|

Description

The model tests the null hypothesis of a generalized Pareto above each threshold in `thresh` against the alternative of a piecewise generalized Pareto model with continuity constraints.

Usage

```
thselect.ncpgp(xdat, thresh, test = "score", plot = FALSE, level = 0.95, ...)
```

Arguments

| | |
|---------------------|--|
| <code>xdat</code> | [vector] observations |
| <code>thresh</code> | [vector] candidate thresholds |
| <code>test</code> | [string] indicating whether to perform score test or likelihood ratio (<code>lr</code>) test. The latter requires fitting the alternative model, and so is more computationally expensive. |
| <code>plot</code> | [logical]; if TRUE, return a plot with the p-value path. |
| <code>level</code> | [double] confidence level for confidence interval, defaults to 0.95 |
| <code>...</code> | additional arguments, for backward compatibility purposes |

Value

an object of class `mev_thselect_ncpgp` containing the test statistic (`stat`), the p-values (`pval`), the threshold candidates (`thresh`) and the selected threshold (`thresh0`).

| | |
|--------------|--|
| thselect.pec | <i>Prediction error C-criterion threshold selection method</i> |
|--------------|--|

Description

This function computes the non-robust Pareto prediction error of Dupuis and Victoria-Feser (2003), termed C-criterion, for the Hill estimator of the shape parameter. The threshold returned is the value of the threshold, taken from order statistics, that minimizes the average prediction error.

Usage

```
thselect.pec(xdat, kmax)
```

Arguments

| | |
|------|---|
| xdat | vector of observations |
| kmax | maximum number of order statistics to consider. Default to sample size if left unspecified. |

Value

a list with the number of exceedances k , the chosen threshold thresh_0 and the corresponding Hill estimator shape estimate shape.

References

Dupuis, D.J. and M.-P. Victoria-Feser (2003). A Prediction Error Criterion for Choosing the Lower Quantile in Pareto Index Estimation, University of Geneva, technical report, <https://archive-ouverte.unige.ch/unige:5789>.

| | |
|-------------------|--|
| thselect.pickands | <i>Pickands' order statistics threshold selection method</i> |
|-------------------|--|

Description

Restricting to the largest fourth of the data, returns the number of exceedances that minimizes the Kolmogorov-Smirnov statistic, i.e., the maximum absolute difference between the estimated generalized Pareto and the empirical distribution of exceedances. Relative to the paper, different estimation methods are proposed.

Usage

```
thselect.pickands(xdat, thresh, method = c("mle", "lmom", "quartiles"))
```

Arguments

| | |
|--------|--|
| xdat | [numeric] vector of observations |
| thresh | [numeric] vector of candidate thresholds. If missing, defaults to order statistics from the 10th to a quarter of the sample size. |
| method | [string] estimation method, either the quartiles of Pickands (1975), maximum likelihood, probability weighted moments or L-moments |

Value

a list with components

- k_0 : number of exceedances
- thresh_0 : selected threshold returned by the procedure
- thresh: vector of candidate thresholds
- dist; vector of Kolmogorov-Smirnoff distance

- method; string for the estimation method
- scale: estimated scale parameter at the chosen threshold
- shape: estimated shape parameter at the chosen threshold

Note

The quartiles estimator of Pickands is robust, but very inefficient. It is provided for historical reasons.

References

James Pickands III (1975). *Statistical inference using extreme order statistics*, Annals of Statistics, 3(1) 119-131, doi:[10.1214/aos/1176343003](https://doi.org/10.1214/aos/1176343003)

See Also

[thselect.goks]

thselect.rbm

Threshold selection for the random block maxima method

Description

Threshold selection for the random block maxima method

Usage

```
thselect.rbm(xdat, kmax = length(xdat))
```

Arguments

| | |
|------|--|
| xdat | [vector] sample exceedances |
| kmax | maximum number of exceedances to consider. |

Value

a list with elements

- k_0 : the number of exceedances at the selected threshold
- thresh0: the selected threshold, or accordingly the $(k_0 + 1)$ th order statistic
- shape: the RBM shape estimate

| | |
|-----------------|---------------------------------------|
| thselect.samsee | <i>Threshold selection via SAMSEE</i> |
|-----------------|---------------------------------------|

Description

Smooth asymptotic mean squared error estimator of Schneider et al. (2021) for threshold selection. The implementation uses a second-order regular variation index of -1

Usage

```
thselect.samsee(xdat)
```

Arguments

xdat vector of positive exceedances

Value

a list with elements

- k0 optimal number of exceedances
- shape Hill estimator of the tail index
- thresh0 selected threshold

References

Schneider, L.F., Krajina, A. and Krivobokova, T. (2021). *Threshold selection in univariate extreme value analysis*, *Extremes*, **24**, 881-913 [doi:10.1007/s10687021004057](https://doi.org/10.1007/s10687021004057)

| | |
|----------------|--|
| thselect.sinfo | <i>Threshold selection diagnostic of Suveges and Davison</i> |
|----------------|--|

Description

The information matrix test (IMT), proposed by Suveges and Davison (2010), is based on the difference between the expected quadratic score and the second derivative of the log-likelihood. The asymptotic distribution for each threshold u and gap K is asymptotically χ^2 with one degree of freedom. The approximation is good for $N > 80$ and conservative for smaller sample sizes. The test assumes independence between gaps.

Usage

```
thselect.sinfo(xdat, thresh, qllev, plot = FALSE, kmax = 1, k = 1)
```

Arguments

| | |
|--------|---|
| xdat | [vector] vector of observations |
| thresh | [vector] candidate thresholds |
| qllev | [vector] probability levels to define threshold if thresh is missing. |
| plot | [logical]; should the graphical diagnostic be plotted? |
| kmax | [int] the largest K-gap under consideration for clusters |
| k | [int] the K-gap for automatic threshold selection |

Details

The procedure proposed in Suveges & Davison (2010) was corrected for erratas. The maximum likelihood is based on the limiting mixture distribution of the intervals between exceedances (an exponential with a point mass at zero). The condition $D^{(K)}(u_n)$ should be checked by the user.

Fukutome et al. (2015) propose an ad hoc automated procedure

1. Calculate the interexceedance times for each K-gap and each threshold, along with the number of clusters
2. Select the (u, K) pairs for which $IMT < 0.05$ (corresponding to a P-value of 0.82)
3. Among those, select the pair (u, K) for which the number of clusters is the largest

Value

an invisible list of class with elements

- thresh a vector of thresholds based on empirical quantiles at supplied levels.
- stat a matrix of test statistics
- pval a matrix of approximate p-values (corresponding to probabilities under a χ_1^2 distribution)
- mle a matrix of maximum likelihood estimates for each given pair of thresholds and gaps
- loglik a matrix of log-likelihood values at MLE for each given pair of elements in thresh and gap in $0, \dots, kmax$
- quantile quantile levels for thresholds, if supplied by the user
- kmax the largest gap number

Author(s)

Leo Belzile

References

Fukutome, Liniger and Suveges (2015), Automatic threshold and run parameter selection: a climatology for extreme hourly precipitation in Switzerland. *Theoretical and Applied Climatology*, **120**(3), 403-416.

Suveges and Davison (2010), Model misspecification in peaks over threshold analysis. *Annals of Applied Statistics*, **4**(1), 203-221.

White (1982), Maximum Likelihood Estimation of Misspecified Models. *Econometrica*, **50**(1), 1-25.

Examples

```
thselect.sinfo(
  xdat = rgp(n = 10000),
  qlcv = seq(0.1, 0.9, length = 10),
  kmax = 3)
```

| | |
|------------------|---|
| thselect.vmetric | <i>Metric-based threshold selection</i> |
|------------------|---|

Description

Adaptation of Varty et al.'s metric-based threshold automated diagnostic for the independent and identically distributed case with no rounding.

This S3 method produces quantile-quantile plots with confidence and tolerance bands on various scale (uniform, exponential, generalized Pareto), or a plot of the metric as a function of the threshold.

Usage

```
thselect.vmetric(
  xdat,
  thresh,
  B = 199L,
  type = c("eqd", "exp", "qq", "pp", "tails"),
  dist = c("l1", "l2"),
  uq = FALSE,
  bootstrap = c("nonparametric", "parametric"),
  pp = ppoints(250),
  level = 0.95,
  plot = FALSE,
  ...
)

## S3 method for class 'mev_thselect_vmetric'
plot(
  x,
  type = c("qq", "pp", "exp", "metric"),
  B = 1000L,
  probs = c(0.025, 0.975),
  ...
)
```

Arguments

| | |
|--------|------------------------|
| xdat | vector of observations |
| thresh | vector of thresholds |

| | |
|-----------|--|
| B | number of simulations for variability of estimation |
| type | string; a single string indicating the choice of plot |
| dist | string indicating norm, either l1 for absolute error or l2 for quadratic error |
| uq | logical; if TRUE, generate bootstrap samples accounting for the sampling distribution of parameters. Only valid when bootstrap = "parametric". |
| bootstrap | string, one of nonparametric (sampling with replacement from exceedances) or parametric (sampling from generalized Pareto). |
| pp | plotting positions for the uniform. If type = "tails", only the values exceeding the threshold probability level are kept. Default to 250 uniform plotting positions on the unit interval. |
| level | level of symmetric confidence interval. Default to 0.95 |
| plot | logical; if TRUE, returns a plot |
| ... | additional arguments, currently ignored |
| x | an object of class mev_thselect_vmetric produced by a call to thselect.vmetric |
| probs | quantile levels for intervals. |

Details

The algorithm proceeds by first computing the maximum likelihood algorithm and then simulating replication datasets using either a parametric or nonparametric bootstrap. For each bootstrap sample, we refit the model and convert the quantiles to exponential or uniform variates depending on type, or else if eqd by calculating the expected plotting positions for the simulated sample.

If uq = TRUE and we specify bootstrap = "parametric", the estimation uncertainty is taken into consideration and each sample is drawn from a generalized Pareto distribution, but with different parameters reflecting the sampling distribution.

The mean absolute or mean squared distance is calculated on each bootstrap sample at each threshold, and then aggregated into a single average at each thresh value. The threshold returned is the one with the lowest average value of the metric.

Collings et al. (2025) recommend to use quantile-quantile plot, but with pp starting from some minimal threshold and going no further than the $1 - 10/n$ probability level. This can be supplied via pp. When choosin type = "tails", only probability points exceeding the threshold level are kept, so the metric is evaluated at the same levels, but with fewer points, as we increase the threshold level.

Value

an invisible list with components

- thresh: scalar threshold minimizing criterion
- thresh0: vector of candidate thresholds
- metric: value of the metric criterion evaluated at each threshold
- type: argument type
- dist: argument dist,
- level: level of confidence interval, from level
- bootstrap: type of bootstrap, either parametric or nonparametric.

References

Varty, Z. and J.A. Tawn and P.M. Atkinson and S. Bierman (2021+), Inference for extreme earthquake magnitudes accounting for a time-varying measurement process.

Murphy, C., Tawn, J. A., & Varty, Z. (2024). *Automated Threshold Selection and Associated Inference Uncertainty for Univariate Extremes*. *Technometrics*, 67(2), 215–224. <doi:10.1080/00401706.2024.2421744>

Collings, T.P., C. Murphy-Barltrop, C. Murphy, I.D. Haigh, P.D. Bates, and N.D. Quinn (2025). *Automated tail-informed threshold selection for extreme coastal sea levels*, *Natural Hazards and Earth System Sciences*, 25(11), 4545–4562, <doi:10.5194/nhess-25-4545-2025>.

Examples

```
## Not run:
xdat <- rexp(1000, rate = 1/2)
thresh <- quantile(xdat, prob = c(0.25,0.5, 0.75))
# Method of Murphy, Tawn and Varty (2024) - EQD
thv <- thselect.vmetric(xdat, thresh, B = 99)
plot(thv)
plot(thv, type = "metric")
print(thv)
# TAILS method
tails <- thselect.vmetric(
  xdat,
  thresh = thresh,
  type = "tails",
  B = 99,
  pp = seq(0.8, 1-10/length(xdat), length.out = 250))

## End(Not run)
```

| | |
|---------------|---|
| thselect.wcvm | <i>Threshold selection via minimization of the weighted Cramer-von Mises distance</i> |
|---------------|---|

Description

For a Pareto-type sample, return the threshold that minimizes a weighted Cramer-von Mises criterion for the exponential sample with scale H_{n,n_u} and the log increments.

Usage

```
thselect.wcvm(xdat, k)
```

Arguments

| | |
|------|--|
| xdat | vector of positive exceedances |
| k | vector of number of exceedances, or integer indicating the maximum value of k , in which case a vector of integers from $k = 10$ to k is constructed |

Value

an object of class `mev_thselect_wcvm` (list) with elements

- `k0`: selected number of order statistics
- `shape`: Hill estimate of the shape at selected threshold
- `thresh`: value of the threshold (the $(k+1)$ st largest order statistic)
- `criterion`: a data frame with columns `k` and `crit` giving the criterion value

References

Goegebeur , Y., Beirlant , J., and de Wet , T. (2008). Linking Pareto-Tail Kernel Goodness-of-fit Statistics with Tail Index at Optimal Threshold and Second Order Estimation. *REVSTAT-Statistical Journal*, 6(1), 51-69. <doi:10.57805/revstat.v6i1.57>

tstab.cv

Coefficient of variation threshold stability plot

Description

This function calculates parametric estimates of the coefficient of variation with pointwise Wald confidence intervals along with empirical estimates and returns a threshold stability plot.

Usage

```
tstab.cv(
  xdat,
  thresh,
  method = c("empirical", "mle", "wcv", "cv"),
  nthresh = 10L,
  nsim = 99L,
  plot = TRUE,
  level = 0.95,
  ...
)
```

Arguments

| | |
|----------------------|--|
| <code>xdat</code> | [vector] vector of observations |
| <code>thresh</code> | [vector] vector of threshold. If missing, set to p^k for $k = 0$ to $k = \text{nthresh}$ |
| <code>method</code> | [string], either moment estimator for the (weighted) coefficient of variation (wcv and cv) or maximum likelihood (mle) |
| <code>nthresh</code> | [integer] number of thresholds, if <code>thresh</code> is not supplied by the user |
| <code>nsim</code> | [integer] number of bootstrap replications |
| <code>plot</code> | [logical] if TRUE, returns a plot of the p-value path |
| <code>level</code> | [numeric] probability level for sequential testing procedure |
| <code>...</code> | additional parameters, notably for package <code>boot</code> , for the type of confidence intervals. |

Examples

```
tstab.cv(
  xdat = rgp(1000),
  thresh = qgp(seq(0,0.9, by = 0.1)),
  method = "cv")
tstab.cv(
  xdat = rgp(1000),
  thresh = qgp(seq(0,0.9, by = 0.1)),
  method = "empirical")
```

tstab.egp

*Threshold stability plots for extended generalized Pareto models***Description**

Threshold stability plots for extended generalized Pareto models

Usage

```
tstab.egp(
  xdat,
  thresh,
  model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
    "logist"),
  param = c("shape", "kappa"),
  type = c("wald", "lrt"),
  transform = FALSE,
  level = 0.95,
  plot = TRUE,
  ...
)
```

Arguments

| | |
|-----------|---|
| xdat | vector of observations, greater than the threshold |
| thresh | threshold value |
| model | a string indicating which extended family to fit |
| param | [string] parameter, either shape or additional parameter kappa |
| type | [string] confidence interval type, either wald or profile. |
| transform | logical; if TRUE and type="wald", intervals for kappa are computed on the log-scale and back-transformed. |
| level | [double] confidence interval level, default to 0.95. |
| plot | [logical] if TRUE (default), return a threshold stability plot |
| ... | additional arguments for the plot function, currently ignored. |

Value

an invisible list object of class `mev_egp_tstab` with elements

- `kappa`: matrix of parameter estimates and confidence intervals for κ , if specified in `param`
- `shape`: matrix of parameter estimates and confidence intervals for the shape parameter ξ , if specified in `param`
- `thresh`: vector of threshold candidates
- `level`: level of the confidence intervals
- `model`: string giving the EGP model family
- `type`: type of confidence interval

Examples

```
xdat <- rgp(n = 1000)
tstab.egp(
  xdat = xdat,
  thresh = c(0, quantile(xdat, 0.5)),
  model = "gj-tnorm",
  param = "kappa",
  transform = TRUE)
```

tstab.gpd

Parameter stability plots for peaks-over-threshold

Description

This function computes the maximum likelihood estimate at each provided threshold and plots the estimates (pointwise), along with 95% confidence/credible intervals obtained using Wald or profile confidence intervals, or else from 1000 independent draws from the posterior distribution under vague independent normal prior on the log-scale and shape. The latter two methods better reflect the asymmetry of the estimates than the Wald confidence intervals.

Usage

```
tstab.gpd(
  xdat,
  thresh,
  method = c("wald", "lrt", "post"),
  level = 0.95,
  plot = TRUE,
  which = c("scale", "shape"),
  changepar = TRUE,
  ...
)
```

Arguments

| | |
|------------|--|
| xdat | a vector of observations |
| thresh | a vector of candidate thresholds at which to compute the estimates. |
| method | string indicating the method for computing confidence or credible intervals. Must be one of "wald", "profile" or "post". |
| level | confidence level of the intervals. Default to 0.95. |
| plot | logical; should parameter stability plots be displayed? Default to TRUE. |
| which | character vector with elements scale or shape |
| changeapar | logical; if TRUE, changes the graphical parameters. |
| ... | additional arguments passed to plot. |

Value

a list with components

- threshold: vector of numerical threshold values.
- mle: matrix of modified scale and shape maximum likelihood estimates.
- lower: matrix of lower bounds for the confidence or credible intervals.
- upper: matrix of lower bounds for the confidence or credible intervals.
- method: method for the confidence or coverage intervals.

plots of the modified scale and shape parameters, with pointwise confidence/credible intervals and an invisible data frame containing the threshold thresh and the modified scale and shape parameters.

Note

The function is hard coded to prevent fitting a generalized Pareto distribution to samples of size less than 10. If the estimated shape parameters are all on the boundary of the parameter space (meaning $\hat{\xi} = -1$), then the plots return one-sided confidence intervals for both the modified scale and shape parameters: these typically suggest that the chosen thresholds are too high for estimation to be reliable.

Author(s)

Leo Belzile

See Also

[gpd.fitrange](#)

Examples

```

dat <- abs(rnorm(10000))
u <- qnorm(seq(0.9,0.99, by= 0.01))
par(mfrow = c(1,2))
tstab.gpd(xdat = dat, thresh = u, changepar = FALSE)
## Not run:
tstab.gpd(xdat = dat, thresh = u, method = "lrt")
tstab.gpd(xdat = dat, thresh = u, method = "post")

## End(Not run)

```

tstab.hill

Threshold stability plot for Hill estimator

Description

Threshold stability plot for Hill estimator

Usage

```
tstab.hill(xdat, kmax, method = "hill", ..., log = TRUE)
```

Arguments

| | |
|--------|---|
| xdat | [vector] sample exceedances |
| kmax | [int] maximum number of order statistics |
| method | [string] name of estimator for shape parameter. Default to hill. |
| ... | additional arguments passed to <code>fit.shape</code> for certain methods. |
| log | [logical] should the x-axis for the number of order statistics used for estimation be displayed on the log scale? Default to TRUE |

Value

a plot of shape estimates as a function of the number of exceedances

Examples

```

xdat <- rgp(n = 250, loc = 1, scale = 2, shape = 0.5)
tstab.hill(xdat)

```

`tstab.lthill`*Threshold stability plots for left-truncated Hill estimators*

Description

Given a vector of exceedances and some potential choices of k for the threshold, compute the left-truncated Hill estimators for each value of k and use these to compute the variance and slope of the estimator

Usage

```
tstab.lthill(xdat, k, which = c("lthill", "var", "slope"), log = TRUE, ...)
```

Arguments

| | |
|--------------------|--|
| <code>xdat</code> | [numeric] vector of positive observations |
| <code>k</code> | [integer] number of order statistics for the threshold |
| <code>which</code> | [string] the type of plot, showing the left-truncated Hill plot on the log, the log of the variance of the estimator, or the log slope |
| <code>log</code> | [logical] if TRUE (default), shows the Hill plot on the log-scale |
| <code>...</code> | additional parameters for color, etc. to be passed to plot |

Value

an invisible list with `lthill`, order statistics, the log variance and the log scale.

References

Bladt, M., Albrecher, H. & Beirlant, J. (2020) *Threshold selection and trimming in extremes*. *Extremes*, 23, 629-665 . [doi:10.1007/s10687020003850](https://doi.org/10.1007/s10687020003850)

Examples

```
xdat <- 10/(1 - runif(n = 1000)) - 10
tstab.lthill(xdat = xdat, k = c(50,100,200))
```

tstab.mrl

*Mean residual life plot***Description**

Computes mean of sample exceedances over a range of thresholds or for a pre-specified number of largest order statistics, and returns a plot with 95% Wald-based confidence intervals as a function of either the threshold or the number of exceedances. The main purpose is the plotting method, which generates the so-called mean residual life plot. The latter should be approximately linear over the threshold for a generalized Pareto distribution

Usage

```
tstab.mrl(
  xdat,
  thresh,
  kmin = 10L,
  kmax = length(xdat),
  plot = TRUE,
  level = 0.95,
  xlab = c("thresh", "nexc"),
  type = c("band", "ptwise"),
  ...
)
```

Arguments

| | |
|--------|--|
| xdat | vector of sample observations |
| thresh | vector of thresholds |
| kmin | integer giving the minimum number of exceedances; ignored if thresh is provided. Default to 10 |
| kmax | integer giving the maximum number of exceedances; ignored if thresh is provided. Default to sample size. |
| plot | logical; if TRUE, call the plot method |
| level | double giving the level of confidence intervals for the plot, default to 0.95 |
| xlab | string indicating whether to use thresholds (thresh) or number of largest order statistics (nexc) for the x-axis |
| type | string whether to plot pointwise confidence intervals using segments ("ptwise") or using dashed lines ("band") |
| ... | additional arguments, currently ignored |

Value

an invisible list with mean sample exceedances and standard deviation, number of exceedances, threshold

References

Davison, A.C. and R.L. Smith (1990). Models for Exceedances over High Thresholds (with discussion), *Journal of the Royal Statistical Society. Series B (Methodological)*, **52**(3), 393–442.

Examples

```
tstab.mrl(
  xdat = rgp(n = 100, shape = -0.5),
  xlab = "thresh",
  kmax = 50)
tstab.mrl(
  rexp(100),
  thresh = qexp(seq(0, 0.9, by = 0.01)))
```

venice

Venice Sea Levels

Description

The venice data contains the 10 largest yearly sea levels (in cm) from 1887 until 2019. Only the yearly maximum is available for 1922 and the six largest observations for 1936.

Format

a data frame with 133 rows and 11 columns containing the year of the measurement (first column) and ordered 10-largest yearly observations, reported in decreasing order from largest (r1) to smallest (r10).

Note

Smith (1986) notes that the annual maxima seems to fluctuate around a constant sea level up to 1930 or so, after which there is potential linear trend. Records of threshold exceedances above 80 cm (reported on the website) indicate that observations are temporally clustered.

The observations from 1931 until 1981 can be found in Table 1 in Smith (1986), who reported data from Pirazzoli (1982). The values from 1983 until 2019 were extracted by Anthony Davison from the City of Venice website (accessed in May 2020) and are licensed under the CC BY-NC-SA 3.0 license. The Venice City website indicates that later measurements were recorded by an instrument located in Punta Salute.

Source

City of Venice, Historical archive <<https://www.comune.venezia.it/node/6214>>. Last accessed November 5th, 2020.

References

- Smith, R. L. (1986) Extreme value theory based on the r largest annual events. *Journal of Hydrology* **86**, 27–43.
- Pirazzoli, P., 1982. Maree estreme a Venezia (periodo 1872-1981). *Acqua Aria* **10**, 1023-1039.
- Coles, S. G. (2001) *An Introduction to Statistical Modelling of Extreme Values*. London: Springer.

See Also

[venice](#)

w1500m

Best 200 times of Women 1500m Track

Description

200 all-time best performance (in seconds) of women 1500-meter run.

Format

a vector of size 200

Source

<http://www.alltime-athletics.com/w_1500ok.htm>, accessed 14.08.2018

xacf

Extremogram

Description

Given a regular time series of observations, compute the pairwise tail correlation between series at different lags. Permutation-based resampling are used to construct confidence envelope (one-sided) for comparison with the independent setting if `confint = TRUE`.

Usage

```
xacf(  
  x,  
  qllev,  
  lag.max = NULL,  
  plot = TRUE,  
  confint = FALSE,  
  B = 100L,  
  level = 0.95,  
  ties.method = "random",  
  na.action = na.fail  
)
```

Arguments

| | |
|--------------------------|---|
| <code>x</code> | vector of observations or time series |
| <code>qllev</code> | quantile level of threshold, a scalar between (0,1) |
| <code>lag.max</code> | integer, maximum lag at which to calculate the extremogram. Default to $10 \log_{10}(n)$ |
| <code>plot</code> | logical; if TRUE, return a plot of the extremogram |
| <code>confint</code> | logical; if TRUE, calculate level pointwise confidence intervals under independence, using a permutation-based approach |
| <code>B</code> | integer, number of simulations for <code>confint</code> |
| <code>level</code> | confidence level requested (default to 0.95). |
| <code>ties.method</code> | string indicating the type of method for rank; see rank for a list of options. Default to "random" |
| <code>na.action</code> | function to be called to handle missing values |

Value

a list with elements `extremogram` for the estimate of tail correlation at different lags, `upper` for the upper bound of the confidence interval for independent data and `level` of the latter.

References

- Davis, R. A., Mikosch, T., and Cribben, I. (2012). Towards estimating extremal serial dependence via the bootstrapped extremogram. *Journal of Econometrics*, 170(1), 142-152, [doi:10.1016/j.jeconom.2012.04.003](https://doi.org/10.1016/j.jeconom.2012.04.003).
- Davis, R. A. and T. Mikosch (2009). The extremogram: A correlogram for extreme events, *Bernoulli*, 15(4), 977-1009, [doi:10.3150/09BEJ213](https://doi.org/10.3150/09BEJ213).

Examples

```
xacf(x = rmar1(n = 1000, theta = 0.2, shape = 0.5),
     qllev = 0.95)
```

xdep.asym

Coefficient of extremal asymmetry

Description

This function implements estimators of the bivariate coefficient of extremal asymmetry proposed in Semadeni's (2021) PhD thesis. Two estimators are implemented: one based on empirical distributions, the second using empirical likelihood.

Usage

```
xdep.asym(
  xdat,
  qlev = NULL,
  nq = 40,
  qlim = c(0.8, 0.99),
  estimator = c("emp", "elik"),
  confint = c("none", "wald", "bootstrap"),
  level = 0.95,
  B = 999L,
  ties.method = "random",
  plot = TRUE,
  ...
)
```

Arguments

| | |
|-------------|---|
| xdat | an n by 2 matrix of observations |
| qlev | vector of quantile levels at which to evaluate extremal asymmetry |
| nq | integer; number of quantiles at which to evaluate the coefficient if u is NULL |
| qlim | a vector of length 2 with the probability limits for the quantiles |
| estimator | string indicating the estimation method, one of emp or empirical likelihood (elik) |
| confint | string for the method used to derive confidence intervals, either none (default) or a nonparametric bootstrap |
| level | probability level for confidence intervals, default to 0.95 or bounds for the interval |
| B | integer; number of bootstrap replicates (if applicable) |
| ties.method | string; method for handling ties. See the documentation of rank for available options. |
| plot | logical; if TRUE, return a plot. |
| ... | additional arguments for backward compatibility |

Details

Let U, V be uniform random variables and define the partial extremal dependence coefficients

$$\varphi_+(u) = \Pr(V > U \mid U > u, V > u),$$

,

$$\varphi_-(u) = \Pr(V < U \mid U > u, V > u),$$

$$\varphi_0(u) = \Pr(V = U \mid U > u, V > u).$$

Define

$$\varphi(u) = \frac{\varphi_+ - \varphi_-}{\varphi_+ + \varphi_-}$$

and the coefficient of extremal asymmetry as $\varphi = \lim_{u \rightarrow 1} \varphi(u)$.

The empirical likelihood estimator, derived for max-stable vectors with unit Frechet margins, is

$$\hat{\varphi}_{\text{el}} = \frac{\sum_i p_i \mathbb{I}(w_i \leq 0.5) - 0.5}{0.5 - 2 \sum_i p_i (0.5 - w_i) \mathbb{I}(w_i \leq 0.5)}$$

where p_i is the empirical likelihood weight for observation i , \mathbb{I} is an indicator function and w_i is the pseudo-angle associated to the first coordinate, derived based on exceedances above u .

Value

an invisible data frame with columns

qllev quantile level of thresholds

coef extremal asymmetry coefficient estimates

lower either NULL or a vector containing the lower bound of the confidence interval

upper either NULL or a vector containing the lower bound of the confidence interval

References

Semadani, C. (2020). Inference on the Angular Distribution of Extremes, PhD thesis, EPFL, no. 8168.

Examples

```
## Not run:
samp <- rmev(n = 1000,
            d = 2,
            param = 0.2,
            model = "log")
xdep.asym(samp, confint = "wald")
xdep.asym(samp, method = "emplik", confint = "none")

## End(Not run)
```

xdep.chi

Coefficient of tail correlation

Description

The coefficient of tail correlation χ is

$$\chi = \lim_{u \rightarrow 1} \frac{\Pr(F_1(X_1) > u, \dots, F_D(X_D) > u)}{1 - u}.$$

Asymptotically independent vectors have $\chi = 0$. The estimator uses an estimator of the survival copula

Usage

```
xdep.chi(
  xdat,
  qlev = NULL,
  nq = 40,
  qlim = c(0.8, 0.99),
  estimator = c("emp", "betacop", "gpd", "hill"),
  confint = c("wald", "lrt"),
  level = 0.95,
  margtrans = c("emp", "none"),
  ties.method = "random",
  plot = TRUE,
  ...
)
```

Arguments

| | |
|-------------|--|
| xdat | an n by d matrix of multivariate observations |
| qlev | vector of percentiles between 0 and 1 |
| nq | number of quantiles of the structural variable at which to form a grid; only used if $u = \text{NULL}$. |
| qlim | limits for the sequence u of the structural variable |
| estimator | string giving estimator to employ |
| confint | string indicating the type of confidence interval, one of "wald" or "lrt" |
| level | the confidence level required (default to 0.95). |
| margtrans | string giving the marginal transformation, one of emp for rank-based transformation or none if data are already on the uniform scale |
| ties.method | string indicating the type of method for rank; see rank for a list of options. Default to "random" |
| plot | logical; if TRUE, return a plot |
| ... | additional arguments to <code>taildep</code> , currently ignored |

Value

a data frame

- qlev: quantile level of estimates
- coef: point estimates
- lower: lower bound of confidence interval
- upper: lower bound of confidence interval

Examples

```
## Not run:
set.seed(765)
# Max-stable model
dat <- rmev(n = 1000, d = 2, param = 0.7, model = "log")
xdep.chi(dat, confint = 'wald')

## End(Not run)
```

xdep.chibar

Coefficient chi-bar

Description

For data with unit Pareto margins, the coefficient $\bar{\chi} = 2\eta - 1$ is defined via

$$\Pr(\min(X) > x) = L(x)x^{-1/\eta},$$

where $L(x)$ is a slowly varying function. Ignoring the latter, several estimators of η can be defined. In unit Pareto margins, η is a nonnegative shape parameter that can be estimated by fitting a generalized Pareto distribution above a high threshold. In exponential margins, η is a scale parameter and the maximum likelihood estimator of the latter is the Hill estimator. Both methods are based on peaks-over-threshold and the user can choose between pointwise confidence obtained through a likelihood ratio test statistic ("lrt") or the Wald statistic ("wald").

Usage

```
xdep.chibar(
  xdat,
  qllev = NULL,
  nq = 40,
  qlim = c(0.8, 0.99),
  estimator = c("emp", "betacop"),
  confint = c("wald", "lrt"),
  level = 0.95,
  margtrans = c("emp", "none"),
  ties.method = "random",
  plot = TRUE,
  ...
)
```

Arguments

| | |
|-------|--|
| xdat | an n by d matrix of multivariate observations |
| qllev | vector of percentiles between 0 and 1 |
| nq | number of quantiles of the structural variable at which to form a grid; only used if $u = \text{NULL}$. |

| | |
|-------------|--|
| qlim | limits for the sequence u of the structural variable |
| estimator | string giving estimator to employ |
| confint | string indicating the type of confidence interval, one of "wald" or "lrt" |
| level | the confidence level required (default to 0.95). |
| margtrans | string giving the marginal transformation, one of emp for rank-based transformation or none if data are already on the uniform scale |
| ties.method | string indicating the type of method for rank; see rank for a list of options. Default to "random" |
| plot | logical; if TRUE, return a plot |
| ... | additional arguments to <code>taildep</code> , currently ignored |

Details

The most common approach for estimation is the empirical survival copula, by evaluating the proportion of sample minima with uniform margins that exceed a given x . An alternative estimator uses a smoothed estimator of the survival copula using Bernstein polynomial, resulting in the so-called betacop estimator. Approximate pointwise confidence intervals for the latter are obtained by assuming the proportion of points is binomial.

Value

a data frame

- `qllev`: quantile level of estimates
- `coef`: point estimates
- `lower`: lower bound of confidence interval
- `upper`: lower bound of confidence interval

References

Ledford, A.W. and J. A. Tawn (1996), Statistics for near independence in multivariate extreme values. *Biometrika*, **83**(1), 169–187.

Ledford, A.W. and J. A. Tawn (1996), Statistics for near independence in multivariate extreme values. *Biometrika*, **83**(1), 169–187.

Examples

```
## Not run:
set.seed(765)
# Max-stable model
dat <- rmev(n = 1000, d = 2, param = 0.7, model = "log")
xdep.chibar(dat, confint = 'wald')

## End(Not run)
```

xdep.eta

*Coefficient of tail dependence***Description**

For data with unit Pareto margins, the coefficient of tail dependence η is defined via

$$\Pr(\min(X) > x) = L(x)x^{-1/\eta},$$

where $L(x)$ is a slowly varying function. Ignoring the latter, several estimators of η can be defined. In unit Pareto margins, η is a nonnegative shape parameter that can be estimated by fitting a generalized Pareto distribution above a high threshold. In exponential margins, η is a scale parameter and the maximum likelihood estimator of the latter is the Hill estimator. Both methods are based on peaks-over-threshold and the user can choose between pointwise confidence obtained through a likelihood ratio test statistic ("lrt") or the Wald statistic ("wald").

Usage

```
xdep.eta(
  xdat,
  qllev = NULL,
  nq = 40,
  qlim = c(0.8, 0.99),
  estimator = c("emp", "betacop", "gpd", "hill", "kj"),
  confint = c("wald", "lrt"),
  level = 0.95,
  margtrans = c("emp", "sp", "none"),
  ties.method = "random",
  plot = TRUE,
  mqlev = NULL,
  ...
)
```

Arguments

| | |
|-----------|--|
| xdat | an n by d matrix of multivariate observations |
| qllev | vector of percentiles between 0 and 1 |
| nq | number of quantiles of the structural variable at which to form a grid; only used if $u = \text{NULL}$. |
| qlim | limits for the sequence u of the structural variable |
| estimator | string giving estimator to employ |
| confint | string indicating the type of confidence interval, one of "wald" or "lrt" |
| level | the confidence level required (default to 0.95). |
| margtrans | string giving the marginal transformation, one of emp for rank-based transformation or none if data are already on the uniform scale |

| | |
|-------------|--|
| ties.method | string indicating the type of method for rank; see rank for a list of options. Default to "random" |
| plot | logical; if TRUE, return a plot |
| mqlev | marginal quantile levels for semiparametric estimation for estimator kj; data above this are modelled using a generalized Pareto distribution. If missing, empirical estimation is used throughout |
| ... | additional arguments to <code>taildep</code> , currently ignored |

Details

The most common approach for estimation is the empirical survival copula, by evaluating the proportion of sample minima with uniform margins that exceed a given x . An alternative estimator uses a smoothed estimator of the survival copula using Bernstein polynomial, resulting in the so-called betacop estimator. Approximate pointwise confidence intervals for the latter are obtained by assuming the proportion of points is binomial.

Value

a data frame

- qllev: quantile level of estimates
- coef: point estimates
- lower: lower bound of confidence interval
- upper: lower bound of confidence interval

References

Ledford, A.W. and J. A. Tawn (1996), Statistics for near independence in multivariate extreme values. *Biometrika*, **83**(1), 169–187.

Examples

```
## Not run:
set.seed(765)
# Max-stable model
dat <- rmev(n = 1000, d = 2, param = 0.7, model = "log")
xdep.eta(dat, confint = 'wald')

## End(Not run)
```

xdep.xcoef

Coefficient of extremal index

Description

These functions estimate the extremal index using an approximate sample from the Frechet distribution.

Usage

```
xdep.xcoef(
  xdat,
  coord = NULL,
  thresh = NULL,
  estimator = c("schlather", "smith", "fmado"),
  margtrans = c("emp", "gev", "none"),
  ties.method = "random",
  prob = 0,
  plot = TRUE,
  ...
)
```

Arguments

| | |
|-------------|---|
| xdat | an n by D matrix of unit Frechet observations |
| coord | an optional d by D matrix of location coordinates |
| thresh | threshold parameter (default is to keep all data if prob = 0). |
| estimator | string indicating which estimator to compute, one of smith, schlather or fmado. |
| margtrans | string indicating which method to use to transform the margins to unit Frechet scale, either emp for nonparametric transformation via rank transform, gev for fit of generalized extreme value distribution to marginals, or none |
| ties.method | method for handling of ties in rank transformation |
| prob | probability of not exceeding threshold thresh |
| plot | logical; should cloud or matrix of pairwise empirical estimates be plotted? Default to TRUE. |
| ... | additional parameters passed to the function, currently ignored. |

Details

The **Smith** estimator: suppose $Z(x)$ is simple max-stable vector (i.e., with unit Frechet marginals). Then $1/Z$ is unit exponential and $1/\max(Z(s_1), Z(s_2))$ is exponential with rate $\theta = \max\{Z(s_1), Z(s_2)\}$. The extremal index for the pair can therefore be calculated using the reciprocal mean.

The **Schlather and Tawn** estimator: the likelihood of the naive estimator for a pair of two sites A is

$$\text{card} \left\{ j : \max_{i \in A} X_i^{(j)} \bar{X}_i > z \right\} \log(\theta_A) - \theta_A \sum_{j=1}^n \left[\max \left\{ z, \max_{i \in A} (X_i^{(j)} \bar{X}_i) \right\} \right]^{-1},$$

where $\bar{X}_i = n^{-1} \sum_{j=1}^n 1/X_i^{(j)}$ is the harmonic mean and z is a threshold on the unit Fréchet scale. The search for the maximum likelihood estimate for every pair A is restricted to the interval $[1, 3]$. A binned version of the extremal coefficient cloud is also returned. The Schlather estimator is not self-consistent. The Schlather and Tawn estimator includes as special case the Smith estimator if we do not censor the data ($p = 0$) and do not standardize observations by their harmonic mean.

The **F-madogram** estimator is a non-parametric estimate based on a stationary process Z ; the extremal coefficient satisfies

$$\theta(h) = \frac{1 + 2\nu(h)}{1 - 2\nu(h)},$$

where

$$\nu(h) = \frac{1}{2} \mathbb{E}[|F(Z(s+h)) - F(Z(s))|]$$

The implementation only uses complete pairs to calculate the relative ranks.

All estimators are coded in plain R and computations are not optimized. The estimation time can therefore be large for large data sets. If there are no missing observations, the routine `fmadogram` from the `SpatialExtremes` package should be preferred as it is noticeably faster.

The data will typically consist of max-stable vectors or block maxima. Both of the Smith and the Schlather–Tawn estimators require unit Fréchet margins; the margins will be standardized to the unit Fréchet scale, either parametrically or nonparametrically unless `margtrans = "none"`. If `margtrans = "gev"`, a parametric GEV model is fitted to each column of `dat` using maximum likelihood estimation and transformed back using the probability integral transform. If `method = "emp"`, using the empirical distribution function. The latter is the default, as it is appreciably faster.

Value

an invisible list with vectors `dist` if `coord` is non-null or else a matrix of pairwise indices `ind`, `extcoef` and the supplied estimator, `fmodo` and `binned`. If `estimator == "schlather"`, an additional matrix with 2 columns containing the binned distance `binned` with the `h` and the binned extremal coefficient.

References

- Schlather, M. and J. Tawn (2003). A dependence measure for multivariate and spatial extremes, *Biometrika*, **90**(1), pp. 139–156.
- Cooley, D., P. Naveau and P. Poncet (2006). Variograms for spatial max-stable random fields, In: Bertail P., Soulier P., Doukhan P. (eds) *Dependence in Probability and Statistics*. Lecture Notes in Statistics, vol. 187. Springer, New York, NY
- R. J. Erhardt, R. L. Smith (2012), Approximate Bayesian computing for spatial extremes, *Computational Statistics and Data Analysis*, **56**, pp.1468–1481.

Examples

```

## Not run:
coord <- 10 * cbind(runif(50), runif(50))
di <- as.matrix(dist(coord))
dat <- rmev(
  n = 1000,
  d = 100,
  param = 3,
  sigma = exp(-di / 2),
  model = 'xstud'
)
res <- xdep.xcoef(xdat = dat, coord = coord)
# Extremal Student extremal coefficient function

XT.extcoeffun <- function(h, nu, corrfun, ...) {
  if (!is.function(corrfun)) {
    stop('Invalid function \'corrfun\'.')
  }
  h <- unique(as.vector(h))
  rhoh <- sapply(h, corrfun, ...)
  cbind(
    h = h,
    extcoef = 2 * pt(sqrt((nu + 1) * (1 - rhoh) / (1 + rhoh)), nu + 1)
  )
}
#This time, only one graph with theoretical extremal coef
plot(res$dist, res$extcoef, ylim = c(1, 2), pch = 20)
abline(v = 2, col = 'gray')
extcoefxt <- XT.extcoeffun(
  seq(0, 10, by = 0.1),
  nu = 3,
  corrfun = function(x) {
    exp(-x / 2)
  }
)
lines(
  extcoefxt[, 'h'],
  extcoefxt[, 'extcoef'],
  type = 'l',
  col = 'blue',
  lwd = 2
)
# Brown--Resnick extremal coefficient function
BR.extcoeffun <- function(h, vario, ...) {
  if (!is.function(vario)) {
    stop('Invalid function \'vario\'.')
  }
  h <- unique(as.vector(h))
  gammah <- sapply(h, vario, ...)
  cbind(h = h, extcoef = 2 * pnorm(sqrt(gammah / 4)))
}
extcoefbr <- BR.extcoeffun(

```

```

    seq(0, 20, by = 0.25),
    vario = function(x) {
      2 * x^0.7
    }
  )
lines(
  extcoefbr[, 'h'],
  extcoefbr[, 'extcoef'],
  type = 'l',
  col = 'orange',
  lwd = 2
)

coord <- 10 * cbind(runif(20), runif(20))
di <- as.matrix(dist(coord))
dat <- rmev(
  n = 1000,
  d = 20,
  param = 3,
  sigma = exp(-di / 2),
  model = 'xstud'
)
res <- xdep.xcoef(
  xdat = dat,
  coord = coord,
  estimator = "smith"
)

## End(Not run)

```

xdep.xindex

Extremal index coefficient

Description

The function implements estimators of the extremal index based on interexceedance time and gap of exceedances. The maximum likelihood estimator and iteratively reweighted least square estimators of Suveges (2007) as well as the intervals estimator. The implementation differs from the presentation of the paper in that an iteration limit is enforced to make sure the iterative procedure terminates. Multiple thresholds can be supplied.

Usage

```

xdep.xindex(
  xdat,
  qllev = 0.95,
  estimator = c("wls", "mle", "intervals"),
  confint = c("none", "wald", "lrt"),
  level = 0.95,

```

```

    plot = FALSE,
    warn = FALSE,
    ...
)

```

Arguments

| | |
|-----------|---|
| xdat | numeric vector of observations |
| qlev | a vector of quantile levels in (0,1) for estimation of the extremal index. Defaults to 0.95 |
| estimator | a string specifying the chosen method (only one allowed). Must be either wls for weighted least squares, mle for maximum likelihood estimation or intervals for the intervals estimator of Ferro and Segers (2003). Partial match is allowed. |
| confint | string indicating the type of confidence interval, one of "wald" or "lrt" for estimator="mle", else "none" |
| level | the confidence level required (default to 0.95). |
| plot | logical; if TRUE, plot the extremal index as a function of q |
| warn | logical; if TRUE, receive a warning when the sample size is too small |
| ... | additional arguments, for backward compatibility |

Details

The iteratively reweighted least square is a procedure based on the gaps of exceedances $S_n = T_n - 1$. The model is first fitted to non-zero gaps, which are rescaled to have unit exponential scale. The slope between the theoretical quantiles and the normalized gap of exceedances is $b = 1/\theta$, with intercept $a = \log(\theta)/\theta$. As such, the estimate of the extremal index is based on $\hat{\theta} = \exp(\hat{a}/\hat{b})$. The weights are chosen in such a way as to reduce the influence of the smallest values. The estimator exploits the dual role of θ as the parameter of the mean for the interexceedance time as well as the mixture proportion for the non-zero component.

The maximum likelihood is based on an independence likelihood for the rescaled gap of exceedances, namely $\bar{F}(u_n)S(u_n)$. The score equation is equivalent to a quadratic equation in θ and the maximum likelihood estimate is available in closed form. Its validity requires however condition $D^{(2)}(u_n)$ to apply; this should be checked by the user beforehand.

A warning is emitted if the effective sample size is less than 50 observations.

Value

a data frame

- qlev: quantile level of estimates
- coef: point estimates
- lower: lower bound of confidence interval
- upper: lower bound of confidence interval

Author(s)

Leo Belzile

References

- Ferro and Segers (2003). Inference for clusters of extreme values, *JRSS: Series B*, **65**(2), 545-556.
- Suveges (2007) Likelihood estimation of the extremal index. *Extremes*, **10**(1), 41-55.
- Suveges and Davison (2010), Model misspecification in peaks over threshold analysis. *Annals of Applied Statistics*, **4**(1), 203-221.
- Fukutome, Liniger and Suveges (2015), Automatic threshold and run parameter selection: a climatology for extreme hourly precipitation in Switzerland. *Theoretical and Applied Climatology*, **120**(3), 403-416.

Examples

```
set.seed(234)
# Moving maxima model with theta=0.5
a <- 1; theta <- 1/(1+a)
sim <- rgev(10001, loc=1/(1+a), scale=1/(1+a), shape=1)
x <- pmax(sim[-length(sim)]*a, sim[-1])
q <- seq(0.9, 0.99, by = 0.01)
xdep.xindex(
  xdat = x,
  qlev = q,
  estimator = "mle",
  confint = "wald")
```

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